

A STUDY OF NON-HOMOGENEOUS ABSORBING MARKOV CHAINS

A Thesis

by

JOHN KEVIN BEAN

Submitted to the Graduate College of
Texas A&M University
in partial fulfillment of the requirement for the degree of

MASTER OF SCIENCE

December 1975

Major Subject: Mathematics

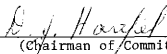
A STUDY OF NON-HOMOGENEOUS ABSORBING MARKOV CHAINS

A Thesis

by

JOHN KEVIN BEAN

Approved as to style and content by:



(Chairman of Committee)



(Head of Department)



(Member)



(Member)

December 1975

439326

ABSTRACT

A Study of Non-Homogeneous Absorbing Markov Chains. (December 1975)

John Kevin Bean, B.S., Texas A&M University

Chairman of Advisory Committee: Dr. D. J. Hartfiel

The objective of this thesis is to investigate general conditions which guarantee the existence of limits in non-homogeneous absorbing Markov chains. The major emphasis of this thesis will be a collection of theorems analogous to the classical results of Markov chain theory concerning the limiting behavior of Markov chains.

A first result generalizes the classical Markov chain result concerning the existence of $\lim_{k \rightarrow \infty} A^k$, where A is the transition matrix of the Markov chain. This result is generalized into a non-homogeneous Markov chain setting by allowing the transition matrix A to change at each step of time. This leads to the study of $\lim_{k \rightarrow \infty} A_1 \dots A_k$. Conditions necessary for the existence of this limit are determined.

A second result generalizes the classical Markov chain result concerning the existence of

$$\lim_{k \rightarrow \infty} \frac{I + A + \dots + A^{k-1}}{k}$$

where A is the transition matrix of the Markov chain. This result is also generalized to a non-homogeneous Markov chain result by allowing the transition matrix to change at each step of time. This leads to the study of

$$\lim_{k \rightarrow \infty} \frac{I + A_1 + \dots + A_1 \dots A_{k-1}}{k} .$$

Conditions are provided guaranteeing the existence of this limit.

The results obtained in this thesis are then applied to problems usually resolved by classical Markov chain theory, but are actually more suited to a non-homogeneous Markov chain solution.

ACKNOWLEDGMENTS

I wish to express particular gratitude to Dr. D. J. Hartfiel for his guidance, assistance and encouragement throughout my studies at Texas A&M.

I wish to thank Dr. James R. Boone and Dr. Ronald A. J. Schorn for being members of my advisory committee, and to thank Dr. George R. Blakley for his assistance also.

I wish to thank my dear friends Bill Slingerland, Beverly Barnes, John Taylor, Ron Mailloux and Charlotte de Spiegel for their encouragement and camaraderie throughout this study.

I wish to express particular appreciation to my parents, Mr. and Mrs. L. M. McGuire, without whose love, patience, encouragement and faith in me, none of this would have been possible.

TABLE OF CONTENTS

Chapter	Page
I. INTRODUCTION	1
Definitions and Notations.	1
Background and Statement of the Problem.	1
II. RESULTS CONCERNING NON-HOMOGENEOUS ABSORBING MARKOV CHAINS	4
Results Concerning the Limit of Non-Homogeneous Absorbing Markov Chains.	4
Results Concerning the Limit of the Arithmetic Mean of Non-Homogeneous Absorbing Markov Chains.	11
III. APPLICATIONS	15
IV. SUMMARY AND CONCLUSION	20
REFERENCES	21
VITA	22

CHAPTER I

INTRODUCTION

Definitions and Notations. The definitions and notations found in this thesis are consistent with those used in Gantmacher [2] and Hohn [5]. Exceptions to the above are the following.

All matrices considered herein are $n \times n$ and nonnegative where $n \geq 2$. For any matrix A_k , $a_{ij}^{(k)}$ will denote the ij^{th} entry of A_k . Likewise $x_i^{(k)}$ will denote the i^{th} entry in a vector x_k . Further, $|A|$ will denote the maximum absolute value of the entries of A and $r_i(A_k)$ will denote the i^{th} row sum of A_k . Finally, e^k will be used to denote the vector with all entries zero except the k^{th} entry which is one.

Background and Statement of the Problem. Probability theory is one of the most important and most studied of the applied mathematical fields. Probability study originated in the mid 1600's, with principle works attributed to Fermat, Pascal and Huygens, with later significant discoveries by Bernoulli, Gauss, Laplace and others. The study of Markov chains was begun in 1907 when Andrei Andreevich Markov investigated a sequence of dependent random variables [8]. The related sequences of probability matrices were called Markov chains. Further investigations established Markov chains as a great theoretical tool

The format of this thesis follows the style of the Proceedings of the American Mathematical Society.

in probability study. In the hands of physicists, this theory was transformed into an active means for the study of natural processes, among these Brownian motion, population growth and games of chance. Economic as well as other applications were also found.

Markov's work dealt only with finite chains. The denumerable extension was developed by A. N. Kolmogorov in 1936. Further research, all dealing with the convergence of homogeneous Markov chains, was done by W. Feller, J. Hadamard and A. Y. Khinchin, all of whom made significant contributions to the theory.

Two classical results of Markov chain theory on the limiting behavior of Markov chains concerns the existence of

$$(1) \lim_{k \rightarrow \infty} A^k$$

and

$$(2) \lim_{k \rightarrow \infty} \frac{I + A + A^2 + \dots + A^{k-1}}{k}$$

where A is the transition matrix of the Markov chain. These two results can be generalized into a non-homogeneous Markov chain setting by allowing the transition matrix to change at each step of time. This leads to the non-homogeneous Markov chain problems considered in this thesis. These concern the existence of

$$(1) \lim_{k \rightarrow \infty} A_1 \dots A_k$$

and

$$(2) \lim_{k \rightarrow \infty} \frac{I + A_1 + \dots + A_1 \dots A_{k-1}}{k}$$

In particular, in this thesis conditions are given which guarantee that

$$\lim_{k \rightarrow \infty} A_1 \dots A_k = \begin{bmatrix} I & 0 \\ N & 0 \end{bmatrix}$$

and conditions are given which guarantee that

$$\lim_{k \rightarrow \infty} \frac{1}{k} (I + A_1 + \dots + A_1 \dots A_{k-1}) = \begin{bmatrix} I & 0 \\ N & 0 \end{bmatrix}$$

where N is some matrix.

Finally the results of this thesis will be applied to problems usually arising in the theory of classical Markov chains, but which, in fact, are more appropriate to the theory of non-homogeneous Markov chains.

CHAPTER II

RESULTS CONCERNING LIMITS IN NON-HOMOGENEOUS ABSORBING

MARKOV CHAINS

This chapter, which constitutes the theoretical portion of the thesis, will deal with the theory of non-homogeneous absorbing Markov chains. The chapter is divided into two sections. The first section deals with the limiting behavior of a non-homogeneous absorbing Markov chain. In the second section, that work is applied to evaluate a limit concerning the arithmetic mean of a non-homogeneous absorbing Markov chain.

Results Concerning Limits in Non-Homogeneous Absorbing Markov Chains. In this section, necessary conditions will be given that will guarantee that $\lim_{k \rightarrow \infty} A_1 \dots A_k$ exists, where A_k is a stochastic matrix in an absorbing Markov chain for all k . A tool which will be employed in the development of these conditions is called the measure of full indecomposability [4] and is defined as follows.

$$\text{Let } \nu(A) = \min_{\substack{|R|+|C|=n \\ R, C \text{ non-empty subsets of } N}} \max_{\substack{i \in R \\ j \in C}} a_{ij}, \text{ where } R \text{ and } C \text{ denote non-}$$

empty subsets of row and column indices, respectively, with $|N|$ being the number of elements in set N .

For example,

$$\text{if } A = \begin{bmatrix} 3 & 1 & 0 \\ 6 & 0 & 3 \\ 0 & 2 & 2 \end{bmatrix}, \text{ then } \nu(A) = 1$$

while

$$\text{if } A = \begin{bmatrix} 2 & 0 & 4 \\ 3 & 3 & 7 \\ 0 & 9 & 3 \end{bmatrix}, \text{ then } \mu(A) = 2.$$

The fundamental result given in this section is achieved through the following sequence of lemmas developing the properties of μ as related to the product of matrices.

Lemma 1. Suppose A_1, A_2, \dots, A_{n-1} is a sequence of matrices with $\mu(A_k) > 0$ for all k . Then, $A_{n-1} \dots A_2 A_1 > 0$.

Proof. Let $A \geq 0$ be a matrix such that $\mu(A) > 0$. For any column vector $x \geq 0$ in \mathbb{R}^n , define $|x|_+$ as the number of nonzero entries in x . Choose any column vector $x \geq 0$ such that x has at least one zero and at least one nonzero entry. Set $|x|_+ = s$, i.e. x has s nonzero entries. Then, there exists a permutation matrix P_1 which permutes the nonzero entries of x to the first s positions, i.e.

$$P_1 x = y = \begin{bmatrix} y_1 \\ \vdots \\ y_s \\ 0 \\ \vdots \\ 0 \end{bmatrix}.$$

Since P_1 is a permutation matrix, $P_1^t P_1 = I$. Thus, $Ax = A P_1 x = (A P_1^t) (P_1 x) = A P_1^t y$. Now, choose a second permutation matrix P_2 such that

$$P_2 A P_1^t = \left(\begin{array}{c|ccc} B & \text{---} & \text{---} & \text{---} \\ \hline 0 & \text{---} & \text{---} & \text{---} \end{array} \right)$$

where B is $r \times s$ and B contains no zero rows. Thus, $P_2 Ax = (P_2 A P_1^t) y$.

Clearly, $|y|_+ = |x|_+ = s$. Since B has at least one nonzero entry

in each row, $|P_2Ax|_+ = |Ax|_+ = r$. Further, the zero-block of $P_2AP_1^t$ is $(n-r) \times s$. As we know by hypothesis that $\mu(A) > 0$, it follows that $n - r + s < n$. Therefore, $r > s$, i.e. $|x|_+ < |Ax|_+$.

Now, consider e^i for $i \in \{1, 2, \dots, n\}$. Clearly, $|e^i|_+ = 1$ for all i . Choose some arbitrary e^i . Then, $|A_1 e^i|_+ > |e^i|_+ = 1$. Hence, by induction $|A_k \dots A_1 e^i|_+ > k$ for all $k < n$ so that $|A_{n-1} \dots A_1 e^i|_+ > n - 1$. Thus, $A_{n-1} \dots A_1 e^i > 0$ and since e^i was arbitrarily chosen, $A_{n-1} \dots A_1 > 0$.

Lemma 2. Suppose A is a matrix with $\mu(A) > 0$. Then, there exists a matrix B such that

- (i) $A \geq B$ and
 (ii) $\min_{b_{ij} > 0} b_{ij} = \mu(B) = \mu(A)$.

Proof. Construct a matrix B as follows: Set

$$b_{ij} = \begin{cases} a_{ij} & \text{if } a_{ij} \geq \mu(A) \\ 0 & \text{otherwise} \end{cases}.$$

Clearly, $A \geq B \geq 0$ and hence $\mu(A) \geq \mu(B)$.

Now, consider an arbitrary submatrix L in B such that L is in rows with indices in R and in columns with indices in C , where $|R| + |C| = n$. Let L' be the submatrix in A with row indices in R and column indices in C . From the definition of $\mu(A)$, we can find some ℓ'_{ij} such that $\ell'_{ij} \geq \mu(A)$. Thus, $\ell_{ij} \geq \mu(A)$ and since L was arbitrarily chosen, $\mu(B) \geq \mu(A)$. Therefore, $\mu(B) = \mu(A)$.

Finally, consider $\min_{b_{ij} > 0} b_{ij}$. By the construction of B,

$\min_{b_{ij} > 0} b_{ij} \geq \mu(A)$. As $\mu(B) = \mu(A)$, there is a $b_{ij} > 0$ so that $b_{ij} = \mu(A)$.

Hence $\min_{b_{ij} > 0} b_{ij} = \mu(B) = \mu(A)$.

Lemma 3. Suppose A_1, A_2, \dots, A_{n-1} is a sequence of matrices with $\mu(A_1) > 0, \dots, \mu(A_{n-1}) > 0$. Then, $\min_{i,j} (A_1 A_2 \dots A_{n-1})_{ij} \geq \mu(A_1) \dots \mu(A_{n-1})$.

Proof. Consider matrices A_1, A_2, \dots, A_{n-1} . Construct matrices

B_1, B_2, \dots, B_{n-1} as follows: Set

$$b_{ij}^{(k)} = \left\{ \begin{array}{ll} a_{ij}^{(k)} & \text{if } a_{ij}^{(k)} \geq \mu(A_k) \\ 0 & \text{otherwise} \end{array} \right\}$$

so that by Lemma 2, $\mu(B_k) = \mu(A_k) > 0$ for all $k \leq n-1$. Let

$$C = B_1 B_2 \dots B_{n-1}.$$

Now, since

$$\min_{i,j} (B_1 B_2)_{ij} = \min_{i,j} \sum_{i_1}^{(1)} b_{i i_1}^{(1)} b_{i_1 j}^{(2)},$$

it follows by induction that

$$\min_{i,j} c_{ij} = \min_{i,j} (B_1 \dots B_{n-1})_{ij} = \min_{i,j} \sum_{i_1 \dots i_{n-2}} b_{i i_1}^{(1)} \dots b_{i_{n-2} j}^{(n-1)}.$$

This fact, in conjunction with Lemma 1, implies

$$c_{ij} = (B_1 \dots B_{n-1})_{ij} = \sum_{i_1 \dots i_{n-2}} b_{i i_1}^{(1)} \dots b_{i_{n-2} j}^{(n-1)} > 0.$$

Thus, some $b_{i_1 i_1}, \dots, b_{i_{n-2} j} > 0$. Now, $b_{i_1 i_1} \geq \mu(A_1), \dots, b_{i_{n-2} j} \geq \mu(A_{n-1})$

so that $c_{ij} \geq \mu(A_1)\mu(A_2)\dots\mu(A_{n-1})$. Hence, $\min_{i,j} c_{ij} \geq \mu(A_1)\dots\mu(A_{n-1})$.

Finally, as $A_k \geq B_k$, $(A_1 \dots A_{n-1})_{ij} \geq (B_1 \dots B_{n-1})_{ij}$ for all i and j so

that

$$\min_{i,j} (A_1 \dots A_{n-1})_{ij} \geq \min_{i,j} (B_1 \dots B_{n-1})_{ij} \geq \mu(A_1)\dots\mu(A_{n-1}).$$

By applying these three lemmas, conditions will now be given that guarantee that a sequence of matrices A_1, \dots, A_k, \dots is such that

$\lim_{k \rightarrow \infty} A_1 \dots A_k = 0$. The conditions will only require certain conditions

imposed on the $\mu(A_k)$ and the row sums of A_k , for each k .

Theorem 1. Suppose $A_1, \dots, A_{n-1}, \dots$ is a sequence of matrices such that $\mu(A_1) \geq \mu, \dots, \mu(A_k) \geq \mu, \dots$ where $\mu > 0$,

$$(i) \max_i r_i(A_k) \leq 1 \text{ and}$$

(ii) there exists $\delta_1, 0 < \delta_1 < 1$, such that for all k there

$$\text{exists } i(k) \text{ such that } r_{i(k)}(A_k) \leq \delta_1.$$

Then $\lim_{k \rightarrow \infty} A_1 \dots A_k = 0$.

Proof. Consider $A_1 \dots A_k \dots$. Choose $B_1 = (A_1 \dots A_{n-1})$,

$$B_2 = (A_n \dots A_{2(n-1)}), \dots, B_s = \prod_{s(n-1)+1}^{(s+1)(n-1)} A_k. \text{ By Lemma 3, } b_{ij}^{(s)} \geq \mu^{n-1}$$

for each i, j and s .

Consider the product $B_1 \dots B_k$. Let $B_s B_{s+1} = C$. Then,

$$\begin{aligned}
\sum_{k=1}^n c_{ik} &= \sum_{k=1}^n \left(\sum_{t=1}^n b_{it}^{(s)} b_{tk}^{(s+1)} \right) = \sum_{t=1}^n \sum_{k=1}^n b_{it}^{(s)} b_{tk}^{(s+1)} \\
&= \sum_{t=1}^n b_{it}^{(s)} r_t(B_{s+1}) \\
&\leq \min_i b_{it}^{(s)} \delta + 1 - \min_i b_{it}^{(s)} \\
&= \mu^{n-1} \delta + 1 - \mu^{n-1} \\
&= 1 - \mu^{n-1}(1-\delta) < 1.
\end{aligned}$$

Hence, $r_k[B_s B_{s+1}] \leq 1 - \mu^{n-1}(1-\delta) < 1$ for all k . Set $[1 - \mu^{n-1}(1-\delta)] = \ell$.

Thus, $\max_i r_i[B_s B_{s+1}] \leq \ell < 1$.

$$\begin{aligned}
\text{Now, } r_i(XY) &= r_i \left(\sum_{t=1}^n X_{it} Y_{tk} \right) = \sum_{k=1}^n \sum_{t=1}^n X_{it} Y_{tk} \\
&= \sum_{k=1}^n X_{ik} r_k(Y) \leq r_i(X) \max_i r_i(Y).
\end{aligned}$$

Since X and Y were arbitrarily chosen, this holds for any X and Y .

By applying these results, we have

$$\begin{aligned}
r_i(B_{s+1} \dots B_{s+2m}) &\leq \max_i r_i(B_{s+1} B_s) \dots \max_i r_i(B_{s+2m-1} B_{s+2m}) \\
&\leq \ell \cdot \ell \cdots \ell = \ell^m \text{ for all } i.
\end{aligned}$$

Now, let m approach ∞ . As $m \rightarrow \infty$, $\ell^m \rightarrow 0$. Hence, for all

i , $r_i(B_1 \dots B_m) \rightarrow 0$ and $\lim_{k \rightarrow \infty} B_1 \dots B_k = 0$. Therefore, $\lim_{k \rightarrow \infty} A_1 \dots A_k = 0$.

This result finds its use in the following corollary which provides conditions guaranteeing the existence of a limit of a non-homogeneous absorbing Markov chain.

Corollary 1. Suppose $C_1, C_2, \dots, C_k, \dots$ is a sequence of stochastic matrices and

$$C_k = \begin{bmatrix} I & 0 \\ B_k & A_k \end{bmatrix}$$

for all k .

Further suppose $\lim_{k \rightarrow \infty} A_1 \dots A_k = 0$. Then, $\lim_{k \rightarrow \infty} C_1 \dots C_k$ exists.

Proof. Set the product

$$C_1 \dots C_k = \begin{bmatrix} I & 0 \\ L_k & A_1 \dots A_k \end{bmatrix}$$

As $\lim_{k \rightarrow \infty} A_1 \dots A_k = 0$, to show $\lim_{k \rightarrow \infty} C_1 \dots C_k$ exists it is only necessary

to prove L_k has a limit. Note that

$$L_2 = B_1 + A_1 B_2$$

$$L_3 = B_1 + A_1 B_2 + A_1 A_2 B_3$$

⋮

$$L_k = B_1 + A_1 B_2 + \dots + A_1 \dots A_{k-1} B_k$$

Hence, $\ell_{ij}^{(1)} \leq \ell_{ij}^{(2)} \leq \dots \leq \ell_{ij}^{(k)}$ for all i and j . Thus, the sequence

L_1, \dots, L_k, \dots is monotonically increasing componentwise.

We know C_k is stochastic for each k . Further, $C_1 C_2$ is stochastic

and hence by induction $C_1 \dots C_k$ is stochastic. Thus, 1 is an upper

bound for $\ell_{ij}^{(k)}$, i.e. $\ell_{ij}^{(k)} \leq 1$, for all i, j and k . Therefore,

$\lim_{k \rightarrow \infty} \ell_{ij}^{(k)}$ exists and hence $\lim_{k \rightarrow \infty} L_k$ exists, which gives the result of

the corollary.

From this corollary, it is seen that

$$\lim_{k \rightarrow \infty} C_1 \dots C_k = \begin{bmatrix} I & 0 \\ N & 0 \end{bmatrix}$$

where N is some matrix. However, by reviewing the proof of the corollary, a recipe for the calculation of the matrix N is not achievable. In fact, for the general problem, it is doubtful that a recipe exists. Some relief for this problem, however, can be obtained by using the work in the proof of Theorem 1. From this work, it is seen that given $\epsilon > 0$ a k can be calculated so that

$$\left| C_1 \dots C_k - \begin{bmatrix} I & 0 \\ N & 0 \end{bmatrix} \right| < \epsilon .$$

Further, this k depends on μ so that μ measures the rate of convergence of $C_1 \dots C_k$ to $\begin{bmatrix} I & 0 \\ N & 0 \end{bmatrix}$.

This then concludes this section. In the next section, the results of this section will be applied to a study of the limit of the arithmetic mean of a non-homogeneous absorbing Markov chain.

Results Concerning the Limit of the Arithmetic Mean of a Non-Homogeneous Absorbing Markov Chain. It is the intent of this section to provide necessary conditions for the existence of the limit of the arithmetic mean of a non-homogeneous absorbing Markov chain. The result is based on the following theorem.

Theorem 2. Suppose C_1, \dots, C_k, \dots is a sequence of stochastic matrices such that

$$C_k = \begin{bmatrix} I & 0 \\ B_k & A_k \end{bmatrix}$$

for all k . Suppose further that $\lim_{k \rightarrow \infty} A_1 \dots A_k = 0$. Then,

$$\lim_{k \rightarrow \infty} \frac{1}{k} (C_1 + C_1 C_2 + \dots + C_1 \dots C_k) = \lim_{k \rightarrow \infty} C_1 \dots C_k .$$

Proof. As in Corollary 1, set

$$C_1 = \begin{bmatrix} I_1 & 0_1 \\ L_1 & A_1 \end{bmatrix}, \quad C_1 C_2 = \begin{bmatrix} I_2 & 0_2 \\ L_2 & A_1 A_2 \end{bmatrix}, \dots, C_1 \dots C_k \\ = \begin{bmatrix} I_k & 0_k \\ L_k & A_1 \dots A_k \end{bmatrix} .$$

We will examine the expression $\frac{1}{k}(C_1 + \dots + C_1 \dots C_k)$ blockwise.

Consider first the identity blocks, which give

$$\frac{1}{k}(I_1 + \dots + I_k) = \left(\frac{1}{k}\right)(k)(I) = I .$$

Likewise, the zero blocks yield

$$\frac{1}{k}(0_1 + \dots + 0_k) = \left(\frac{1}{k}\right)(0) = 0 .$$

Now, consider the A blocks. By hypothesis, given $\epsilon > 0$ there exists

$R > 0$ such that for $R \leq k$, $|A_1 \dots A_k| \leq \epsilon$. Thus,

$$\begin{aligned} \left| \frac{A_1 + \dots + A_1 \dots A_k}{k} \right| &\leq \frac{|A_1|}{k} + \dots + \left| \frac{A_1 \dots A_{R-1}}{k} \right| + \left| \frac{A_1 \dots A_R}{k} \right| + \dots + \left| \frac{A_1 \dots A_k}{k} \right| \\ &\leq \frac{R-1}{k} + \epsilon \left(\frac{k-R+1}{k} \right). \end{aligned}$$

Hence, for k sufficiently large

$$\left| \frac{A_1 + \dots + A_1 \dots A_k}{k} \right| \leq 2\epsilon.$$

Since ϵ was arbitrary, $\lim_{k \rightarrow \infty} \frac{1}{k} (A_1 + \dots + A_1 \dots A_k) = 0$.

Finally, consider the L blocks. We know from Corollary 1 that $L_k \rightarrow N$ for some matrix N . Let $\epsilon > 0$. Choose M such that for $M \leq k$, we have

$$|L_k - N| \leq \epsilon$$

and thus

$$N - \epsilon P \leq L_k \leq N + \epsilon P$$

where $p_{i,j} = 1$ for all i and j . Hence,

$$\begin{aligned} \frac{L_1 + \dots + L_{M-1} + (N - \epsilon P) + \dots + (N - \epsilon P)}{k} &\leq \frac{L_1 + \dots + L_k}{k} \\ &\leq \frac{L_1 + \dots + L_{M-1} + (N + \epsilon P) + \dots + (N + \epsilon P)}{k} \end{aligned}$$

and

$$\frac{L_1 + \dots + L_{M-1}}{k} + \frac{k-M+1}{k}(N-\varepsilon P) \leq \frac{L_1 + \dots + L_k}{k} \leq \frac{L_1 + \dots + L_{M-1}}{k} + \frac{k-M+1}{k}(N+\varepsilon P)$$

Hence, for sufficiently large k ,

$$N - 2\varepsilon P \leq \frac{1}{k}(L_1 + \dots + L_k) \leq N + 2\varepsilon P.$$

As ε was arbitrary, $\lim_{k \rightarrow \infty} \frac{1}{k}(L_1 + \dots + L_k) = N$. Therefore,

$$\lim_{k \rightarrow \infty} \frac{C_1 + \dots + C_1 \dots C_k}{k} = \begin{bmatrix} I & 0 \\ N & 0 \end{bmatrix}.$$

The existence of the arithmetic mean of a non-homogeneous absorbing Markov chain can now be established.

Corollary 2. Suppose C_1, \dots, C_k, \dots is a sequence of stochastic matrices such that

$$C_k = \begin{bmatrix} I & 0 \\ B_k & A_k \end{bmatrix}$$

for all k . Suppose further that $\lim_{k \rightarrow \infty} A_k = 0$. Then,

$$\lim_{k \rightarrow \infty} \frac{I + C_1 + \dots + C_1 \dots C_{k-1}}{k} = \lim_{k \rightarrow \infty} C_1 C_2 \dots C_k.$$

Proof. The proof follows by noting that

$$\lim_{k \rightarrow \infty} \frac{I + \dots + C_1 \dots C_{k-1}}{k} = \lim_{k \rightarrow \infty} \frac{I}{k} + \lim_{k \rightarrow \infty} \left(\frac{k-1}{k}\right) \frac{C_1 + \dots + C_1 \dots C_{k-1}}{k-1} = \lim_{k \rightarrow \infty} C_1 \dots C_k.$$

This then concludes our theoretical work on non-homogeneous absorbing Markov chains. In the following chapter, the use of these results will be shown by applying this work in the solution of several practical problems.

CHAPTER III

APPLICATIONS

The main objective of this chapter is to provide several applications of non-homogeneous absorbing Markov chain theory. It is the intent to apply the results of Chapter II to problems which have in the past been resolved by classical Markov chain theory, but which in fact are more of the non-homogeneous Markov chain variety.

The initial problem concerns actual problems incurred in space flight. To describe this problem simply, suppose three light bulbs are utilized in the lighting system of a lunar module. As there is no room for excess baggage, no replacement bulbs are brought on the trip, even though the bulbs may burn out, thus leaving the module in darkness. Hence, there is a need for computing the probability of the lights continuing to burn throughout the flight period. Time is a factor in this problem as the filaments become thinner as time passes, thus the probability of the lights burning out increases with time.

Let the k^{th} transition matrix for this non-homogeneous absorbing Markov chain be

$$C_k = \begin{bmatrix} 1 & 0 & 0 & 0 \\ .4\left(\frac{2k}{k+1}\right) & .6\left(\frac{k+5}{3k+3}\right) & 0 & 0 \\ .3\left(\frac{2k+1}{3k+3}\right) & .3\left(\frac{6k+1}{3k+3}\right) & .4\left(\frac{k+4}{2k+2}\right) & 0 \\ .1\left(\frac{2k+3}{k+1}\right) & .2\left(\frac{k+2}{k+1}\right) & .3\left(\frac{4k+1}{3k+3}\right) & .4\left(\frac{k+2}{2k+2}\right) \end{bmatrix}$$

for all k , where

- $c_{11}^{(k)}$ = probability zero lights are on, given zero lights were on the previous hour,
- $c_{12}^{(k)}$ = probability one light is on, given zero lights were on the previous hour,
- $c_{13}^{(k)}$ = probability two lights are on, given zero lights were on the previous hour,
- $c_{14}^{(k)}$ = probability three lights are on, given zero lights were on the previous hour,
- $c_{21}^{(k)}$ = probability zero lights are on, given one light was on the previous hour,
- $c_{22}^{(k)}$ = probability one light is on, given one light was on the previous hour,
- $c_{23}^{(k)}$ = probability two lights are on, given one light was on the previous hour,
- $c_{24}^{(k)}$ = probability that three lights are on, given one light was on the previous hour,
- $c_{31}^{(k)}$ = probability zero lights are on, given two lights were on the previous hour,
- $c_{32}^{(k)}$ = probability one light is on, given two lights were on the previous hour,
- $c_{33}^{(k)}$ = probability two lights are on, given two lights were on the previous hour,
- $c_{34}^{(k)}$ = probability three lights are on, given two lights were on the previous hour,
- $c_{41}^{(k)}$ = probability zero lights are on, given three lights were on the previous hour,
- $c_{42}^{(k)}$ = probability one light is on, given three lights were on the previous hour,
- $c_{43}^{(k)}$ = probability two lights are on, given three lights were on the previous hour,
- $c_{44}^{(k)}$ = probability three lights are on, given three lights were on the previous hour.

Using the results developed in Chapter II, we can predict the

long-run behavior of these light bulbs. Clearly, the light bulbs will eventually all burn out and

$$\lim_{k \rightarrow \infty} C_1 \dots C_k = \begin{bmatrix} 1 & 0 & 0 & 0 \\ n_1 & 0 & 0 & 0 \\ n_2 & 0 & 0 & 0 \\ n_3 & 0 & 0 & 0 \end{bmatrix}$$

where $n_3 = 1$ is the probability that all lights are burned out in the long run. However, space flight time is not infinite, and the time of flight is usually known. Suppose the flight time is 100 hours. Then, a computer can be used to compute $C_1 \dots C_{100}$, which will contain $\ell_{31}^{(100)}$ the probability that all of the lights will be burned out after the 100 hours in flight.

A second problem concerns a physiological disturbance. For this, assume there is a small tar particle trapped in the lungs of an individual. For the particle to be removed from his body, it must pass through his throat into the atmosphere. Once it escapes into the atmosphere, we will assume it cannot return into his body. We will also assume that his lungs will grow weaker at each passing exhalation, the unit of time utilized in these experiments, thus making the problem time dependent. Let the k^{th} transition matrix for this non-homogeneous absorbing Markov chain be

$$C_k = \begin{bmatrix} 1 & 0 & 0 \\ .4 \left(\frac{k+2}{2k+2} \right) & .6 \left(\frac{3k+1}{3k+3} \right) & .2 \left(\frac{k+2}{k+1} \right) \\ 0 & .01 \left(\frac{k+55}{k+1} \right) & .99 \left(\frac{22k+1}{22k+22} \right) \end{bmatrix}$$

where

$c_{11}^{(k)}$ = probability tar is in air, given it was in air at previous exhalation,

$c_{12}^{(k)}$ = probability tar is in throat, given it was in air at previous exhalation,

$c_{13}^{(k)}$ = probability tar is in lungs, given it was in air at previous exhalation,

$c_{21}^{(k)}$ = probability tar is in air, given it was in throat at previous exhalation,

$c_{22}^{(k)}$ = probability tar is in throat, given it was in throat at previous exhalation,

$c_{23}^{(k)}$ = probability tar is in lungs, given it was in throat at previous exhalation,

$c_{31}^{(k)}$ = probability tar is in air, given it was in lungs at previous exhalation,

$c_{32}^{(k)}$ = probability that tar is in throat, given it was in lungs at previous exhalation,

$c_{33}^{(k)}$ = probability that tar is in lungs, given it was in lungs at previous exhalation.

In this example, Corollary 1 yields

$$\lim C_1 \dots C_k = \begin{bmatrix} 1 & 0 & 0 \\ n_1 & 0 & 0 \\ n_2 & 0 & 0 \end{bmatrix}$$

where n_1 is the probability of the particle going from the throat into the atmosphere and n_2 is the probability of the particle going from the lungs into the atmosphere in the long run. Since $C_1 \dots C_k$ is stochastic, $n_1 + n_2 = 1$. Therefore, the tar particle will eventually be in the atmosphere.

These are but two applications indicating the use of non-homogeneous absorbing Markov chain theory. Countless others also exist. For example, in games of chance, in mass transportation, as well as in economics. Hence, while non-homogeneous absorbing Markov chains have great theoretical value, there is also a practical viewpoint which is of interest.

CHAPTER IV

SUMMARY AND CONCLUSION

The work contained in this thesis provides conditions for the existence of limits in non-homogeneous absorbing Markov chains. The two primary results yield conditions which insure the existence of the limits:

$$(1) \lim_{k \rightarrow \infty} A_1 \dots A_k$$

$$(2) \lim_{k \rightarrow \infty} \frac{I + A_1 + \dots + A_1 \dots A_{k-1}}{k},$$

where each A_k is stochastic. In addition, several practical applications of non-homogeneous absorbing Markov chains have been described.

Further investigation of the theory of non-homogeneous absorbing Markov chains will no doubt be undertaken. This work can attempt to generalize the results in this thesis. Further, as this thesis has dealt only with finite Markov chains, the extension of the results of this thesis to infinite chains is also a goal.

Finally, little work has been done in the area of providing applications for non-homogeneous Markov chain theory and practically none are given for the absorbing variety. This area can also be expanded. It seems that uses can be found in countless occurrences, as indicated in Chapter III.

REFERENCES

1. J. C. Davis, "A study of strong ergodicity in non-homogeneous Markov chains," M.S. thesis, Texas A&M University, College Station, Texas, 1974.
2. F. R. Cantmacher, The theory of matrices, vol. II, Chelsea Publishing Company, New York, 1964.
3. B. V. Gnedenko and A. Y. Khinchin, An elementary introduction to the theory of probability, 5th Edition, Dover Publications, New York, 1962.
4. D. J. Hartfiel, Results on measures of irreducibility and full indecomposability, Trans. Amer. Math. Soc., vol. II, 1975, 357-368.
5. F. E. Hohn, Elementary matrix algebra, The Macmillan Company, New York, 1964.
6. J. G. Kemeny and J. L. Snell, Finite Markov chains, D. Van Nostrand Company, New York, 1960.
7. L. E. Maistrov, Probability theory, a historical sketch, Academic Press, New York, 1974.
8. A. A. Markov, Investigation of a noteworthy case of dependent trials, Isv. Ros. Akad. Nauk, vol. I, 1907, 61-80.

VITA

My name is John Kevin Bean. I was born in Washington, D.C., on August 2, 1952. My parents are John William Bean and Gloria C. McGuire. I graduated from Italy High School, Italy, Texas, in 1970. I received a B.S. degree in mathematics from Texas A&M University, College Station, Texas, on December 15, 1973. I have five semesters college teaching experience as a graduate teaching assistant with the Department of Mathematics at Texas A&M University. My permanent address is Route 2, Italy, Texas, 76651.

The typist for this thesis was Jan Want.