

Charge-separated atmospheric neutrino-induced muons in the MINOS far detector

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We found 140 neutrino-induced muons in 854.24 live days in the MINOS far detector, which has an acceptance for neutrino-induced muons of $6.91 \times 10^6 \text{ cm}^2 \text{ sr}$. We looked for evidence of neutrino disappearance in this data set by computing the ratio of the number of low momentum muons to the sum of the number of high momentum and unknown momentum muons for both data and Monte Carlo expectation in the absence of neutrino oscillations. The ratio of data and Monte Carlo ratios, \mathcal{R} , is $\mathcal{R} = 0.65_{-0.12}^{+0.15}(\text{stat}) \pm 0.09(\text{syst})$, a result that is consistent with an oscillation signal. A fit to the data for the oscillation parameters $\sin^2 2\theta_{23}$ and Δm_{23}^2 excludes the null oscillation hypothesis at the 94% confidence level. We separated the muons into μ^- and μ^+ in both the data and Monte Carlo events and found the ratio of the total number of μ^- to μ^+ in both samples. The ratio of those ratios, $\hat{\mathcal{R}}_{CPT}$, is a test of *CPT* conservation. The result $\hat{\mathcal{R}}_{CPT} = 0.72_{-0.18}^{+0.24}(\text{stat})_{-0.04}^{+0.08}(\text{syst})$ is consistent with *CPT* conservation.

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I. INTRODUCTION

Measurements of atmospheric neutrinos by the Super-Kamiokande experiment have shown that there is a deficit of ν_μ when compared to expectations [1–9]. The hypothesis that best describes this deficit is the oscillation of $\nu_\mu(\bar{\nu}_\mu) \rightarrow \nu_\tau(\bar{\nu}_\tau)$ [10,11], with the oscillation probability given by

$$P_{\nu_\mu \rightarrow \nu_\tau} = \sin^2 2\theta_{23} \sin^2(1.27 \Delta m_{23}^2 L/E), \quad (1)$$

where θ_{23} is the mixing angle, $\Delta m_{23}^2 = |m_3^2 - m_2^2|$ is the mass squared difference in eV^2 between the neutrino mass states, L is the distance in km traveled by the neutrino, or its baseline, and E is the energy of the ν_μ in GeV [12]. The Super-Kamiokande data are best fit by the oscillation hypothesis with parameters $(\sin^2 2\theta_{23}, \Delta m_{23}^2) = (1.0, 2.4 \times 10^{-3} \text{ eV}^2)$ [8,9]; the ranges for these parameters given by the 90% confidence contours of the zenith angle oscillation fit are $\sin^2 2\theta_{23} > 0.92$ and $1.5 < \Delta m_{23}^2 < 3.4 \times 10^{-3} \text{ eV}^2$. The MACRO [13,14] and Soudan 2 [15,16] results are consistent with those obtained by Super-Kamiokande. For the MINOS analysis of atmos-

pheric neutrinos with an interaction vertex in the detector, the parameter ranges are $\sin^2 2\theta_{23} > 0.2$ and $7 \times 10^{-5} < \Delta m_{23}^2 < 5 \times 10^{-2} \text{ eV}^2$ [17]. Below, we extend the MINOS atmospheric analysis to neutrino-induced muons observed in the MINOS far detector.

The oscillation hypothesis for the atmospheric neutrino deficit has received strong support from the first results of the MINOS long baseline experiment. MINOS, which sends ν_μ produced at Fermilab to a detector 735 km away in northern Minnesota [18], finds $(\sin^2 2\theta_{23}, \Delta m_{23}^2) = (1.00, 2.74 \times 10^{-3} \text{ eV}^2)$ [19], when fitting their ν_μ events; the ranges of these parameters given by the 90% confidence contours are $\sin^2 2\theta_{23} > 0.72$ and $2.2 < \Delta m_{23}^2 < 3.8 \times 10^{-3} \text{ eV}^2$. The K2K long baseline experiment, which sent ν_μ produced at KEK to the Super-Kamiokande detector located 250 km away, also found consistent results— $(1.0, 2.8 \times 10^{-3} \text{ eV}^2)$ [20,21] with 90% confidence ranges for the parameters of $\sin^2 2\theta_{23} > 0.55$ and $1.9 < \Delta m_{23}^2 < 3.6 \times 10^{-3} \text{ eV}^2$.

Since the magnetized MINOS far detector distinguishes μ^- from μ^+ , MINOS data can be used as a probe of *CPT* conservation in the neutrino sector. *CPT* conservation requires that the probability of an atmospheric ν_μ of a given energy remaining a ν_μ after traveling from its origin to its point of interaction be equal to the probability that an atmospheric $\bar{\nu}_\mu$ of the same energy remains a $\bar{\nu}_\mu$ after traveling the same distance [11]. We consequently test *CPT* conservation by comparing ν_μ -induced μ^- and $\bar{\nu}_\mu$ -induced μ^+ with respect to expectations, as a measure of whether the atmospheric ν_μ and $\bar{\nu}_\mu$ are disappearing at the same rate. Effects induced by charged current interactions of the ν_e components of the neutrino eigenstates with the Earth could masquerade as apparent *CPT* violation. However, distortions from matter effects are estimated to

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be small for most of the θ_{13} range compatible with the CHOOZ limit [22]. The average induced $\nu_{\mu}/\bar{\nu}_{\mu}$ event rate asymmetry is estimated to be less than 10% [23]. Therefore, at our current level of statistical accuracy we can ignore matter effects and describe $\nu_{\mu}(\bar{\nu}_{\mu}) \rightarrow \nu_{\tau}(\bar{\nu}_{\tau})$ oscillations using the two flavor approximation in Eq. (1). Previous tests of *CPT* conservation with atmospheric neutrinos have been made by Super-Kamiokande [24] and MINOS [17]. As Super-Kamiokande cannot distinguish between ν_{μ} and $\bar{\nu}_{\mu}$, their measurement was only an indirect test of *CPT* conservation. This result was consistent with no violation of *CPT* conservation. The previous MINOS measurement was made using events in which the neutrino interaction vertex was contained within the detector fiducial mass. This test also showed no evidence for *CPT* violation in atmospheric neutrinos.

After briefly discussing the detector in Sec. II, we describe our event sample in Sec. III A and the Monte Carlo event generation in Sec. III B. In Sec. III C, we describe the cuts used in selecting the sample of muons to be analyzed. We present the oscillation analysis in Sec. IV and the probe of *CPT* conservation using charge-separated neutrino-induced μ^{-} and μ^{+} in Sec. V.

II. THE MINOS FAR DETECTOR

The MINOS far detector is a steel-scintillator sampling calorimeter located at a depth of 2070 meters-water-equivalent (m.w.e.) in the Soudan mine in northern Minnesota [18]. The detector is made up of 486 vertical octagonal planes of 2.54 cm thick steel laminates, interleaved with 484 planes of 1 cm thick extruded polystyrene scintillator strips and a 2.5 cm wide air gap. Each scintillator plane has 192 strips of width 4.1 cm. The length of each strip depends on its position in the plane and varies between 3.4 and 8.0 m. The scintillator strips in alternating detector planes are oriented at $\pm 45^{\circ}$ to the vertical. The modular detector consists of two supermodules (SM) separated by a gap of 1.1 m. The detector coordinate system has the y axis pointing up, the z axis along the detector long axis (pointing away from Fermilab), and the x axis forms a right-handed coordinate system.

Scintillation light from charged particles traversing the MINOS plastic scintillator is collected with wavelength shifting (WLS) plastic fibers embedded in the scintillator strips. The WLS fibers are coupled to clear optical fibers at both ends of a strip and are read out using 16-pixel multi-anode photomultiplier tubes (PMTs). The signals from eight strips, separated by approximately 1 m within the same plane, are optically summed, or multiplexed, and read out by a single PMT pixel. The multiplexing pattern is different for the two sides of the detector, enabling the resulting eightfold ambiguity to be resolved for single particles. For all other types of events, ambiguities are resolved effectively using additional information from timing and event topology.

To measure the momentum of muons traversing the detector, the steel has been magnetized into a toroidal field configuration using a current-carrying coil that runs through the central axis of each SM. A finite element analysis calculation shows that each SM is magnetized to an average value of ~ 1.3 T by the 15 kA current loop. The field is saturated near the coil hole at a strength of approximately 1.8 T, falling to about 1 T near the edges. Further details about the MINOS detector can be found in [18].

To distinguish upward-going neutrino-induced muons from the downward cosmic ray background requires excellent timing resolution. The smallest increment of time that the MINOS timing system is able to measure, or its least count, is 1.56 ns. However, the timing resolution is dominated by the 8 ns fluorescence decay time in the WLS fiber [25]. The overall resolution is determined by the convolution of the least count with the fluctuations in the arrival times of the photons at the PMT. We calibrated the timing system by measuring the time offsets between each channel. For this determination [26,27], we selected a sample of downward through-going cosmic ray muons with well-reconstructed tracks. For each hit along a track we measured the travel times from the track entrance point in the detector to the hit location and corrected those values for both the rise time of the signal and the propagation of light along the fibers. We then compared these times with the time expected for a relativistic muon traversing the same distance and computed the offset between the measured and expected times for each hit, Δt . The timing calibration constants are obtained using an iterative method to minimize the mean Δt for each channel. After calibration, linear timing fits are again applied to the times and positions of the hits on each muon track. The root mean square (r.m.s.) deviation between the measured and fitted times is calculated for each track. Figure 1 shows the

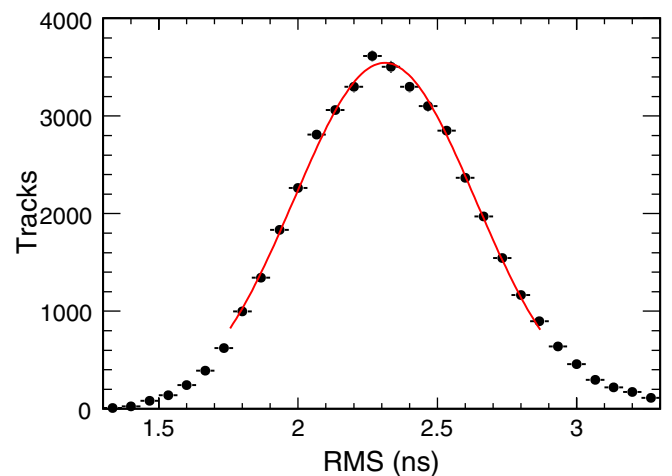


FIG. 1 (color online). Distribution of the r.m.s. deviations of the measured hit times from the calculated hit times for calibration tracks. The resolution of the system is the mean of a Gaussian fit to the distribution, 2.31 ± 0.03 ns.

distribution of these r.m.s. deviations. In this figure, the distribution has been fit to a Gaussian with a mean of 2.31 ± 0.03 ns, a value that represents the overall timing resolution of the detector.

III. DATA ANALYSIS

In the analyses presented here we use neutrino-induced muons, which are defined as events that come from below or slightly above the horizon. These events are essentially uncontaminated by the background of downward-going atmospheric muons. We analyze two types of events: those that pass completely through the detector (“through-going muons”) and those that enter and stop in the detector (“stopping muons”). Earlier MINOS results for events with an interaction vertex in the detector (“contained events”) appear in [17]. We distinguish between neutrino-induced muons with well-determined momenta and neutrino-induced muons whose charge and momentum are undetermined (cf. Sec. III D).

A. Data sample

The data presented here represent 854.24 live days and were taken between August 1, 2003 and April 30, 2006. The geometric acceptance of the detector for neutrino-induced muons is 6.91×10^6 cm² sr. We found a total of 140 neutrino-induced muons in this data set.

B. Simulated muons

1. Atmospheric neutrino-induced muons

We generated a large sample of simulated neutrino-induced muons in the MINOS far detector using the Bartol 96 [28] neutrino flux tables along with the default NUANCE neutrino interaction model [29] and the GRV94 [30] parton distribution functions. We first simulated neutrino-induced muons passing through the surface of a box surrounding the MINOS far detector and then propagated these muons through the detector with the MINOS GEANT-based detector Monte Carlo simulation. A total of 6.5×10^5 neutrino-induced muon events, the equivalent of 2500 years of live time, were generated in this way. The momentum distributions simulated for the neutrino-induced μ^- and μ^+ in MINOS are shown in Fig. 2 for neutrinos without oscillations and for neutrinos with an oscillation signal using the Super-Kamiokande parameters [8]. This figure shows that the oscillation signal affects mostly muons with low momenta.

2. Cosmic ray muons

We simulated each cosmic ray muon event by first choosing its arrival direction from a distribution uniform in solid angle and then associating this direction with the overburden found in the Soudan 2 slant depth map [31]. The surface energy of the muon was selected from the known distribution [32]. Once the surface energy and over-

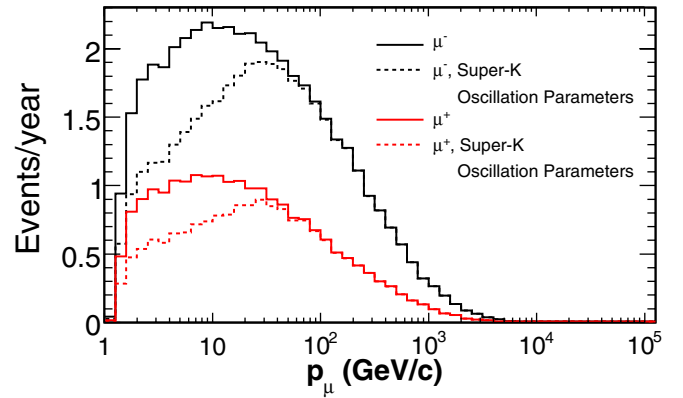


FIG. 2 (color online). Input momentum distribution for neutrino-induced μ^- and μ^+ calculated with the NUANCE simulation package. The distributions in the MINOS far detector without oscillations and with the Super-Kamiokande determined values of $\sin^2 2\theta_{23}$ and Δm_{23}^2 are shown. The oscillation affects mostly muons with low momenta.

burden were chosen, we tested the muon to see whether it survived energy loss processes between the surface and the detector [32]. We placed the muons that survived on an imaginary box surrounding the detector [33] and then propagated them through the detector with the MINOS GEANT-based detector simulation.

C. Event selection

All data and simulated muons were analyzed using the standard MINOS reconstruction algorithms and a uniform set of event selection cuts. The selection cuts summarized below are described in further detail in [26].

1. Muon selection cuts

We first selected muon events using criteria developed for the study of cosmic ray muons with the MINOS far detector [34]. The first two cuts require that there was a track successfully reconstructed in the event (“No track” cut), and that there is only a single track found by the track-fitting algorithm (“Multiples” cut). The next set of cuts excludes random collections of hits that could be mistaken for muon tracks. These cuts require that (1) the track cross at least 20 planes in the detector (“20 plane” cut), (2) the track have a path length of at least 2 m (“2.0 m length” cut), (3) the earliest recorded hit of the track be no more than 15 cm from the front or back surface of the detector or 30 cm from any of the remaining surfaces (“Fiducial” cut), and (4) the track fit have a $\chi^2_{\text{fit}}/\text{ndf} < 1.5$.

The next selection cuts were designed to remove events with poor timing information. To ensure the presence of sufficient timing information in the events, we first excluded tracks if fewer than half of the hits come from strips with signals on both ends (“Double-ended strip cut”). We determined the direction of the track by plotting the time

difference ΔT (ns) of each hit along the track as a function of its distance ΔS (m) from the first hit. If the y positions of the hits increase along the length of the track, ΔS is positive; for y decreasing along the track, ΔS is negative. The slope of the linear fit to the $\Delta T/\Delta S$ distribution is $1/\beta = c/v$. The top panel of Fig. 3 shows the $\Delta T/\Delta S$ distribution with the linear fit superposed for a typical cosmic muon; the bottom panel displays the (x, y) positions of the hits in the detector for this muon.

In Fig. 4 we show the distributions of $\chi^2_{1/\beta}/\text{ndf}$ values for the linear fits to $1/\beta$ for both data and cosmic ray muon Monte Carlo events passing the above cuts. The distribution does not peak at 1 as might be expected for a χ^2/ndf statistic. However, that is unimportant to the analysis since we are using the $\chi^2_{1/\beta}/\text{ndf}$ value only to define a cut. We defined a cut that requires $\chi^2_{1/\beta}/\text{ndf} < 3.0$ for an event (“ $\chi^2_{1/\beta}/\text{ndf} < 3.0$ ” cut). The cut was selected at the value of $\chi^2_{1/\beta}/\text{ndf}$ where the number of events falls to $\sim 1\%$ of the peak. This cut maximizes the number of selected events while excluding those events with the worst $\chi^2_{1/\beta}/\text{ndf}$

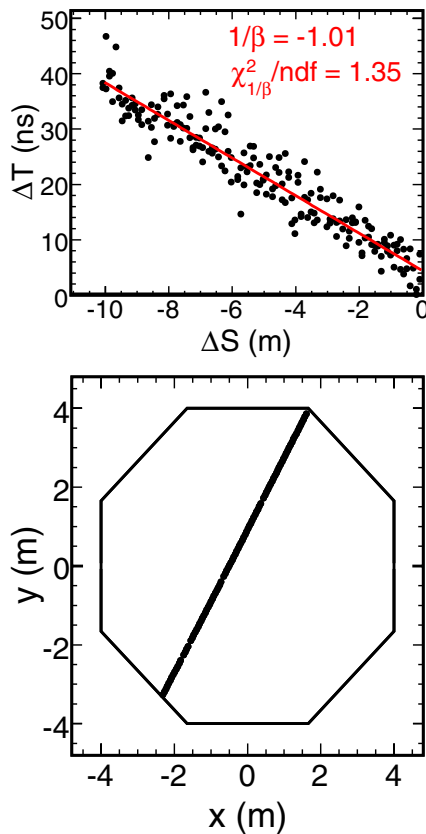


FIG. 3 (color online). A typical cosmic ray muon in the MINOS far detector. The top panel shows the timing information for the hits along the track with a straight line fit superposed. The legend gives $\Delta T/\Delta S = 1/\beta$ and the $\chi^2_{1/\beta}/\text{ndf}$ of the fit for this muon. The bottom panel shows the (x, y) hit positions of this track. The resolution of the positions is the 4 cm width of the strips.

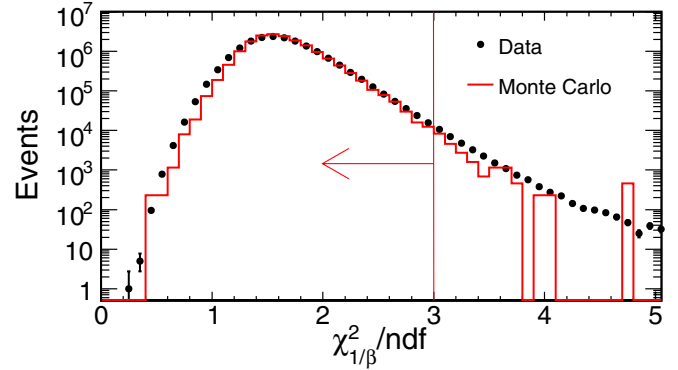


FIG. 4 (color online). Distribution of $\chi^2_{1/\beta}/\text{ndf}$ values from the $1/\beta$ determinations for data and cosmic ray muon Monte Carlo events. In this figure, the Monte Carlo distribution has been normalized to the same area as the data distribution. Events with $\chi^2_{1/\beta}/\text{ndf} < 3$ are used in the analysis.

values. Although the data and Monte Carlo simulation deviate at low values of $\chi^2_{1/\beta}/\text{ndf}$, the agreement is excellent near the peak and the cut value. Therefore, the Monte Carlo sample can be used to effectively study the systematic uncertainty introduced into the analysis by this cut.

In the last step of the muon selection, we separated the downward-going from upward-going events. Upward-going events have a positive slope for the straight line fit to the $\Delta T/\Delta S$ distribution. The final muon cut is a check on the up/down separation. The entrance point of a track, and therefore its incoming direction, is determined by the slope of the times of the hits in the detector as a function of their z position. For a few short tracks this information is sufficiently ambiguous that the entrance point can be confused with the exit point of the track causing the reconstruction to interchange the two. As a result, downward-going muons are incorrectly reconstructed as upward-going. We remove these failures by checking the times of the hits in the reconstructed track as a function of their y position. The slope of the line fit to the hits as a function of their y positions must agree with the reconstructed incoming direction of the track (“Directionality” cut).

Table I shows the effect of the cuts on the data and the two Monte Carlo distributions. We normalize the cosmic ray muon Monte Carlo distributions to the data value at the 20 plane cut because the Monte Carlo simulation does not account for multiple muon events which make up $\sim 5\%$ of downward-going cosmic ray muons [35]. After cuts the selected events consist of muon tracks whose directionality is well determined. The distribution of $1/\beta$ values for the selected muon events is shown in Fig. 5. Since the sample includes both downward-going and upward-going muons and these muons are relativistic, there are two peaks in the $1/\beta$ distribution. The peak at $1/\beta = -1$ shows downward-going muons and the peak at $1/\beta = 1$ shows upward-going muons. Integrating the events in the two

TABLE I. Effect of cuts on data and simulated samples. The fraction of the total events remaining is shown. N gives the total number of events in each sample.

Cut	Data $N = 3.81 \times 10^7$	μ MC $N = 1.20 \times 10^6$	ν -induced MC $N = 6.47 \times 10^5$
No cuts	1.000	...	1.000
No track	0.800	...	0.862
Multiples	0.743	...	0.848
20 plane	0.561	0.561	0.606
2.0 m length	0.557	0.557	0.578
Fiducial	0.534	0.538	0.559
$\chi^2_{fit}/ndf < 1.5$	0.429	0.447	0.497
Double-ended strips	0.429	0.447	0.497
$\chi^2_{1/\beta}/ndf < 3.0$	0.428	0.447	0.497
Directionality	0.428	0.447	0.478

peaks shows that the fraction of (upward-going muons)/(downward-going muons) $\sim 10^{-5}$.

2. Upward-going neutrino-induced muons

As in previous experiments [4,16], we identify upward-going muons that enter the detector as neutrino-induced muons. Events in the range $0.7 < 1/\beta < 1.3$ are included in our neutrino-induced muon data set. The cut was determined using downward-going cosmic ray muons. It was set as the $1/\beta$ range that includes 99% of the events and which is centered on the peak. In the left panel of Fig. 6 we show a well-characterized upward-going muon event. The right panel shows an event excluded by the $1/\beta$ cut.

Table II shows that our data sample includes 130 upward-going neutrino-induced muons in the range $0.7 < 1/\beta < 1.3$. In this table we also show the background from contained vertex interactions where the neutrino interacts close to the detector edge and cannot be separated from muons entering the detector. This background was estimated by calculating the rate of Monte Carlo atmospheric

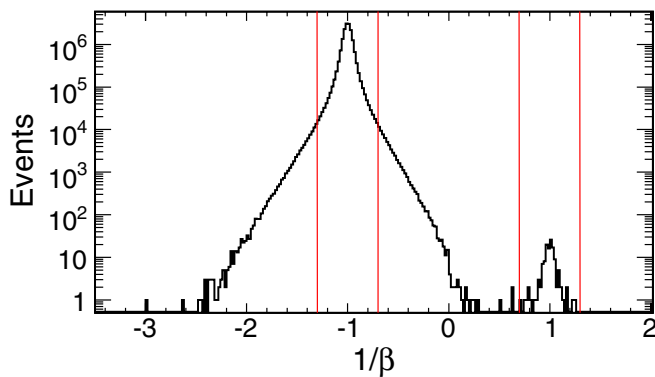


FIG. 5 (color online). Distribution of $1/\beta$ for upward-going neutrino-induced muons, with a peak at 1, and downward-going cosmic ray muons, with a peak at -1 . The vertical lines at 0.7 and 1.3 bracket the events included in the upward-going muon sample.

neutrino events that interact inside the detector volume [17]. The background includes events that pass the cuts in Table I and satisfy the upward-going selection cut. The Monte Carlo events were assumed to oscillate with the best-fit oscillation parameters from Super-Kamiokande [8]. The 20 plane cut has been shown to be effective at removing the background due to upward-going pions produced by downward-going cosmic ray muons interacting in the rock surrounding the detector.

3. Horizontal muons

In this analysis, we also include muons coming from along and slightly above the detector's horizon. The flat overburden of the Soudan site [31] makes this search feasible. The slant depth of rock between the detector and the surface for incoming directions above the horizon increases approximately as $\sec\theta$, where θ is the zenith angle. Only the highest energy cosmic ray muons ($E_\mu > 100$ TeV) have sufficient energy to penetrate the large column of rock present for muons coming from directions near the horizon. Since the intensity of cosmic ray muons falls as $E^{-2.7}$, very few cosmic ray muons survive to reach the detector, implying that muons from near the detector's

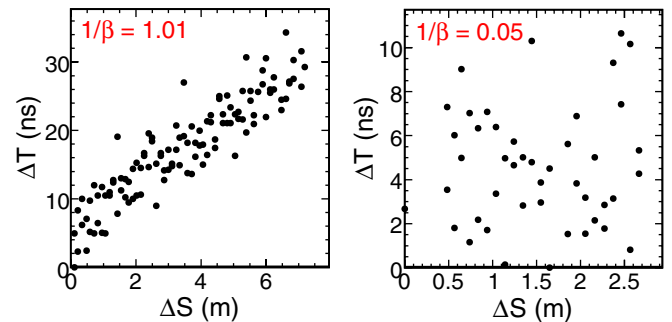


FIG. 6 (color online). Distributions of $\Delta T/\Delta S$ as in Fig. 3. The left panel shows a typical upward-going muon and the right panel shows a muon excluded by the $1/\beta$ cut from the upward-going sample.

TABLE II. Observed upward-going and horizontal muons. The background accounts for both contained vertex neutrino interactions and downward-going cosmic ray muons.

Muon type	Events	Background	
		Contained	Cosmic μ
Upward-going	130	4.2	0.0
Horizontal	10	0.1	0.3

horizon are neutrino induced. These events are important to the analysis because they sample neutrino-induced muons with lower values of L/E than the upward-going muons. Equation (1) shows that the average oscillation probability for these horizontal muons is lower than for upward-going neutrino-induced muons. Thus the horizontal muons add information important in determining the neutrino flux normalization, a parameter used in the oscillation analysis in Sec. IV C.

We require that the measured velocity of the downward-going muons be in the range $-1.3 < 1/\beta < -0.7$. To determine the maximum angle above the horizon from which we can distinguish horizontal muons from cosmic ray muons, we use the zenith angle distribution for all events shown in Fig. 7. The distribution is steeply falling for $\cos\theta < 0.25$ but becomes approximately constant for $\cos\theta < 0.1$. The rate of neutrino-induced muons is expected to be approximately independent of $\cos\theta$ [32]. We select events with $\cos\theta < 0.05$ to minimize the background from cosmic ray muons. To estimate this background we fit an exponential to the distribution of Fig. 7 for $0.1 \leq \cos\theta \leq 0.2$, a similar procedure to that used by Super-Kamiokande [4]. The exponential fit is shown in the figure; it has the form $y = \exp(a + b \cos\theta)$, where $a = -3.59$ and $b = 51.65$. Figure 7 shows that the background at

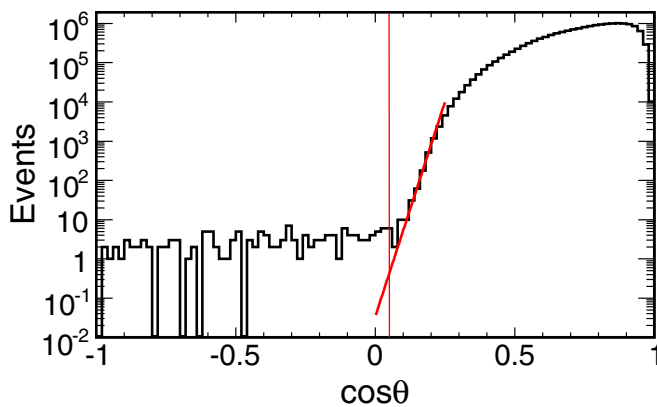


FIG. 7 (color online). The $\cos\theta$ distribution for all incoming muons. The data are fit to an exponential between $0.1 \leq \cos\theta \leq 0.2$ to estimate the background of cosmic ray muons in the horizontal neutrino-induced muon signal region. The fit is shown in the figure. As indicated, we select muons coming from directions with $\cos\theta < 0.05$.

$\cos\theta = 0.05$ is 0.3 events and is negligible for smaller values of $\cos\theta$.

In Table II we show that there are 10 horizontal neutrino-induced muons in the data sample. The background shown includes both cosmic ray downward-going muons and the contained vertex events discussed previously.

D. Charge-sign and momentum determination

Previous oscillation analyses based on neutrino-induced muons have typically divided the event sample into through-going and stopping muons. The experiments that presented these analyses could only measure the muon momentum directly for stopping muons. Since the MINOS far detector is magnetized, we can measure the muon momentum and charge sign for both types of neutrino-induced muons.

To determine whether the charge and momentum of a neutrino-induced muon have been determined accurately, we have developed a criterion based on the measured curvature of the muon track. We first draw a straight line between the end points on the track and then use the deviations of the remaining track points from the line to form a χ^2 statistic, $\chi_{\text{line}}^2/\text{ndf}$. Large values of $\chi_{\text{line}}^2/\text{ndf}$ indicate significant bending in the magnetic field which leads to a good determination of the charge sign and momentum.

We use our Monte Carlo simulation to set the value of $\chi_{\text{line}}^2/\text{ndf}$ that selects neutrino-induced muons with good momentum and charge-sign determination. As a test of whether the $\chi_{\text{line}}^2/\text{ndf}$ values are well simulated, we compare in Fig. 8 the distributions of $\chi_{\text{line}}^2/\text{ndf}$ for stopping muons in the cosmic ray data, in the cosmic ray muon Monte Carlo simulation, and in the neutrino-induced muon

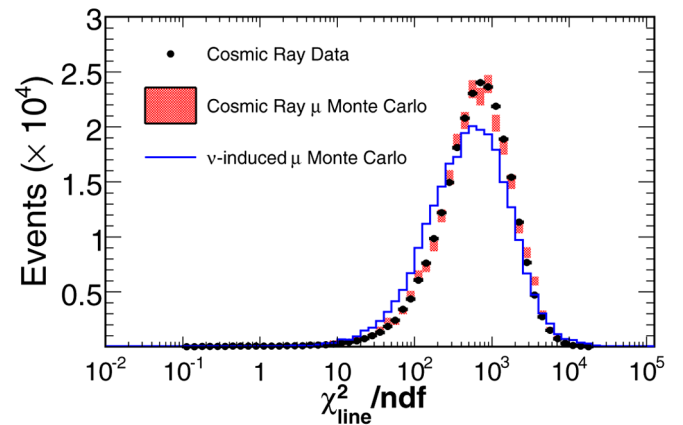


FIG. 8 (color online). Comparison of the $\chi_{\text{line}}^2/\text{ndf}$ distributions for stopping muons in the cosmic ray data (points), in the cosmic ray Monte Carlo simulation (boxes), and in the neutrino-induced Monte Carlo simulation (line). The number of events in each bin for the cosmic ray Monte Carlo simulation is shown by the location of the center of the box for the bin, and the vertical extent of the box indicates the statistical uncertainty in the bin.

Monte Carlo simulation. The Monte Carlo distributions have been normalized to have the same number of events as in the data. Figure 8 shows that the cosmic ray muon Monte Carlo simulation provides a high statistics match to the cosmic ray data. Since the stopping cosmic ray muons have energies <20 GeV, the maximum energy of a normally incident stopping muon, these low energy cosmic events provide a sample of muons with a distribution in energy similar to neutrino-induced muons. Also superposed on these distributions is that of the Monte Carlo neutrino-induced muons. This figure suggests that the neutrino-induced Monte Carlo simulation can be used to determine an appropriate value of $\chi_{\text{line}}^2/\text{ndf}$ for selecting events with good charge-sign and momentum determination.

In Fig. 9 we show the fraction of Monte Carlo neutrino-induced muons with correct charge identification, or purity, as a function of $\chi_{\text{line}}^2/\text{ndf}$. The purity is approximately 97% at $\chi_{\text{line}}^2/\text{ndf} = 10$ and rises to over 99% with increasing values of $\chi_{\text{line}}^2/\text{ndf}$. We identify muons with $\chi_{\text{line}}^2/\text{ndf} > 10$ as having good charge-sign and momentum determination; muons with $\chi_{\text{line}}^2/\text{ndf} < 10$ are assumed to have poor charge-sign and momentum determination.

We further divided the Monte Carlo muons with $\chi_{\text{line}}^2/\text{ndf} > 10$ into two samples based on their momentum, a separation that roughly distinguishes muons whose parent neutrinos have a relatively large probability of oscillation from those with a lower probability. The “low momentum” (*L*) sample includes muons with fit momentum, p_{fit} , in the range $1 \leq p_{\text{fit}} < 10$ GeV/*c*. The “high momentum” (*H*) sample includes muons in the range $10 \leq p_{\text{fit}} < 100$ GeV/*c*. No muons with $\chi_{\text{line}}^2/\text{ndf} > 10$ have $p_{\text{fit}} > 100$ GeV/*c*. The muons with $\chi_{\text{line}}^2/\text{ndf} < 10$ are in the “unknown momentum” (*U*) sample.

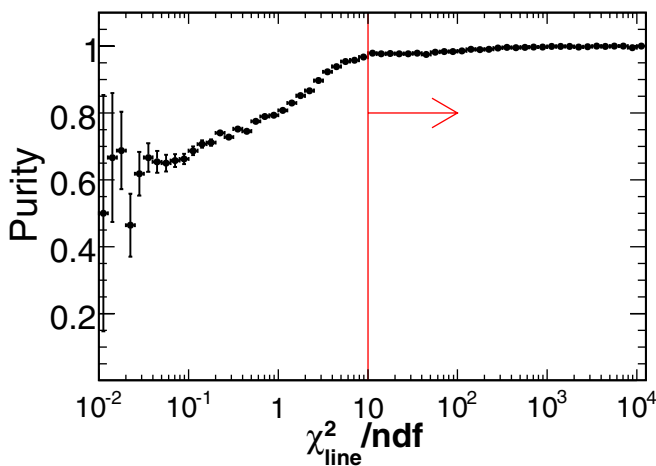


FIG. 9 (color online). Purity of the charge-sign determination for neutrino-induced muons as a function of $\chi_{\text{line}}^2/\text{ndf}$. Events with $\chi_{\text{line}}^2/\text{ndf} > 10$ have well-determined charge sign and momentum.

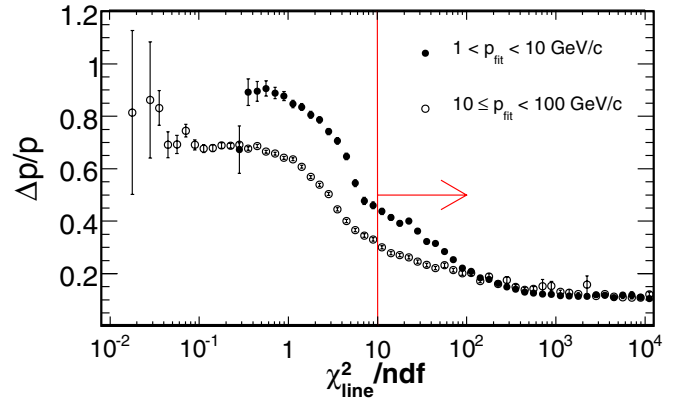


FIG. 10 (color online). Momentum resolution for Monte Carlo neutrino-induced muons as a function of $\chi_{\text{line}}^2/\text{ndf}$. The line at $\chi_{\text{line}}^2/\text{ndf} = 10$ shows the cut used in the analyses.

The quality of the momentum determination can be deduced from the Monte Carlo simulation using the r.m.s. momentum resolution, $(\Delta p/p) = \sqrt{\langle (p_{\text{truth}} - p_{\text{fit}})^2 / p_{\text{truth}}^2 \rangle}$, where p_{truth} is the known momentum of the muons. In Fig. 10 we show $\Delta p/p$, as a function of $\chi_{\text{line}}^2/\text{ndf}$ for the low momentum and high momentum samples. The low momentum Monte Carlo muons have a momentum resolution of less than 50% at $\chi_{\text{line}}^2/\text{ndf} = 10$ and the resolution decreases to 10% with increasing values of $\chi_{\text{line}}^2/\text{ndf}$. For the high momentum muons the momentum resolution is less than 30% at $\chi_{\text{line}}^2/\text{ndf} = 10$ and it decreases to 10% as $\chi_{\text{line}}^2/\text{ndf}$ increases.

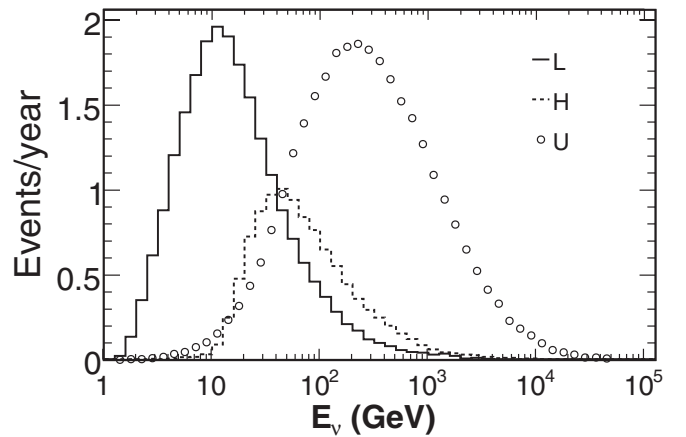


FIG. 11. Distribution of energies for neutrinos producing neutrino-induced muons observed in the MINOS detector as determined by the Monte Carlo simulation. The neutrinos producing low momentum muons are shown by the solid line, those producing high momentum muons are shown by the dashed line, and those producing unknown momentum muons are shown by open circles. Neutrinos producing muons detected by MINOS have energies ≥ 2 GeV.

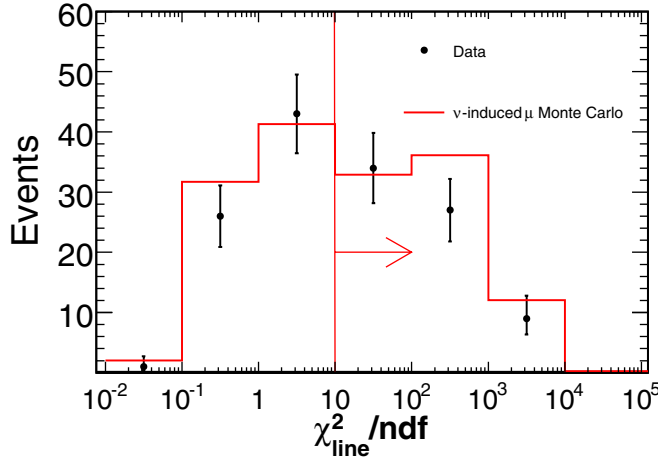


FIG. 12 (color online). Comparison of the $\chi^2_{\text{line}}/\text{ndf}$ distributions for neutrino-induced muon data and unoscillated Monte Carlo simulation.

Figure 11 shows the distributions of parent neutrino energies for the low momentum, high momentum, and unknown momentum muons as determined by the Monte Carlo simulation. The neutrinos producing low momentum muons have energies that peak near 10 GeV. The high momentum muons are produced by neutrinos with energies that peak near 50 GeV. The unknown momentum muons are produced by even higher energy neutrinos, with a peak energy near 250 GeV. Comparing the peak energies of the various neutrino distributions, we see from Eq. (1) that the low momentum muons are those that are expected to show the strongest oscillation signal, a result also seen in Fig. 2.

In Fig. 12 we show the distributions of $\chi^2_{\text{line}}/\text{ndf}$ for the selected neutrino-induced muons from the data and unoscillated Monte Carlo simulation normalized by live time. The muons with $\chi^2_{\text{line}}/\text{ndf} < 10$ are from high energy neutrinos and are not expected to show a strong oscillation signal. Those muons with $\chi^2_{\text{line}}/\text{ndf} > 10$ are from lower

TABLE III. Momentum and charge sign of selected neutrino-induced muons. The calculated background is shown as well as the Monte Carlo expectation in the absence of oscillations.

p_{fit} (GeV)	Data	Background	MC
		μ^-	
1–10 (<i>L</i>)	21	2.2	37.5
10–100 (<i>H</i>)	20	0.2	17.5
		μ^+	
1–10 (<i>L</i>)	16	1.3	19.3
10–100 (<i>H</i>)	13	0.2	8.6
		<i>U</i>	
Unknown (<i>U</i>)	70	0.7	76.5

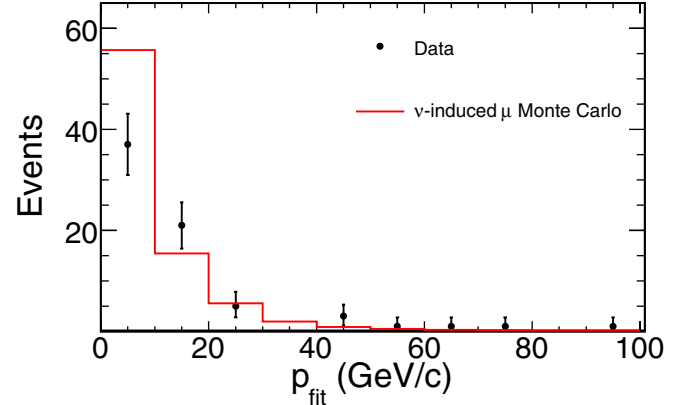


FIG. 13 (color online). Distribution of fit momenta for events in the combined low momentum and high momentum data samples. The Monte Carlo expectation for no oscillations is shown by the solid line. The unknown momentum muons are not included in this figure.

energy neutrinos, a sample expected to show an oscillation signal.

The results of applying the $\chi^2_{\text{line}}/\text{ndf}$ cut to the sample of 140 neutrino-induced muons are given in Table III. In this table, the events with $\chi^2_{\text{line}}/\text{ndf} > 10$ have been further separated by charge identification and momentum. The calculated background contribution is shown as well as the Monte Carlo expectation in the absence of oscillations. Figure 13 shows the fit momentum distributions for data and unoscillated Monte Carlo simulation using the combined low momentum and high momentum muon samples. The unknown momentum muons are not included in this figure. The first bin in Fig. 13 shows a deficit of events in the data relative to Monte Carlo expectation without oscillations (cf. Fig. 2).

IV. OSCILLATION ANALYSIS

After first discussing the systematic uncertainties associated with the analysis procedure, we test the neutrino-induced muons in Table III for evidence of neutrino oscillations.

A. Systematic uncertainties

There are several sources of systematic uncertainty in this analysis. These are due to both the event reconstruction and the physics modeling. These uncertainties are summarized in Table IV.

The sources of systematic uncertainty related to reconstruction are those associated with the data selection cuts (1) $\chi^2_{\text{fit}}/\text{ndf}$, (2) $\chi^2_{1/\beta}/\text{ndf}$, and (3) $\chi^2_{\text{line}}/\text{ndf}$, where the numbers refer to Table IV. The systematic uncertainties for these cuts were all computed in a similar manner. For example, we computed the systematic uncertainty on $\chi^2_{1/\beta}/\text{ndf}$ by establishing that the cut value of $\chi^2_{1/\beta}/\text{ndf} = 3$ selects 98.9% of events seen in Fig. 4. We divided the

TABLE IV. Sources of systematic uncertainty in the low-to-high momentum event ratio \mathcal{R} .

Source	σ	$\Delta\mathcal{R}_k$
Reconstruction systematics:		
(1) $\chi_{\text{fit}}^2/\text{ndf} < 1.5$	0.01	$< 5 \times 10^{-4}$
(2) $\chi_{1/\beta}^2/\text{ndf} < 3.0$	0.01	$< 1 \times 10^{-4}$
(3) $\chi_{\text{line}}^2/\text{ndf} < 10$	0.27	0.02
Model systematics:		
(4) Normalization	0.15	$< 1 \times 10^{-4}$
(5) Spectral index	0.03	0.08
(6) Cross section ($E < 30$ GeV)	0.07	0.03
(7) Cross section ($E > 30$ GeV)	0.02	0.01
$\Delta\mathcal{R} = \sqrt{\sum(\Delta\mathcal{R}_k)^2}$		0.09

total Monte Carlo data set into 12 subsamples of 20 000 events each. For each subsample we found the value of $\chi_{1/\beta}^2/\text{ndf}$ that accepts 98.9% of the events. The variance in these values of $\chi_{1/\beta}^2/\text{ndf}$ divided by the nominal value of the $\chi_{1/\beta}^2/\text{ndf}$ cut is the 1σ uncertainty given in Table IV. The 1σ values for the uncertainties in $\chi_{\text{fit}}^2/\text{ndf}$ and $\chi_{\text{line}}^2/\text{ndf}$ were computed in a similar manner.

The first source of systematic uncertainty in the physics modeling is the (4) overall normalization of the calculated neutrino flux. Uncertainties in the primary cosmic ray fluxes and the hadronic production are the main contributors to the overall uncertainty in the normalization. Combining these sources gives an uncertainty in the flux normalization of 15% [36]. There are three sources of systematic uncertainty in the ratio of the number of muons induced by low energy neutrinos to those induced by high energy neutrinos. One contribution comes from (5) the uncertainty in the spectral index of the atmospheric neutrino energy spectrum. The neutrino flux is proportional to $E_\nu^{-(\gamma+1)}$, with the value of the spectral index, $\gamma = 1.7 \pm 0.05$ [35], a 3% uncertainty in the spectral index. The uncertainties in the cross sections for neutrinos and antineutrinos are energy dependent, with much larger uncertainties for neutrinos with energies < 30 GeV than for those with energies > 30 GeV [10]. This difference in the uncertainties leads to a natural breakpoint in the specification of systematic uncertainties in the cross sections for the parent neutrino interactions. With this division the uncertainties in the neutrino and antineutrino cross sections contribute 7% [37] to the uncertainty in the rate of muons coming from neutrinos with energies < 30 GeV (6) and 2% [10] for muons coming from neutrinos with energies > 30 GeV (7).

B. Low-to-high and unknown momentum event ratio

One way to look for evidence of neutrino oscillations in the neutrino-induced muons is to take the ratio of the number of low momentum muons, which are more likely

to show an oscillation signal, to the sum of the number of high momentum and unknown momentum muons, which are less likely, and compare this ratio with its Monte Carlo expectation including backgrounds. In the data, this ratio is given by

$$R_{L/H+U}^{\text{data}} = \sum_L (N_{\mu^-} + N_{\mu^+}) / \sum_{H+U} (N_{\mu^-} + N_{\mu^+}), \quad (2)$$

where N_{μ^-} is the number of μ^- observed in a bin and N_{μ^+} is the number of μ^+ observed in a bin. The sum over L includes events in the range $1 < p_{\text{fit}} < 10$ GeV/ c , and the sum over $(H + U)$ includes the remaining high momentum and unknown momentum events. In the Monte Carlo simulation, a similar ratio, $R_{L/H+U}^{\text{MC}}$, is defined. In the absence of oscillations, the ratio of these two quantities, \mathcal{R} , will be consistent with unity; if an oscillation signal is present, \mathcal{R} will be less than unity.

We computed the systematic uncertainty in this ratio with our Monte Carlo simulation by varying the value of each of the parameters in Table IV by $\pm 1\sigma$ from the values used in the analysis. The change in the Monte Carlo ratio resulting from the variation in each parameter, $\Delta\mathcal{R}_k$, was assumed to be the uncertainty in the ratio. The total uncertainty in $\Delta\mathcal{R}$ was found by adding the uncertainties from the individual parameters in quadrature.

As the reconstruction uncertainties are based on how different the cut efficiencies are for data and Monte Carlo simulation, we determined how the value of those selection criteria affected the ratio by varying the value of the cuts. Varying $\chi_{\text{fit}}^2/\text{ndf}$ between 1.485 and 1.515 gives $\Delta\mathcal{R}_1 < 5 \times 10^{-4}$. The value $\Delta\mathcal{R}_2 < 1 \times 10^{-4}$ is as expected from Fig. 4. Changing the value of $\chi_{\text{line}}^2/\text{ndf}$ between 7.3 and 12.7 shows that $\Delta\mathcal{R}_3 = \pm 0.02$. The flux normalization simply scales the number of low momentum, high momentum, and unknown momentum muons by a constant so this uncertainty cancels in the ratio. Varying the spectral index by $\pm 3\%$ alters the relative numbers of low momentum to high momentum and unknown momentum muons which leads to $\Delta\mathcal{R}_5 = \pm 0.08$. Varying the cross section for neutrinos with $E < 30$ GeV by $\pm 7\%$ gives $\Delta\mathcal{R}_6 = \pm 0.03$. A variation of $\pm 2\%$ in the cross section for neutrino interactions with $E > 30$ GeV gives $\Delta\mathcal{R}_7 = \pm 0.01$. Adding these uncertainties in quadrature gives the total systematic uncertainty, $\Delta\mathcal{R} = 0.09$. These results are given in Table IV.

From Tables III and IV we find

$$\mathcal{R} = \frac{R_{L/H+U}^{\text{data}}}{R_{L/H+U}^{\text{MC}}} = 0.65_{-0.12}^{+0.15}(\text{stat}) \pm 0.09(\text{syst}). \quad (3)$$

The upper and lower limits on the data event rate ratio are estimated accounting for the Poisson fluctuations in the numerator and denominator [38]. The ranges quoted are calculated to give coverage at 68% C.L. Adding the upper statistical and systematic uncertainties in quadrature, the upper uncertainty is +0.17 which results in a value for \mathcal{R}

that differs from the no oscillation expectation of unity by 2.0σ . This result is consistent with neutrino oscillations.

C. Oscillation fit

In the following section we test the significance of the neutrino disappearance suggested by the value of \mathcal{R} in Eq. (3) by fitting for the oscillation parameters $\sin^2 2\theta_{23}$ and Δm_{23}^2 .

1. Method

We used the data and Monte Carlo samples from Table III for the oscillation fit. We first divided the momentum-separated muons (L, H, U) into three $\cos\theta$ bins: $-1.0 < \cos\theta < -0.6$, $-0.6 < \cos\theta < -0.2$, and $-0.2 < \cos\theta < 0.05$. To find the oscillation parameters ($\sin^2 2\theta_{23}$, Δm_{23}^2), we compared the data and Monte Carlo simulation with the χ^2 statistic,

$$\chi^2 = 2 \sum_{i=1}^9 \left(N_i^{\text{MC}} - N_i^{\text{D}} + N_i^{\text{D}} \ln \frac{N_i^{\text{D}}}{N_i^{\text{MC}}} \right) + \sum_{k=1}^7 \frac{\epsilon_k^2}{\sigma_k^2}. \quad (4)$$

The first sum is the χ^2 for Poisson distributed data [39]. The number of data events seen in bin i is N_i^{D} and the number of events predicted in bin i by the Monte Carlo simulation is N_i^{MC} . The parameters ϵ_k in the second sum account for the systematic uncertainties described in Sec. IVA. The ϵ_k are included in the fit as penalty terms and allow the fit to adjust the predicted number of muons in each bin i based on the systematic uncertainties. The ϵ_k are assumed to be Gaussian distributed with a mean of 0 so that excursions of the fit values from the expected values contribute an amount ϵ_k^2/σ_k^2 to the χ^2 statistic, where the σ_k are given in Table IV.

The N_i^{MC} depend on the effects of neutrino oscillations, the systematic uncertainties due to the reconstruction and input physics model, and the background present for bin i . We define the expected number of muons in bin i as $N_i^{\text{MC}} = N_i^{\text{MC}}(\sin^2 2\theta_{23}, \Delta m_{23}^2, \vec{\epsilon}) + N_i^{\text{BG}}$. The background events due to neutrino interactions inside the detector, N_i^{BG} , have been oscillated with the best-fit Super-Kamiokande parameters and $\vec{\epsilon} = (\epsilon_1, \dots, \epsilon_7)$ correspond to the uncertainties listed in Table IV. We calculated the values of χ^2 at each point on a grid in $(\sin^2 2\theta_{23}, \Delta m_{23}^2)$ space. The fit varied the ϵ_k at each point to find the minimum χ^2 [40] in Eq. (4) [8].

2. Results of oscillation fit

The best fit in the physical region is found at $\sin^2 2\theta_{23} = 1$ and $\Delta m_{23}^2 = 0.93 \times 10^{-3} \text{ eV}^2$ with $\chi^2/\text{ndf} = 5.9/7$. The values of the ϵ_k were all found to be less than $1\sigma_k$. The results of the fit are shown in Fig. 14, where the best fit to the data in the physical region (solid line), the Monte Carlo prediction for the null oscillation hypothesis (dotted line), and the prediction derived from the MINOS result with the NuMI neutrino beam (dashed line) [19] are

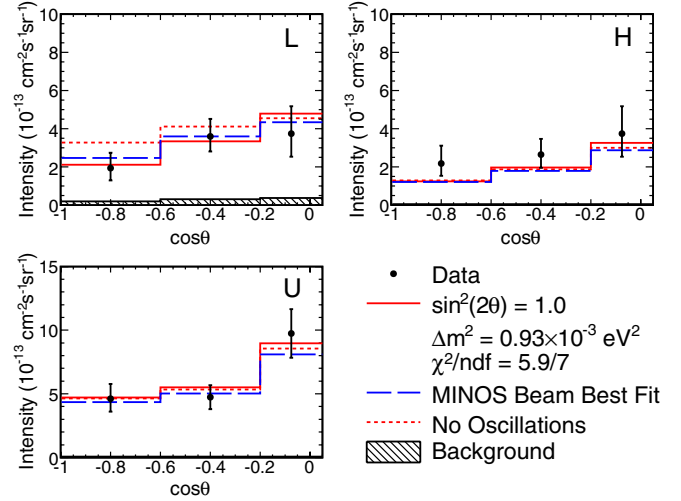


FIG. 14 (color online). The intensity of neutrino-induced muons as a function of $\cos\theta$. The data are shown by the points, the best fit is shown by the solid line, and the null oscillation hypothesis is shown by the dotted line. The prediction using the MINOS result with the NuMI neutrino beam is shown by the dashed line. The top left panel shows the events with $1 < p_{\text{fit}} < 10 \text{ GeV}/c$ (L), the top right shows the events with $10 \leq p_{\text{fit}} < 100 \text{ GeV}/c$ (H), and the bottom left shows events with unknown momentum (U).

superposed onto the observed intensity of neutrino-induced muons. For the results from the null oscillation hypothesis and the MINOS beam analysis, we varied the ϵ_k to minimize Eq. (4) at the fixed values of $\sin^2 2\theta_{23}$ and Δm_{23}^2 .

Figure 15 shows the 68% and 90% confidence intervals for the oscillation parameter fit. The points within the 68% contour have a difference in χ^2 with the best-fit point, $\Delta\chi^2$, of less than 2.3. The points within the 90% contour have $\Delta\chi^2 < 4.61$. The null oscillation hypothesis has $\Delta\chi^2 = 5.5$ and is excluded at the 94% confidence level. Also shown in Fig. 15 is the 90% confidence interval for the MINOS contained vertex analysis [17] based on the first

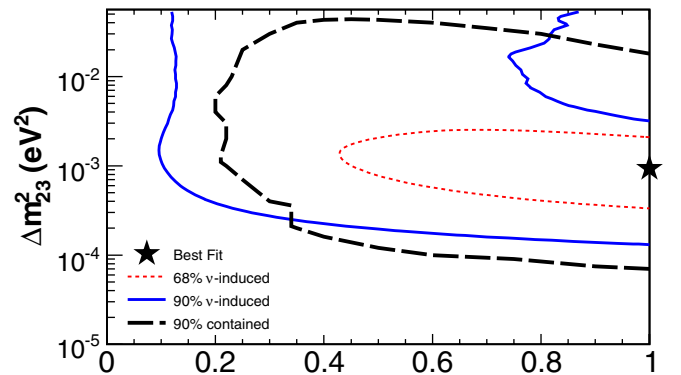


FIG. 15 (color online). The 68% (dotted line) and 90% (solid line) confidence intervals for the oscillation parameter fit. The best-fit point is indicated by the star. Also shown is the 90% confidence interval for the MINOS contained vertex analysis (dashed line) based on the first 418 days of data taking with the far detector.

418 days of data taking with the far detector. The results of these two analyses are consistent.

V. TESTS OF *CPT* CONSERVATION

To probe *CPT* conservation with MINOS data, we use two tests. After discussing the systematic uncertainties, we first compare the strength of the oscillation signals for the neutrino-induced μ^- and μ^+ samples individually. Second, we divide the charge-separated muons into low and high momentum samples and further test for differing rates of neutrino disappearance. We show that this second test is the one that minimizes systematic uncertainties.

A. Systematic uncertainties

Here we discuss the additional sources of systematic uncertainties associated with the *CPT* analysis. One additional source of systematic uncertainty is due to misidentified charge sign or the charge purity of the data sample. We determined the charge purity systematic by comparing the underground cosmic ray muon charge ratio for events passing the $\chi^2_{\text{line}}/\text{ndf}$ cut with the charge ratio found by MINOS, $N_{\mu^+}/N_{\mu^-} = 1.371$ [34]. For this comparison we divided the cosmic ray muons into two data samples, one with muons having momenta <30 GeV/ c and the other with momenta in the range 30–100 GeV/ c . For cosmic ray muons in the lower momentum sample, we calculated the difference in the charge ratio to be 0.047, a difference that can be attributed to an impurity of 5%. For cosmic ray muons in the high momentum sample, we calculated the difference in the charge ratio to be 0.12, which can be attributed to an impurity of 14%. By taking a weighted average of these impurities, where the weights are given by the number of neutrino-induced muons in each momentum range, we calculated the impurity in the neutrino-induced muon data to be 6%. Using similar methods we found the weighted average impurity for the neutrino-induced Monte Carlo events to be 1%. These weighted impurities are given in Table V (4).

There are also additional sources of systematic uncertainty due to the physics model. These arise from uncertainties in the ratios of the $\bar{\nu}_\mu$ cross section, $\sigma_{\bar{\nu}_\mu}$, to the ν_μ cross section, σ_{ν_μ} , and we estimate the uncertainties to be 9% [37] for energies <30 GeV (9) and 2% [41] for energies >30 GeV (10). The uncertainty in the ratio of the number of ν_μ to $\bar{\nu}_\mu$ (11) is 1% [36].

B. Charge-separated event ratio

Our first test of *CPT* conservation comes from the comparison of the ratio of the total number of μ^- events to μ^+ events in the data and the Monte Carlo simulation including backgrounds. In the data, this ratio of charge-separated muons is given by

TABLE V. Sources of systematic uncertainty considered and their effects on the ratio \hat{R}_{CPT} .

Source	σ	$(\Delta\hat{R}_{CPT})_k$
Reconstruction systematics:		
(1) $\chi^2_{\text{fit}}/\text{ndf} < 1.5$	0.01	$<5 \times 10^{-4}$
(2) $\chi^2_{1/\beta}/\text{ndf} < 3.0$	0.01	$<1 \times 10^{-4}$
(3) $\chi^2_{\text{line}}/\text{ndf} < 10$	0.27	$<5 \times 10^{-3}$
(4) Charge purity (data, MC)	(0.06, 0.01)	+0.07
Model systematics:		
(5) Normalization	0.15	$<1 \times 10^{-4}$
(6) Spectral index	0.03	$<1 \times 10^{-4}$
(7) Cross section ($E < 30$ GeV)	0.07	$<5 \times 10^{-3}$
(8) Cross section ($E > 30$ GeV)	0.02	$<5 \times 10^{-3}$
(9) $\sigma_{\bar{\nu}_\mu}/\sigma_{\nu_\mu}$ ($E < 30$ GeV)	0.09	0.04
(10) $\sigma_{\bar{\nu}_\mu}/\sigma_{\nu_\mu}$ ($E > 30$ GeV)	0.02	$<5 \times 10^{-3}$
(11) $\nu_\mu/\bar{\nu}_\mu$	0.01	0.01
$\Delta\hat{R}_{CPT} = \sqrt{\sum(\Delta\hat{R}_{CPT})_k^2}$		-0.04, +0.08

$$R_{\mu^-/\mu^+}^{\text{data}} = \sum(N_{\mu^-}) / \sum(N_{\mu^+}). \quad (5)$$

The Monte Carlo ratio is defined in a similar way. If the ν_μ oscillate with the same parameters as the $\bar{\nu}_\mu$ then the ratio of the data and Monte Carlo ratios, \hat{R}_{CPT} , will be consistent with unity. The systematic uncertainties in \hat{R}_{CPT} are shown in Table V. This table shows how changes in the default parameters corresponding to the sources of systematic uncertainty change the value of \hat{R}_{CPT} . The main contributions to the uncertainty in \hat{R}_{CPT} are the purity, the cross section ratio for neutrinos with energies <30 GeV, and the ratio of $\nu_\mu/\bar{\nu}_\mu$. Since the ratio in Eq. (5) is greater than 1, more negative than positive muons will be misidentified and the measured charge ratio decreases towards unity. Consequently, charge misidentification leads to a one-sided (positive) error on the measured charge ratio. The uncertainty due to the impurity is $(\Delta\hat{R}_{CPT})_4 = 0.07$. The cross section ratio for neutrinos with energies <30 GeV changes the ratio by $(\Delta\hat{R}_{CPT})_9 = \pm 0.04$. The uncertainty in the ratio of the number of ν_μ to $\bar{\nu}_\mu$ corresponds to $(\Delta\hat{R}_{CPT})_{11} = \pm 0.01$. We find from Tables III and V

$$\hat{R}_{CPT} = \frac{R_{\mu^-/\mu^+}^{\text{data}}}{R_{\mu^-/\mu^+}^{\text{MC}}} = 0.72^{+0.24}_{-0.18}(\text{stat})^{+0.08}_{-0.04}(\text{syst}), \quad (6)$$

where the statistical uncertainties were calculated using the method for Eq. (3). This value of \hat{R}_{CPT} is consistent with *CPT* conservation.

C. Charge-separated, low-to-high momentum event ratio

Our second test of CPT conservation is based on the charge-separated ratios of low momentum to high momentum events compared to the Monte Carlo expectation including backgrounds. First we define ratios of data to Monte Carlo expectation for the ν_μ and $\bar{\nu}_\mu$ individually,

$$\mathcal{R}_- = \left(\frac{R_{L/H}^{\text{data}}}{R_{L/H}^{\text{MC}}} \right)_{\mu^-} = 0.49_{-0.14}^{+0.20}(\text{stat}) \pm 0.07(\text{syst}) \quad (7)$$

and

$$\mathcal{R}_+ = \left(\frac{R_{L/H}^{\text{data}}}{R_{L/H}^{\text{MC}}} \right)_{\mu^+} = 0.55_{-0.19}^{+0.30}(\text{stat}) \pm 0.07(\text{syst}). \quad (8)$$

The systematic uncertainties in these ratios are listed in Table VI, and the statistical uncertainties were again calculated using the same method as for Eq. (3). We define the ratio of ratios $\tilde{\mathcal{R}}_{CPT}$,

$$\tilde{\mathcal{R}}_{CPT} = \mathcal{R}_- / \mathcal{R}_+ = 0.89_{-0.33}^{+0.54}(\text{stat}) \pm 0.03(\text{syst}). \quad (9)$$

We found the upper and lower limits for the statistical uncertainty in Eq. (9) by fitting the observed L and H event rates for μ^- and μ^+ to a model which used $\tilde{\mathcal{R}}_{CPT}$ and three of the four event rates as inputs. The data and model were compared using a χ^2 function appropriate for Poisson distributed data [39]. The 68% C.L. interval quoted includes all values of $\tilde{\mathcal{R}}_{CPT}$ which produce a χ^2 within 1 of the best-fit value. Again within the statistical and systematic uncertainties, the value of $\tilde{\mathcal{R}}_{CPT}$ is consistent with unity and CPT conservation.

The value of this test is clearly demonstrated in Table VI where it can be seen that several sources of systematic uncertainty cancel in the calculation of $\tilde{\mathcal{R}}_{CPT}$. These cancellations occur because the μ^- and μ^+ ratios are affected

similarly for these systematic uncertainties and so divide out in the ratio. The uncertainty in the ratio of the number of ν_μ to the number of $\bar{\nu}_\mu$ does not contribute to the uncertainty in $\tilde{\mathcal{R}}_{CPT}$. As seen in Table VI, the only sources of uncertainty which do not cancel are those in the relative magnitude of the ν_μ and $\bar{\nu}_\mu$ cross sections at low (9) and high (10) energies.

This approach should prove valuable in high statistics accelerator-based tests of CPT conservation because it minimizes systematic uncertainties so effectively.

VI. CONCLUSIONS

The MINOS far detector has taken 854.24 days of data in its search for the atmospheric neutrino-induced muons. We found a total of 140 neutrino-induced muons in the data set. We have performed four analyses. The ratio of the number of low momentum muons to the sum of high momentum and unknown momentum muons has been compared to the same ratio for Monte Carlo events. The resulting ratio, \mathcal{R} , suggests neutrino disappearance at the 2.0σ level. The best oscillation fit to the data gives $\sin^2 2\theta_{23} = 1.0$ and $\Delta m_{23}^2 = 0.93 \times 10^{-3} \text{ eV}^2$. This fit excludes the null oscillation hypothesis at the 94% confidence level. We used two methods to look for evidence of different rates of neutrino disappearance as a test of CPT conservation. Both tests are consistent with CPT conservation. The cancellation of many systematic uncertainties in the computation of $\tilde{\mathcal{R}}_{CPT}$ suggests that this test would provide a precision test of CPT conservation with a suitably large data set.

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TABLE VI. Sources of systematic uncertainty considered and their effects on the ratios \mathcal{R}_- , \mathcal{R}_+ , and $\tilde{\mathcal{R}}_{CPT}$.

Source	σ	$\Delta \mathcal{R}_-$		$\Delta \mathcal{R}_+$		$(\Delta \tilde{\mathcal{R}}_{CPT})_k$
		-1σ	$+1\sigma$	-1σ	$+1\sigma$	
Reconstruction systematics:						
(1) $\chi_{\text{fit}}^2/\text{ndf} < 1.5$	0.01	$< 5 \times 10^{-4}$	$< 5 \times 10^{-4}$	$< 5 \times 10^{-4}$	$< 5 \times 10^{-4}$...
(2) $\chi_{1/\beta}^2/\text{ndf} < 3.0$	0.01	$< 1 \times 10^{-4}$	$< 1 \times 10^{-4}$	$< 1 \times 10^{-4}$	$< 1 \times 10^{-4}$...
(3) $\chi_{\text{line}}^2/\text{ndf} < 10$	0.27	0.05	0.05	0.05	0.05	...
(4) Charge purity (data, MC)	(0.06, 0.01)	$< 1 \times 10^{-4}$	$< 1 \times 10^{-4}$	$< 1 \times 10^{-4}$	$< 1 \times 10^{-4}$...
Model systematics:						
(5) Normalization	0.15	$< 1 \times 10^{-4}$	$< 1 \times 10^{-4}$	$< 1 \times 10^{-4}$	$< 1 \times 10^{-4}$...
(6) Spectral index	0.03	0.04	0.03	0.04	0.04	...
(7) Cross section ($E < 30 \text{ GeV}$)	0.07	0.02	0.02	0.02	0.02	...
(8) Cross section ($E > 30 \text{ GeV}$)	0.02	0.01	0.01	0.01	0.01	...
(9) $\sigma_{\bar{\nu}_\mu}/\sigma_{\nu_\mu}$ ($E < 30 \text{ GeV}$)	0.09	0.02	0.02	0.03
(10) $\sigma_{\bar{\nu}_\mu}/\sigma_{\nu_\mu}$ ($E > 30 \text{ GeV}$)	0.02	0.01	0.01	0.01
$\Delta \tilde{\mathcal{R}}_{CPT} = \sqrt{\sum (\Delta \tilde{\mathcal{R}}_{CPT})_k^2}$						0.03

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