# COST OF BEING A MEXICAN IMMIGRANT AND BEING A MEXICAN NON - CITIZEN IN CALIFORNIA AND TEXAS

A Thesis

by

ISAO TAKEI

Submitted to the Office of Graduate Studies of Texas A&M University in partial fulfillment of the requirements for the degree of

MASTER OF SCIENCE

August 2005

Major Subject: Sociology

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Approved by:

Chair of Committee, Rogelio Saenz Committee Members, Don Albrecht

William Alex. McIntosh

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#### **ABSTRACT**

Cost of Being a Mexican Immigrant and Being a Mexican Non-Citizen in California and Texas. (August 2005)

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Chair of Advisory Committee: Dr. Rogelio Saenz

The purpose of this thesis is to examine hourly wage differences across different groups of Mexican-origin workers. First, I assess the cost of foreign-born status by comparing the hourly wages of Mexican immigrant workers with those of native-born Mexican American workers. Second, I assess the cost of non-citizenship status by comparing the hourly wages of non-citizens with those of Mexican-born U.S. naturalized citizens. I also seek to determine if these costs are greater in California than in Texas. The data are drawn from the 2000 5% Public Use Microdata Sample (PUMS) U.S. Census. The results from multiple linear regression analyses show that being an immigrant, particularly a non-citizen immigrant, is associated with lower hourly wages, especially in California. Thus, Mexicanorigin workers, especially those in California, bear dual costs for being foreign-born and not being naturalized citizens. Furthermore, I focus on length of U.S. residence to assess the social and economic impact of the different periods on the costs associated with foreign-born status. First, those who came to the United States before the IRCA of 1986 and a series of California propositions during the 1990s have higher hourly wages than those who arrived later, because of more stable labor market conditions and the effect of the duration of stay in

the United States. Second, those who arrived during the last decade have much lower hourly wages because of their disadvantaged labor market contexts.

#### **ACKNOWLEDGEMENTS**

My thanks go to the many persons who have supported me during the years. I would like to thank Dr. Rogelio Saenz, my committee chair, for his valuable advice and continuing support in my studies at Texas A&M University. He definitely gave me solid perspectives on my Mexican demographic research. His support for my presentations at major sociological conferences became the base of this study, and extended my research skills in demography and labor market analyses. I also thank Dr. William Alex. McIntosh and Dr. Don Albrecht, my excellent committee members, for their insightful comments, guidance and support throughout the course of this study. Their important feedback and advice encouraged my future studies in demography and labor markets.

Thanks also to my friends and other professors. I thank Dr. Wesley R. Dean for offering me a research assistantship position in the spring and summer of 2005. I learned various statistical analyses and survey methodology in the project. Dr. Jon P. Alston willingly worked with me for cultural studies during my three years of studies in Texas A&M University. I would like to thank him for his time and efforts in publishing a journal article and a book with me. Among my friends in the United States, I offer special thanks to Jing Li for her generous help. I also thank Dr. Arthur Sakamoto, Professor in the Department of Sociology at the University of Texas at Austin, for his technical advice for my statistics questions. In Japan, Professor Saburo Sato, my former academic adviser of the College of International Relations in Nihon University, has been encouraging and supporting of my research for years with his broad knowledge and kindness. Finally, I want to say, "Thank you" to my parents Kazuji Takei and Etsuko Takei for their continuous support for twenty-eight years.

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#### CHAPTER I

#### **INTRODUCTION**

Due to its major growth, the Mexican-origin population has attracted much attention in labor market research (Chiswick, 1986). However, much of this research has been based on Mexican-Anglo comparisons (*e.g.*, Bean *et al.*, 1987; Borjas, 1983, 1983-1984, 1990, 1999; Borjas, 2000; Borjas and Freeman, 1992; Borjas and Tienda, 1987; Chapa, 1990; Chiswick, 1978; Frienberg and Hunt, 1995; Huddle, 1993; Kritz and Nogle, 1994; Marshall, 1984, 1987; Melendez *et al.*, 1991; Simon *et al.*, 1993; Talyor *et al.*, 1988; Winegarden and Khor, 1991; Wright *et al.*, 1997). Yet, this research approach is based on the aggregation of the entire Mexican-origin population, suggesting that the group is homogeneous.

For the comprehensive examination of the heterogeneity of Mexican-origin workers, the thesis focuses on internal hourly wage differences within this group. Among a number of human capital attributes, I focus on nativity status (native- and foreign-born) and immigrants' possession of U.S. citizenship (foreign-born with citizenship and non-citizen foreign-born status) as two core factors. Specifically, the thesis focuses on the cost of being an immigrant (the hourly wage differences between the native-born and the foreign-born) and the cost of being an immigrant without U.S. citizenship (the wage differences between naturalized immigrants and non-naturalized immigrants).

Demographic trends point to the increasing importance of Mexican immigrants in research on labor markets, especially those involving workers of Mexican origin. Indeed, among persons of Mexican origin, the foreign-born (96%) increased nearly three times faster

The thesis follows the style of *International Migration Review*.

than did the native-born (36%) between 1990 and 2000 (Saenz, 2004a). Consequently, the share of the foreign-born in the Mexican-origin population rose from 36 percent in 1990 to 40 percent in 2000 (Saenz, 2004a).

Furthermore, foreign-born Mexicans have a significant attachment to the U.S. labor force. For instance, 71.4 percent of foreign-born males were in the labor force in 2000. However, they are especially likely to be in certain segments of the labor market. About 65% of the Mexican immigrant workers were employed in specific industrial niches, such as construction, farming, repair and maintenance, and manufacturing industries (Saenz, 2004a). The word "Latinoization of the U.S." denotes the disproportionate presence of Mexicans and other Latinos in selected occupations and industries. The waves of Mexican immigrants also include a small but steady increasing portion that have specialized skills (Alarcon, 1999; Castells, 1996; Clark, 1998; Keely, 1974; Kritz, 1987; Pellegrino, 2001; Roberts *et al.*, 1999; Saxenian, 1996).

The majority of Mexican immigrants has historically been concentrated in the Southwest (Bean and Tienda, 1987; Durand *et al.*, 2000), especially in California and Texas (Bustamante, 1997; Jargowsky, 1997; Massey, 1996; McCall, 2000a; Portes and Bach, 1980; Saenz, 1991; Stolzenberg, 1990). In these two primary destination states, large Spanish-speaking communities are well established (e.g., El Paso, Houston and San Antonio, TX and Los Angeles, San Francisco, San Ysidro and San Diego, CA) (Johnson and Oliver, 1989). Because of this large concentration, Clark (1998) points out that Los Angeles and Dallas have higher poverty rates for all immigrants than the U.S. as a whole. Kandel and Cromartie (2004) report that in the majority of southwestern counties, Latinos make up over 10 percent

of the total population and continues to grow through a combination of high natural increase and net immigration.

Although both California and Texas continue to attract the majority of Mexican immigrants, these persons have encountered hasher treatment in California at least over the last decade. The Immigration Reform and Control Act (IRCA) of 1986 caused an expansion of the labor supply in the state; while the law was not particularly successful in reducing Mexican illegal migration to the U.S., about 2.3 million newly legalized Mexicans entered local labor markets (Durand *et al.*, 2000). Due to a fear triggered by the massive growth of the immigrant population, mainly low-skilled undocumented Mexicans, California passed a series of anti-immigrant laws such as Proposition 198, 209, and 227 during the 1990s.

The passage of these welfare reform policies also represents social hostility that induced unfavorable treatment toward Mexican-origin workers in particular. The fiscal imperatives have directed their attention to Mexican migrant workers by unreasonably blaming them as "undeserving poor," because of their "culture of dependency," suggesting that they migrate to the U.S. to draw societal resources (Kurthen, 1997). Furthermore, as can be seen in Los Angeles riots of 1992, high population concentration of immigrants, especially undocumented immigrants, generated social unrest among the native-born who worried about rising crime rates (Sanchez, 1997).

It is likely that these harsh political and social contexts in California are translated to worse labor market outcomes for Mexican immigrants in California than in Texas. First, employer sanctions by IRCA and sentiments toward Mexicans have induced general discrimination against Mexican-origin workers in general (Davila *et al.*, 1998). Second, after IRCA, Mexican immigrants faced wage penalties based on the absence of legal status

(Donato and Massey, 1993; Philip and Massey, 1999; Sorensen and Bean, 1994). As greater concentration of immigrants is associated with lower hourly wage (Borjas, 1987a, 1987b, 1994, 1995; Tienda and Lii, 1987; Topel, 1994), hourly wages are expected to be lower in California than in Texas, all else equal.

The literature on anti-immigrant sentiments and laws in California remind us of the importance of considering not only nativity status and legal status of immigrants, but also non-economic factors (e.g., social and political contexts) for a comprehensive analysis of the labor market position of Mexicans. Therefore, to closely examine the state and local labor market characteristics, the thesis focuses on the hourly wage differences between California and Texas, instead of examining much broader social and economic contexts, such as comparing different standard Census regions (e.g., Midwest and West) and examining the entire country. I regard the word "labor market" as geographic areas beyond occupational and industrial categories to examine the impact of regional differences on wage.

Associated with the major factors mentioned above, another central element of the analysis is immigrants' length of U.S. residence. The literature suggests important links between the social and economic outcomes of immigrants and their temporal presence in this country. This focus gives the analyses of the costs associated with foreign-born status and a lack of U.S. citizenship a time dimension in considering variations across California and Texas with respect to the timing and volume of Mexican immigrants.

The data are drawn from the 2000 5% Public Use Microdata Sample (PUMS). This study examines geographic differences in state-specific labor market contexts that are major determinants of wage (*see* Semyonov, 1988), comparing separate models of the two largest concentration states of Mexicans: California and Texas.

This thesis contains five chapters. The first chapter presented here provides a brief introduction to the purpose and relevance of this study to the labor market research of Latinos/as, and discusses the value of this within-group wage analysis for a better understanding of the Mexican-origin population in the U.S.

Chapter II presents a detailed literature review and theoretical frameworks derived from past empirical and theoretical studies of ethnic labor markets and immigration. First, major labor market perspectives are reviewed, and the significance of examining nativity status and broad labor market contexts is emphasized. Second, I introduce the social and economic impacts of anti-immigrant sentiments and laws in California, and argue that the two largest concentration states of the Mexican-origin people in the Southwest, California and Texas, are associated with distinct labor market conditions that contribute to disparate labor market outcomes. Third, major individual-level attributes are introduced according to the human capital perspective. Based on the literature review, chapter II also summarizes theoretical approaches that are used in the analysis, and outlines the hypotheses drawn from theory.

Chapter III presents the methodology, data, variables, and statistical procedures used to conduct the analysis. Chapter IV presents the findings from the multiple linear regression models, and detailed discussions are given in line with the literature. Chapter V presents an overview of the findings, along with contribution of this study to Latino/a demography, public policy implications, strengths and limitations of this study, and provides a set of recommendations for future research.

#### CHAPTER II

## LITERATURE REVIEW, THEORETICAL PERSPECTIVES AND HYPOTHESES

As noted in chapter I, much of the research on wages involving Mexican-origin workers has been based on Mexican-Anglo wage differences. This research approach derives from labor market theories, which suggest that ethnic minorities tend to have comparatively fewer human capital resources and to receive lower returns to these resources than whites. Relatively little research has examined internal differentiation within the Mexican-origin population. This chapter is divided into five sections. In the literature review section, I first introduce the dual labor market and ethnic enclave economy theoretical perspectives, and point out that they do not sufficiently focus on both social context (e.g., regional differences) and individual differences (particularly native-/foreign-born status). Second, I introduce literature on anti-immigrant policies and sentiments in California during the last decades, to illustrate the fact that different groups of Mexican-origin workers faced variations in their environments which affect the opportunity structures of immigrants. Third, I provide a brief discussion of two key elements of the analysis—immigrant status and naturalization status. Fourth, major individual-level attributes related to wages, drawn from the human capital perspective, are introduced into the discussion. The last portion of the chapter presents a series of hypotheses drawn from the insights of the literature review and theoretical perspectives.

#### Theories of Ethnic Labor Markets

Theories of ethnic labor markets emphasize the economic structure associated with labor markets over human capital attributes in determining wages (Cornelius, 1981; Kalleberg and Sorensen, 1979; Portes and Bach, 1985; Portes and Truelove, 1987; Reed,

2001; Sakamoto and Chen, 1991). A major prediction of the dual labor market theory (*e.g.*, Bustamante, 1976; Edwards *et al.*, 1975; Espenshade, 1995; Gordon, 1972; Kossoudji, 1989; Massey and Espinosa, 1997; Massey *et al.*, 1993; Piore, 1979; Portes, 1981; Portes and Bach, 1980), for example, is that racial minorities and immigrants are disproportionately found in low-wage and unstable occupations identified with the secondary labor markets. Guided by the dual-system explanation of the U.S. labor market, past empirical studies have a strong tendency to aggregate Mexican-origin workers with different backgrounds as homogeneous, emphasizing their economic subordination in comparison to Anglos.

The ethnic enclave economy (e.g., the barrios of East Los Angeles, *see* Borjas, 1999), in contrast, suggests that immigrant and ethnic employees receive higher economic returns to their human capital resources as well as a range of non-monetary rewards (e.g., prospects of upward socioeconomic mobility) *within* mono-ethnic segmented labor markets (*e.g.*, Bayer, 1968; Bonacich, 1972, 1973; Espenshade and King, 1994; Kossoudji, 1989; Light *et al.*, 1994; Massey, 2001; Model, 1992; Nee and Sanders, 1987; Portes and Jensen, 1989, 1992; Sanders and Nee, 1987, 1992; Wilson and Martin, 1982; Zhou and Logan, 1989). For instance, a number of studies describe the relative strength of Mexican social networks for their employment compared to other minority groups (*e.g.*, Bailey and Waldinger, 1991; Donato *et al.*, 1992a; Enchautegui, 1998; Holzer, 1998; Gurak and Caces, 1992; Roberts *et al.*, 1999).

However, these two theoretical approaches have some shortcomings (*see* Portes and Zhou, 1993). For example, researchers claim that the potential benefits of employment in ethnic enclaves are ambiguous because the perspective does not consider the relation of ethnic enclaves to the larger labor market and ignores workers engaged in this sector of the

<sup>&</sup>lt;sup>1</sup> For Mexican immigrants' cases, see Bustamante (1976) and Portes and Bach (1980).

economy (Massey *et al.*, 1994; Reitz *et al.*, 1981, cited in Roos and Hennessy, 1987; Tienda and Lii, 1987; Waldinger, 1986).

In their book titled *The Latin Journey*, Portes and Bach (1985) assert that there is no enclave economy for Mexican immigrants; instead Mexican immigrants participate as low-wage labor in the open economy. Wilson and Portes (1980) explain that the development of enclaves requires two conditions: first, the presence of immigrants with sufficient capital and initial entrepreneurial skills; and second, the renewal of the enclave labor force through a steady stream of new arrivals (*also see* Massey, 2001).<sup>2</sup> The continuing influx of newcomers (Lien, 1994) due to the interconnectedness of the U.S. and Mexican economies (Massey and Espinosa, 1997) supports the second condition. However, as Mexican immigrants tend to have relatively limited human capital (Bean and Stevens, 2003; Cohen, 1989; Madhavan, 1985; Morales and Ong, 1993; Saenz, 1999; Tienda, 2002), the first condition is more tenuous in the case of Mexican immigrants. Nonetheless, Borjas (1999) observes the existence of Mexican ethnic enclaves.

Yet, ethnic labor market perspectives often ignore the following two important factors. First, these perspectives tend to ignore the differences between native- and foreign-born workers in the ethnic communities. Rather, labor market studies of Latinos/as focus on their collectivity, assuming that native- and foreign-born workers are connected in some way. Second, these perspectives do not necessarily consider the effect of broad social contexts (e.g., regional differences) on wages encompassing different types of labor markets (*see* Zucker and Rosenstein, 1981). Instead, their empirical evidence is limited to minority labor

<sup>&</sup>lt;sup>2</sup> For further discussions of the ethnic enclave economy, see Bailey and Waldinger (1991) and Logan *et al.* (1994).

market concentration in metropolitan areas and specific occupations, assuming that their status attainment is low.

Thus, much research has been conducted within a relatively limited scope of empirical questions, which inhibit the capturing of a large portrait of the diversity of the Mexican labor force. Accordingly, within-group wage differences between native- and foreign-born Mexican-origin workers have not yet been comprehensively analyzed (Bean *et al.*, 1988; Bradshaw and Frisbie, 1983; Saenz, 2004b; Semyonov, 1988).

# State Differences in Anti-Immigrant Policies and Sentiments

Kalleberg and Sorensen (1979) argue that "labor markets" can be used to denote geographic areas instead of occupational and industrial groups. Indeed, the description of the literature on the impact of regional differences on wages reminds us of the importance of considering the state and local labor market characteristics where people reside. This perspective enables us to better understand the different labor market situations of Mexicanorigin workers because such perspective also takes into account non-economic factors (e.g., social issues and politics) that often influence the economic outcomes of minority workers.

To closely examine the state and local labor market characteristics, the thesis focuses on hourly wage variations within Mexican-origin workers across California and Texas. In spite of the fact that both states continue to attract the majority of Mexican immigrants (Bean and Tienda, 1987; Bustamante, 1997; Clark, 1998; Durand *et al.*, 2000; Jargowsky, 1997; Johnson and Oliver, 1989; Kandel and Cromartie, 2004; Massey, 1996; McCall, 2000a; Portes and Bach, 1980; Saenz, 1991; Stolzenberg, 1990), it is likely that variations in the strength of anti-immigrant policies and sentiments between California and Texas during the

last decades resulted in different labor market conditions for Mexican-origin immigrants with those in California experiencing harsher conditions than their counterparts in Texas.

The Immigration Reform and Control Act (IRCA) of 1986 directly affected the economic situation of different groups of Mexican-origin workers (*see* Baker, 1997; Davila *et al.*, 1998; Donato and Massey, 1993).<sup>3</sup> IRCA for the first time made it illegal for employers to knowingly hire undocumented workers, imposing both civil and criminal penalties against those who did (Durand *et al.*, 2000). Because of the disproportionate concentration of Mexican immigrants in the Southwest (especially in California and Texas), the impact of IRCA was especially obvious in this region.

Although IRCA was intended to reduce the number of undocumented migrants, it did not accomplish this goal. Indeed, research suggests that IRCA was not particularly successful in reducing Mexican illegal migration to the United States (Baker, 1997; Donato and Massey, 1993; Donato *et al.*, 1992b); most employers mainly continued to hire undocumented migrants at lowered wages through increasing subcontract arrangements. Yet, IRCA did have other consequences. For example, about 2.3 million Mexicans acquired legal documents through IRCA, allowing many to enter the larger traditional labor market (Donato *et al.*, 1992a; Durand *et al.*, 2000; Philip and Massey, 1999; Valdes, 1995). In addition, employer sanctions induced general discrimination against undocumented-appearing Mexican-origin workers, including U.S.-born workers (Davila *et al.*, 1998).

<sup>&</sup>lt;sup>3</sup> Philip and Massey (1999) explain that IRCA sought to control undocumented immigration from Mexico by (1) sanctions against employers who knowingly hired undocumented migrants; (2) additional resources for the U.S. border patrol; and (3) authorizing amnesty for undocumented migrants who could prove continuous residence in the United States after January 1, 1982.

<sup>&</sup>lt;sup>4</sup> Although the number arrested by the Immigration and Naturalization Service (INS) declined between 1986 and 1989, it surpassed pre-IRCA figures in 1990 (Williams, 1991; cited in Valdes, 1995).

Many studies have examined IRCA's effects on the economic outcomes of Mexicanorigin workers, mostly the immigrant population (Davila *et al.*, 1998; Philip and Massey,
1999; Sorensen and Bean, 1994). Such research has indicated that IRCA resulted in wage
penalties against workers with undocumented status (Davila *et al.*, 1998; Donato and Massey,
1993; Donato *et al.*, 1992a); 5 some undocumented migrants worked fewer hours after 1986
because of limited opportunities (Donato *et al.*, 1992a), while others increased their work
effort in order to compensate for the new shortfall in wages (Davila *et al.*, 1998). The wage
penalty for being undocumented suggests that there is a reward for being a legalized
immigrant. In fact, research by Aguilera (2004) shows that Mexicans who became legal
through IRCA experienced upward mobility. 6

In the post-IRCA period, the effect of human capital (e.g., occupation and the duration of trips to the U.S.) declined in determining wages of both legal and undocumented migrants (Philip and Massey, 1999), with legal status becoming the primary determinant of wage rates (Donato and Massey, 1993). More specifically, Massey (1987) found that before IRCA, legal status had no effect on wage rates among Mexican migrants once selectivity and background differences between documented and undocumented migrants were controlled. Among both legal and undocumented migrants, wage rates tended to increase with rising age, education, labor force experience, U.S. migrant experience, and length of stay, and being urban origin migrants and nonagricultural workers (Massey, 1987b). These factors become less important determinants of wages compared to legal status following the implementation of IRCA.

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<sup>&</sup>lt;sup>5</sup> After IRCA, undocumented migrants working in the nonagricultural sector earned wages that were 22% lower than those earned by documented migrants with similar characteristics, and those working in the agricultural sector earned an additional 33% less (Philip and Massey, 1999).

<sup>&</sup>lt;sup>6</sup> However, Philip and Massey (1999) and Sorensen and Bean (1994) argue that IRCA adversely affected the wages of legal immigrants.

Although IRCA is a federal law, Mexican immigrant workers in California might have faced harsher labor market conditions than their counterparts in Texas, due to a greater concentration of Mexican immigrants in California. Indeed, California initiatives such as Propositions 187, 209 and 227<sup>7</sup> represent societal reactions and fears related to the perceived social and economic costs associated with the increasing presence of immigrants, mainly low-skilled undocumented Mexicans (Kurthen, 1997; Martin, 1995; Tolbert and Hero, 1996). The fiscal imperatives have directed their attention to Mexican migrant workers who are often labeled as the "undeserving poor" because of their so-called "culture of dependency" (Kurthen, 1997). Moreover, as can be seen in the Los Angeles riots of 1992, a number of new immigrants were involved in social unrest which escalated worries about rising crime rates (Sanchez, 1997).

The passage of anti-immigrant policies and hostility suggest that the costs associated with foreign-born status and the lack of naturalized status are greater in California than in Texas. However, it is important to examine the broader human capital literature to obtain a more expansive view of the factors that affect the labor market earnings of Mexican immigrants.

#### **Immigrant Status and Naturalization Status**

Past research shows that for individuals, immigrant status is a key factor affecting their wages (Chiswick, 1978; Cohen, 1989; Enchautegui, 1998; Heer, 1990; Madhavan,

<sup>7</sup> Proposition 187 in 1994 blocked the access of illegal immigrants to public education (from kindergarten through university), welfare, and non-emergency health care services. Proposition 209 in 1996 eliminated affirmative action programs. Proposition 227 in 1998 officially abolished bilingual education programs in public schools, as public education is the most costly service used by illegal aliens in California (Martin, 1995). For detailed explanations of these propositions, see Purcell (1997) and Tolbert and Hero (1996).

<sup>&</sup>lt;sup>8</sup> Borjas (1999) argues that more recent immigrant waves are also more likely to use welfare than earlier waves. Borjas (1999:12) notes, for example, that "a comprehensive study by the National Academy of Sciences concluded that immigration raised the annual taxes of the typical native household in California by about \$1,200 a year."

1985; Meisenheimer, 1992; Schoeni, 1998; Sehgal, 1985). In fact, Mexican immigrants in the United States, whether legal or undocumented, tend to have lower wages than U.S.-born Mexican Americans (Chiswick, 1986; Saenz, 2004b). Moreover, naturalization status well reflects immigrants' labor market values. Thus, the rate at which an immigrant group acquires citizenship is important (Portes and Mozo, 1985). However, data from the 2000 5% PUMS show that only 22.5 percent of Mexican immigrants are naturalized, compared to 30.2 percent of other immigrants from Latin America. The Mexican's low naturalization rate suggests that foreign-born Mexican workers face more wage disadvantages than native-born Mexican American workers. Another major reason accounting for immigrants' disadvantageous position is that human capital acquired outside the United States is imperfectly remunerated in the U.S. labor market (Borjas, 1999; Chiswick, 1978; Chiswick *et al.*, 1997; Massey and Espinosa, 1997; Smith, 1984).

The disadvantage of being an immigrant varies by length of U.S. residence. In the short run, immigrants earn significantly less than native-born workers (Kossoudji, 1989). In all industrialized countries, there is apparently a tendency for citizens to be less willing to take certain menial and low-status jobs as economies advance (Marshall, 1984; Massey *et al.*, 1994). Thus, immigrants compensate for their lower wages by putting more working hours with their strong motivation to work (Dominguez and Fernandez de Castro, 2001).<sup>10</sup>

However, immigrants generally improve their wage returns to their human capital with increasing length of stay in this country (*see* Chiswick, 1986, 1978; Chiswick *et al.*,

<sup>&</sup>lt;sup>9</sup> Baker (1996:465) reports that "in 1994, nearly 40% of the Latino population in the United States was foreignborn, and only 18% of that foreign-born population had naturalized." This argument is supported by the potential "reversibility" of migration; immigrants who are difficult to return tend to naturalize at higher rates than those for whom return easily (Jasso and Rosenzweig, 1985, Portes and Mozo, 1985; cited in Portes and Truelove, 1987).

<sup>&</sup>lt;sup>10</sup> Chiswick (1978) argues that for the same schooling, age, and other demographic characteristics immigrants to the U.S. have more ability relevant to the labor market than native-born persons.

1997; Jensen, 1988; Simon and Sullivan, 1988; Tienda, 1983), although Borjas and Tienda (1993) argue that the disadvantages of undocumented relative to documented immigrants increase with age. With increasing time in the United States, Mexican immigrants have higher rates of English fluency, higher levels of education, higher presence in higher-status occupations, higher labor market wages, and lower poverty rates (Saenz, 2004a). In turn, these socioeconomic improvements increase the likelihood of naturalization (Liang, 1994). It is expected that all else equal, the costs associated with immigrant status and non-naturalization status are greater in California than in Texas due to the larger concentration of immigrants (Borjas, 1987a, 1987b, 1994, 1995; Tienda and Lii, 1987; Topel, 1994) and more recent immigrants in California.

## **Human Capital Perspective**

In order to assess the wage differences between native- and foreign-born Mexicanorigin workers, this study incorporates human capital factors into the analysis. The human
capital perspective emphasizes that humans invest in the accumulation of human capital
resources to maximize their labor market earnings (Becker, 1962, 1975; Borjas, 1983, 19831984; Cain, 1975; Chavez, 1991; Enchautegui, 1998; Espenshade, 1995; Kossoudji, 1989;
Lee, 1966; Massey, 1987a, 1987b; Mincer, 1974; Schultz, 1961; Tickamyer and Bokemeier,
1993; Trejo, 1997). Moreover, it is suggested that the levels of human capital significantly
differ by nativity status. In the following section, I briefly introduce selected human capital
factors and note the major differences between the Mexican foreign-born and native-born
with respect to human capital levels.

### **Educational Attainment**

It is generally agreed that education provides a means for upward socioeconomic

mobility (Becker, 1975; Borjas, 1999; Lacy *et al.*, 1989; Ruiz-Quintanilla and Claes, 1996). Educational attainment is an important investment because this is the base of other human capital factors such as specific vocational training (Massey, 1987b; Mincer, 1974), on-the-job experience and occupational skills (Juhn *et al.*, 1993; Kalleberg and Sorensen, 1979; Oi, 1962). Further, as a result of technological changes, specialized skills are thought to be rewarded with higher wages, even among workers with the same education (*e.g.*, Chiswick *et al.*, 1997; Nonini and Ong, 1997; McCall, 2000a).

However, past research demonstrates that educational attainment has a weaker impact on the labor market outcomes of immigrants compared to the native-born, suggesting that pre-immigration skills are less relevant in the U.S. labor market than are the skills acquired by the native-born (Chiswick *et al.*, 1997). Furthermore, Borjas (1999) argues that recent immigrants who arrived during the 1990s remain economically disadvantaged due to a lack of education, and he asserts that they reduce the wages of natives who compete with immigrant workers. Considering that foreign-born Mexicans tend have fewer years of formal schooling than native-born Mexican Americans (Chiswick, 1986), differences in education could be a key factor associated with the within-group wage gap.

# Age

Age is also an important form of human capital (*see* Freeman, 1979; Hughes and Hutchinson, 1988; Jolly *et al.*, 1978; Osberg *et al.*, 1986; Pissarides and Wadsworth, 1990; Rones, 1983). Thus, Joll *et al.* (1991) point out that younger workers tend to have lower pay and higher labor market turnover rates compared to older workers. Youth unemployment is accordingly a common problem across different racial and ethnic groups (Borowski, 1984;

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<sup>&</sup>lt;sup>11</sup> Although they differ in their effects on earnings, in the amounts typically invested, in the size of returns, and in the extent to which the connection between investment and return is perceived, all these investments help improve skills, and thereby contribute to raising earnings (Becker, 1975).

Brown and Sessions, 1997). <sup>12</sup> In general, unemployed workers who enter the labor market at an early age tend to be less educated, inexperienced, and lack enough networks to locate and secure jobs compared to older workers (Clogg and Shockey, 1985; Lacy *et al.*, 1989). Further, search theory assumes that relatively inexperienced younger workers, trying to find their most preferred match, are more likely to engage in voluntary unemployment as well as "jobhopping" than their older counterparts (Layard *et al.*, 1991). The relative youthfulness of Mexican immigrants compared to native-born Mexican Americans (Chiswick, 1986) suggests that the foreign-born workers face higher risk of unemployment and lower wages than their native-born counterparts.

## English Language Proficiency

English language proficiency is a crucial step in the broader process of economic mobility, and this human capital factor is strongly and positively related to earnings (*see* Chavez, 1991; Davila and Mora, 2000; Espinosa and Massey, 1997; Jasso and Rosenzweig, 1990; Massey, 1987a; McConnell and Leclere, 2002). Borjas (1999) notes that Latino immigrants who speak English earn 17 percent more than those who do not, even after adjusting for differences in education and other socioeconomic characteristics between the two groups. The degree of English language acquisition is a good indicator of acculturation, and linguistic isolation denotes lack of interaction outside the ethnic community (Stevens, 1992).

Researchers report that the majority of Mexican immigrants are monolingual Spanish speakers (Chiswick, 1986; Moore and Pachon, 1985; Saenz and Morales, 2005). This result is, in part, due to the large presence of recent immigrants among the Mexican foreign-born

<sup>&</sup>lt;sup>12</sup> In the U.S., less than half of the total pool of 16- to 19-year-old high school dropouts and graduates not going to college held full-time jobs in October 1988 (Lacy and Heffeman, 1989).

population (Saenz, 2004b). On the other hand, data from the 2000 5% PUMS show that the majority of native-born Mexican Americans speak English.

#### Additional Control Variables

Aside from human capital variables, other control factors are introduced into the analysis—metropolitan/nonmetropolitan residence, occupational categories and selfemployment status. Nonmetropolitan area residents tend to have lower hourly wages because of limited employment and occupational choices (Saenz and Torres, 2003). Although Mexican immigrants are disproportionately found in the low-skilled and low-wage labor force (Espenshade and Goodis, 1985; Massey and Schnabel, 1983; Portes and Truelove, 1987; Roos and Hennessy, 1987; Saenz and Torres, 2003; Waldinger, 1989a, 1989b, 1994), recent Mexican immigrants include a relatively small portion that have specialized skills (Alarcon, 1999; Castells, 1996; Clark, 1998; Keely, 1974; Kritz, 1987; Pellegrino, 2001; Roberts et al., 1999). Thus, there seems to be a large wage gap between the professional and non-professional occupational categories. Although self-employment is a classical form of labor force participation for ethnic minorities and immigrants (see Borjas, 1986; Fairlie and Meyer, 1996; Sanders and Nee, 1996; Spencer and Bean, 1999), Mexicans have relatively low self-employment rates (Saenz and Torres, 2003) because their self-employment is disadvantage-based, with a lack of large financial investment in a business and related resources (Borjas, 1986; Spencer and Bean, 1999).

#### **Summary and Hypotheses**

As Kalleberg and Sorensen (1979) argue that annual earnings and wage rates are the most easily quantified job rewards, this thesis seeks to assess internal hourly wage differences within the Mexican-origin workers for a better understanding of their

heterogeneity. I draw on the literature and theoretical perspectives discussed above to develop a more comprehensive understanding of the labor market outcomes of Mexican immigrants in the key states of California and Texas. In particular, the thesis seeks to assess the extent to which the association between wages and immigration status/citizenship status among persons of Mexican-origin varies across California and Texas. Put simply, I investigate the cost of being a Mexican immigrant, or the hourly wage differences between the native-born and foreign-born, in California and Texas. In addition, among the foreignborn, I estimate the cost of being a Mexican immigrant without U.S. citizenship, or the wage differences between naturalized immigrants and non-naturalized immigrants, in the two states

The thesis examines two sets of hypotheses. First, the following two major hypotheses regarding the costs associated with foreign-born status and the lack of naturalized status are examined.

- H<sub>1</sub>: Foreign-born persons of Mexican-origin have lower hourly wages than native-born Mexican Americans across both states (cost of being an immigrant).
- H<sub>2</sub>: Among foreign-born Mexican workers, those who are not naturalized U.S. citizens have lower hourly wages than U.S. naturalized citizens across both states (cost of being an immigrant non-citizen).

Second, following the literature on the impact of the California social and economic context (established anti-immigrant hostility and laws) on Mexican-origin workers' wages, the above two hypotheses are further specified as follows:

H<sub>3</sub>: The cost of being an immigrant is higher in California than in Texas.

H<sub>4</sub>: The cost of being a non-citizen is higher in California than in Texas.

#### **CHAPTER III**

#### **METHODS**

The hypotheses outlined in the previous chapter are examined using data from the 2000 5% Public-Use Microdata Sample (PUMS). This chapter presents an overview of the methodology used to conduct the analysis. The chapter discusses the data set, operationalization of the variables used in the analysis, analytical models, and the statistical procedures used to conduct the analysis.

#### Data

The data used in the analysis are drawn from the 2000 5-Percent PUMS. The PUMS has state-level Census 2000 data containing individual records for a five percent sample of people and housing units. The PUMS is widely used in earnings studies as it provides a large, nationally representative sample of all sectors of the labor force. The PUMS also ensures adequate sample sizes for the minority populations.

The population from which I draw my sample is persons of Mexican origin (both native- and foreign-born) living in California and Texas at the time of the census. The samples of California and Texas have a total of 221,366 Mexican-origin residents: California, 140,381; and Texas, 80,985. There are a total of 119,740 immigrants in the sample: 84,447 in California; and 35,293 in Texas. Only persons of Mexican origin who worked at least 1,040 hours in 1999 are included in the sample. This restriction ensures that only workers that are attached to the labor force are included in the analysis. Note that persons who worked 1,040 hours in 1999 include full-time workers who were employed half the year as well as half-time workers who were employed year-round. The hourly wage is used instead of the annual

wage income to account for the varying hours that people worked over the course of 1999. Analyses are restricted to individuals within the prime working age range from 16 to 64.

#### Variables

The dependent variable is the natural logarithm of the hourly wage based on earnings in 1999. This transformation procedure is used to minimize the effect of outliers. Because I use the natural logarithm of wages, the coefficients can be interpreted as the percentage change in hourly wages, given a one-unit change in the independent variable. A detailed discussion about the natural logarithm and the logarithmic transformation is presented below.

The first independent variable **imm** measures immigrant status (foreign-born versus native-born status), scored "1" if the person is foreign-born (including both U.S. citizen by naturalization and non-citizen of the United States), and scored "0" if the person is U.S.-born. The construction of this variable allows for the assessment of "the cost of being a foreign-born Mexican," which is the heart of the analysis. The second independent variable **nocit** measures the immigrants' citizenship status, coded "1" if the person is not a naturalized citizen of the United States, and "0" if the person is a U.S. citizen by naturalization. This particular variable allows for the estimation of "the cost of not being a naturalized citizen," which is another major interest of the analysis.

Furthermore, for the analysis of the immigrants' duration of U.S. residence, the immigrant population is partitioned into four cohorts: arrivals in 1990-2000, arrivals in 1980-1989, arrivals in 1970-1979, and immigrants who arrived prior to 1970. Namely, the variable **imm** is replaced with four years-since-migration dummy variables (**imm0069**, **imm7079**, **imm8089**, and **imm9000**) using the variable YR2US (Year of Entry to United States) in the PUMS.

Several control variables are used in the analysis. These include four human capital variables (age, sex, education, and self-reported English-language ability), one residential area variable (metropolitan/nonmetropolitan residence), and two variables based on occupation and self-employment status. The measurement and description of the dependent variable and the independent variables are presented in Table 1.

#### **Analytical Models**

The independent, control, and dependent variables are analyzed using an Ordinary Least Square (OLS) multiple linear regression model due to the interval-level dependent variable. The first two state-specific models (1A for California and 1B for Texas) investigate the cost of being a Mexican immigrant, that is, the hourly wage differences between the native-born and the foreign-born. The second set of models (2A for California and 2B for Texas), based solely on the foreign-born population, analyzes the cost of being a Mexican immigrant without U.S. citizenship, that is, the wage differences between naturalized immigrants and non-naturalized immigrants (see Figure 1).<sup>13</sup>

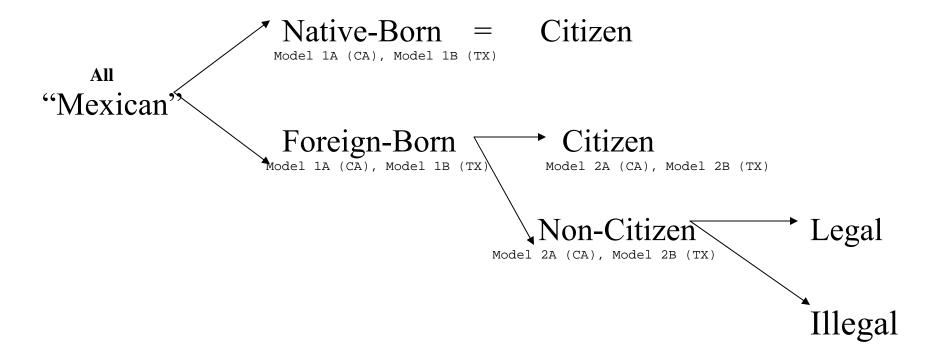
The third set of models (Model 3A for California and Model 3B for Texas) assesses the costs associated with foreign-born status in terms of the four different periods in which immigrants came to the United States. The fourth set of models (Model 4A for California and Model 4B for Texas) estimates the cost associated with the lack of naturalized status controlling for immigrants' length of U.S. residence. The focus on immigrants' length of U.S. residence in models 3 and 4 will allow us to assess the extent to which the initial results (models 1 and 2) hold when length of U.S. residence is taken into account.

<sup>&</sup>lt;sup>13</sup> I refer to the PUMS definition of "citizen" to categorize immigrant or non-immigrant (see Data Dictionary (5-Percent), 7-54, U.S. Census Bureau, Census 2000).

Table 1. Measurement and Description of Dependent Variable and Independent Variables Used in the Analyses of Hourly Wage Differences among Persons of Mexican Origin in California and Texas, by 2000 5% PUMS

| Variable Names              | Measurement and Description of Variables or Categories                                       |
|-----------------------------|--|
| Dependent Variable          | Natural logarithm of hourly wages = Log (annual income/the total working hours in year 1999) |
| Lghrwage                    |  |
| Independent Variables       |  |
| Age                         | Numerical Variable; Age Range: 16-64   |
| Male                        | Dummy Variable; Male = 1; Female = 0   |
| Engabil                     | Dummy Variable; Speak English = 1; does not Speak English = 0                                |
| Metrores                    | Dummy Variable; Living in a Metropolitan Area= 1; Not Living in a Metropolitan Area = 0      |
| Imm                         | Dummy Variable; Immigrant=1; Nonimmigrant=0  |
| Length of U.S. residence fo | or native-born and foreign-born comparison   |
| Imm0069                     | Dummy Variable; Foreign-Born Came to the U.S. before 1970=1; otherwise=0                     |
| Imm7079                     | Dummy Variable; Foreign-Born Came to the U.S. between 1970 and 1979=1; otherwise=0           |
| Imm8089                     | Dummy Variable; Foreign-Born Came to the U.S. between 1980 and 1989=1; otherwise=0           |
| Imm9000                     | Dummy Variable; Foreign-Born Came to the U.S. between 1990 and 2000=1; otherwise=0           |
| (Reference Group=Native     | e-Born)  |
| Noncit                      | Dummy Variable; Not a U.S. Citizen= 1; U.S. Citizen by Naturalization= 0                     |
| Educational Attainment      |  |
| Somehs                      | Dummy Variable; Some High School but no Diploma =1; otherwise=0                              |
| Hsgrad                      | Dummy Variable; High School Graduate =1; otherwise=0   |
| Somecoll                    | Dummy Variable; Some College but no Degree =1; otherwise=0                                   |
| Collgrad                    | Dummy Variable; College Graduate=1; otherwise=0  |
| (Reference Group = Som)     | e Middle School)   |
| Occupation                  |  |
| Mgrprorl                    | Dummy Variable; Management, Professional and Related Occupations=1; otherwise=0              |
| Service                     | Dummy Variable; Service Occupations=1; otherwise=0   |
| Salesoff                    | Dummy Variable; Sales and Office Occupations=1; otherwise=0                                  |
| Cnstexmn                    | Dummy Variable; Construction, Extraction, and Maintenance Occupations=1; otherwise=0         |
| Prtrmtmv                    | Dummy Variable; Production, Transportation, and Material Moving Occupations=1; otherwise=0   |
|                             | ning, Fishing, and Forestry Occupations)   |
| Selfemp                     | Dummy Variable; Selfemployed =1; Not Selfemployed=0  |

Figure 1: Native-Born, Foreign-Born, Citizen, and Non-Citizen



The thesis also analyzes other individual-level determinants of wage differences that are included in the models as controlled variables. For example, reflecting the recent major industrial changes associated with labor market situations of Mexican-origin workers (e.g., polarization of wages from middle-wage to low-or high-wage jobs), the effects of education and professional occupational categories on wages would show significant impacts on hourly wage gaps. Suggestions and discussions of other control variables are also given within the scope of the analysis.

#### **Statistical Procedures**

In this section I explain the statistical reasoning in the background of the natural logarithm and the logarithmic transformation. The natural logarithm of hourly wage is one of the most widely used dependent variables in the statistical analyses of labor market earnings (e.g., Donato and Massey, 1993; Portes and Zhou, 1996; Philip and Massey, 1999; Sakamoto et al., 2000; Tienda, 1983) because (1) hourly wage is highly skewed, and (2) a very useful multiplicative interpretation emerges in terms of the ratio of population medians (Ramsey and Schafer, 1997). Logged data tend to minimize the influence of the extremely low hourly wages of numerous workers. The logarithmic transformation can be expressed in the following formula:

$$\log_b(x) = y \text{ if } b^y = x$$

$$\log_b(x + y) = \log_b(x) \log_b(y)$$

$$\log_{10}(x) \approx 2.30 \log_e(x)$$

$$e \approx 2.718$$

First, a common use of logarithm in statistics is as a data transformation procedure to make samples more normally distributed and variances less unequal. The distribution looks much more symmetric after the log transformation (with less skew and fewer outliers).

Sakamoto *et al.* (2000) note that the distribution of log-wage is concave; it does not need to be normally distributed but has to be symmetric.<sup>14</sup>

Second, the log transformation helps us treat multiplicative effects, which are difficult to handle using classical statistical methods, as additive effects. For example, after the antilog to interpret coefficients, we understand how many times one group earns more (or less) than another group (e.g., male versus female), instead of only getting information about how many dollars one group earns more (or less) than another group.

For ease of interpretation, we antilog to go back to the original scale. The dependent variable does not change throughout the process; it is always "the natural logarithm of hourly wage." However, it does not make sense or it is difficult to interpret the natural logarithm without the antilog because the natural logarithm of hourly wage does not exist in the real world. That is why the data themselves suggest the need for a logarithm transformation. Only after the antilog, are we able to get the ratio of median hourly wage of one group to another group, that is, how many times one group's median hourly wage is more (or less) than that of another group, instead of only getting the information about how many dollars one group earns more (or less) than another group. <sup>15</sup>

# **Diagnostics**

The skewness and distribution of the hourly wages before and after the logarithmic transformation are compared. The original data have the median hourly wage of \$9.62, and \$11.36 standard deviation. The skewness is 8.9, variance is 128.98, and kurtosis is 141.81. The distribution is therefore highly skewed to the right, with a number of low hourly wages. After the logarithmic transformation, on the other hand, the data are more normally

<sup>&</sup>lt;sup>14</sup> Refer to Display 3.2 in Ramsey and Schafer (1997:55).

<sup>&</sup>lt;sup>15</sup> See Display 3.7 in Ramsey and Schafer (1997:66); they explain how the log transformation works.

distributed with fewer outliers. The median hourly wage is 2.26, and the standard deviation is 0.62. The skewness is 0.12, variance is 0.38, and kurtosis is 2.44. Accordingly, these improvements in distribution encourage me to use the natural logarithm of hourly wage as the dependent variable.

Collinearity between the independent variables was also tested. The correlation matrices for the whole sample and the immigrant-only sample show that collinearity is not a major problem, as all correlations are 0.4 or lower. While all the correlations are statistically significant (p < 0.0001), large sample sizes can contribute to small correlations being highly significant (Pett *et al.*, 2003).

Finally, heteroscedasticity tests are conducted using White's test. One of the assumptions for OLS multiple linear regression is that error terms have equal variances. If there is heteroscedasticity, although an OLS regression is still unbiased, it will not be the most efficient. Heteroscedasticity also inflates the standard errors, thus obscuring potentially significant relationships. For each of the eight multiple linear regression models (Model 1 – 4), I tested the null hypothesis of White's test.

$$H_0: \sigma_i^2 = \sigma^2$$
 for all  $i$ 

Diagnostics show that there is no serious heteroscedasticity issue across all of the models in the analysis, as the Chi-Square tests are statistically significant (p < .05). Hence, these are good models with moderate adjusted R-squares (see Tables 3 - 8, 10 and 11). Reflecting the large sample size, almost all p-values of the coefficients are very small (p < 0.0001).

#### CHAPTER IV

#### **FINDINGS**

This chapter presents findings related to the assessment of the cost of being a Mexican immigrant and being a Mexican non-citizen in California and Texas. The chapter consists of various sections. First, I provide an overview of the descriptive statistics associated with the variables used in the analysis. Second, I describe the distribution of the Mexican-origin population along the lines of nativity and naturalization status across the two states. Third, I discuss the relationship between immigrants' duration of stay in the United States and hourly wages. Fourth, I provide a discussion of the multivariate analyses related to the cost of being a Mexican immigrant and being a Mexican non-citizen in California and Texas. Finally, I present a series of supplementary analyses.

# **Descriptive Statistics**

Table 2 illustrates the descriptive statistics based on variables used in the analysis. U.S.-born Mexican Americans (G2) had an average hourly wage that was \$3.61 higher than that of their immigrant counterparts (G3) in California compared to an advantage of only \$1.98 in Texas. Similarly, among foreign-born Mexican workers, U.S. naturalized citizens (G4) in California had a greater average hourly wage advantage (\$3.65) over their non-citizen counterparts (G5) in California than was the case in Texas (\$2.44). As such, the descriptive results suggest that there is a greater cost in being an immigrant and in being a non-citizen immigrant in California than in Texas, with the cost being 1.8 and 1.5 times higher, respectively, in California than in Texas.

Nonetheless, at least part of the hourly wage variations between the native- and foreign-born and between citizen and non-citizen immigrants reflects differences in social

**Table 2. Descriptive Statistics Based on Variables Used in the Analysis** 

|                  |        |        | California | 1      |        |        |        | Texas  |        |        |
|------------------|--------|--------|------------|--------|--------|--------|--------|--------|--------|--------|
| Variables        | G1     | G2     | G3         | G4     | G5     | G1     | G2     | G3     | G4     | G5     |
| Mean Hourly Wage | 12.47  | 14.64  | 11.03      | 13.66  | 10.01  | 11.31  | 12.17  | 10.19  | 11.93  | 9.49   |
| Mean Age         | 35     | 34     | 36         | 40     | 34     | 35.7   | 36     | 36     | 40     | 34     |
| Male             | 63.35% | 54.67% | 69.10%     | 60.93% | 72.28% | 62.42% | 55.18% | 71.79% | 64.80% | 74.59% |
| Engabil          | 71.50% | 97.84% | 54.08%     | 76.28% | 45.42% | 77.21% | 96.91% | 51.72% | 72.66% | 43.35% |
| Metrores         | 87.89% | 87.70% | 88.02%     | 88.44% | 87.86% | 65.35% | 59.57% | 72.84% | 67.33% | 75.04% |
| Education        |        |        |            |        |        |        |        |        |        |        |
| Some High School | 20.87% | 16.12% | 24.01%     | 21.23% | 25.10% | 20.80% | 18.06% | 24.34% | 20.98% | 25.69% |
| HS Graduate      | 22.51% | 29.97% | 17.58%     | 20.15% | 16.58% | 25.35% | 31.93% | 16.84% | 19.97% | 15.58% |
| Some College     | 23.13% | 38.35% | 13.06%     | 21.30% | 9.85%  | 21.88% | 30.46% | 10.78% | 17.21% | 8.20%  |
| College Graduate | 7.20%  | 11.98% | 4.04%      | 7.14%  | 2.83%  | 9.03%  | 12.50% | 4.55%  | 7.34%  | 3.43%  |
| Reference Group  | 26.29% | 3.59%  | 41.31%     | 30.18% | 45.64% | 22.94% | 7.05%  | 43.50% | 34.51% | 47.10% |
| Occupation       |        |        |            |        |        |        |        |        |        |        |
| Mgrprorl         | 14.43% | 23.66% | 8.35%      | 14.63% | 5.90%  | 17.02% | 23.93% | 8.10%  | 13.65% | 5.88%  |
| Service          | 18.66% | 14.37% | 21.49%     | 18.07% | 22.82% | 18.40% | 15.75% | 21.82% | 19.29% | 22.83% |
| Salesoff         | 21.81% | 33.30% | 14.23%     | 19.23% | 12.29% | 21.64% | 29.10% | 12.02% | 16.29% | 10.31% |
| Cnstexmn         | 13.00% | 10.62% | 14.57%     | 12.69% | 15.30% | 18.68% | 12.44% | 26.75% | 19.88% | 29.49% |
| Prtrmtmv         | 25.97% | 16.61% | 32.14%     | 30.55% | 32.76% | 22.17% | 17.76% | 27.88% | 27.97% | 27.84% |
| Reference Group  | 6.13%  | 1.44%  | 9.21%      | 4.83%  | 10.92% | 2.08%  | 1.03%  | 3.44%  | 2.93%  | 3.65%  |
| Selfemp          | 1.96%  | 1.70%  | 2.13%      | 2.51%  | 1.98%  | 2.23%  | 1.90%  | 2.66%  | 3.52%  | 2.31%  |

G1: All Mexican-Origin Workers

G2: All Native-Born G3: All Foreign-Born

G4: Foreign-Born U.S. Naturalized Citizens G5: Foreign-Born without U.S. Citizenship

Source: 2000 5% PUMS.

and economic attributes that are associated with wages. For example, while the majority of native-born Mexican Americans speak English, only slightly more than half of the foreign-born speak English, although naturalized citizens (70%) are more likely to speak English compared to non-citizens (approximately 45%). In addition, the native-born have higher education than their foreign-born counterparts in both states. For instance, the native-born are 2.3 times more likely to be high school graduates compared to their foreign-born counterparts of which only about one-third have a high school diploma. However, among the foreign-born, naturalized citizens are approximately 1.7 times more likely to have completed high school compared to their compatriots who have not become naturalized citizens.

There are also significant differences along nativity lines with respect to the types of jobs that Mexican-origin workers perform. In particular, more than half of the native-born work in higher status occupations (sales and office; and managerial and professional related occupations). In contrast, a significant majority of the foreign-born (76% in Texas and 68% in California) work in three blue-collar occupations (production, transportation, and material moving; construction, extraction, and maintenance; and service occupations). There are some differences, however, among the foreign-born population with naturalized citizens being less likely to be concentrated in these blue-collar occupations compared to those who are not naturalized citizens.

Overall, there are some differences among the Mexican-origin population across the two states. For example, those in California, regardless of nativity/citizenship status, are slightly more likely speak English and to be high school graduates compared to those in Texas. Moreover, the foreign-born are more likely to be employed in construction, extraction, and maintenance occupations compared to their counterparts in California.

# Cost of Being a Mexican Immigrant and Being a Mexican Non-Citizen in California and Texas

The analysis reported above provides preliminary evidence for the hypotheses—namely that immigrants and those that are not naturalized citizens experience wage penalties for their status and that this cost is greater in California than in Texas. However, because of the compositional differences pointed out above, it is necessary to examine the hypotheses using multivariate analysis. Table 3 presents the first set of two multiple linear regression models comparing the difference in predicted log of hourly wages between the foreign-born and native-born Mexican-origin workers in California and Texas. All variables included in the model, except for self-employment (**selfemp**) in the Texas model, are statistically significant.

As hypothesized, being a Mexican immigrant (**imm**) is disadvantageous with respect to earnings in both states, and the cost of being a foreign-born worker is higher in California than in Texas. Specifically, in the California model, it is estimated that the median hourly wage of a given foreign-born worker is 0.9399 times as high as that of a given native-born worker. Thus, the median hourly wage of foreign-born workers is estimated to be 6 percent lower than that of their native-born peers. In the Texas model, on the other hand, the median hourly wage of a given foreign-born worker is estimated to be only 2 percent lower than that of a given native-born worker. The relative nativity gap is therefore three times larger in California. Although the issue of uncontrolled differences in wages and costs of living across the states remains, the between-group difference in the costs of being an immigrant is statistically significant:

$$t = \beta_1 - \beta_2 / \sqrt{\sigma_1^2 + \sigma_2^2}$$
  
= 0.9826 - 0.9399 /  $\sqrt{0.0047^2 + 0.0035^2} = 7.30 > 1.96$  (p < .05, two-tailed)

Table 3. The Cost of Being a Mexican Immigrant in California and Texas: Multiple Linear Regression Results for the Dependent Variable "the Natural Logarithm of the Hourly Wage"

 $\begin{array}{c} \textbf{Model 1B: TX Whole} \\ \textbf{Model 1A: CA Whole Sample} \\ N=140{,}381 & N=80{,}985 \end{array}$ 

|            |              |           |        |   |             | Std.   |       |
|------------|--------------|-----------|--------|---|-------------|--------|-------|
| Variable   | Exp (Coef.)  | Std. Err. | t      | _ | Exp (Coef.) | Err.   | t     |
| Age        | 1.0151****   | 0.0001    | 109.99 |   | 1.0114****  | 0.0002 | 66.79 |
| Male       | 1.2144****   | 0.0032    | 60.88  |   | 1.2409****  | 0.0043 | 50.64 |
| Engabil    | 1.1712****   | 0.0038    | 41.21  |   | 1.1712****  | 0.0054 | 20.36 |
| Metrores   | 1.04656****  | 0.0045    | 10.15  |   | 1.0970****  | 0.0039 | 23.6  |
| Imm        | 0.9399****   | 0.0035    | -17.62 |   | 0.9826***   | 0.0047 | -3.78 |
| Education  |              |           |        |   |             |        |       |
| Somehs     | 1.0497****   | 0.0044    | 10.97  |   | 1.0375****  | 0.0059 | 6.26  |
| Hsgrad     | 1.739****    | 0.0047    | 34.38  |   | 1.1617****  | 0.0061 | 24.79 |
| Somecoll   | 1.3370****   | 0.005     | 58.67  |   | 1.3224****  | 0.0066 | 42.59 |
| Collgrad   | 1.6839****   | 0.007     | 74.21  |   | 1.7924****  | 0.0086 | 66.72 |
| Occupation |              |           |        |   |             |        |       |
| Mgrpror1   | 1.5731****   | 0.0077    | 58.79  |   | 1.5786****  | 0.0138 | 32.98 |
| Service    | 1.1271****   | 0.0068    | 17.54  |   | 1.1094****  | 0.0132 | 7.88  |
| Salesoff   | 1.3344***    | 0.0071    | 40.4   |   | 1.3340****  | 0.0135 | 21.43 |
| Cnstexmn   | 1.4670****   | 0.0072    | 53.16  |   | 1.3936****  | 0.0131 | 25.36 |
| Prtrmtmv   | 1.2705****   | 0.0066    | 36.39  |   | 1.3728****  | 0.013  | 24.44 |
| Selfemp    | 0.9550****   | 0.0102    | -4.5   |   | 1.011       | 0.0124 | 0.89  |
| Constant   | 3.127769**** | 0.0092    | 123.54 |   | 3.0009****  | 0.0156 | 70.39 |

Source: 2000 5% PUMS.

<sup>\*\*\*\*</sup>Significant at the .0001 level;

<sup>\*\*\*</sup>Significant at the .0005 level.

CA Model Adj R-Sq = .2720; TX Model Adj R-Sq = .2503

As can be seen, being an immigrant is a major source of the internal wage gap, and social and economic contexts matter with respect to the labor market outcomes of Mexican immigrants. I argue that the more hostile environment that immigrants face in California than in Texas—represented in a series of California propositions and related anti-immigrant sentiments—contribute to the greater wage penalty that immigrants experience in California.

Table 4 presents the second set of multiple linear regression models comparing the costs of being a Mexican non-citizen between the two states. As hypothesized, being a non-citizen significantly drops foreign-born workers' predicted hourly wages across the states, and this cost is again higher in California than in Texas. In California, the median hourly wage of a given Mexican non-citizen worker is estimated to be 14 percent lower than that of a given naturalized-citizen worker. In Texas, on the other hand, a given non-citizen worker's median hourly wage is estimated to be 10 percent lower than that of a given naturalized-citizen worker. The relative citizenship gap is therefore 1.4 times larger in California. The cross-state difference in the cost associated with non-naturalized status is statistically significant, although cross-state differences in wages and costs of living are not controlled:

$$t = \beta_1 - \beta_2 / \sqrt{\sigma_1^2 + \sigma_2^2}$$

= 
$$0.8998 - 0.8581 / \sqrt{0.0044^2 + 0.0044^2} = 6.73 > 1.96$$
 (p < .05, two-tailed)

The "dual costs" of being a Mexican immigrant without U.S. citizenship are indicated in the literature. In the post-IRCA period, legal status—rather than human capital attributes—has emerged as the primary determinant of immigrants' wage rates (Davila *et al.*, 1998; Donato and Massey, 1993; Donato *et al.*, 1992a; Massey, 1987b; Philip and Massey, 1999).

To supplement the argument of wage gaps based on nativity difference and the possession of U.S. citizenship, four additional models with a California dummy variable

Table 4. The Cost of Being a Mexican Non-Citizen in California and Texas: Multiple Linear Regression Results for the Dependent Variable "the Natural Logarithm of the Hourly Wage"

Model 2A: CA Immigrant-Only Sample (N= 84 447)

Model 2B: TX Immigrant-Only Sample (N=35.293)

| (11-04,447) |             |           |        | (IN-33,293) |           |        |  |
|-------------|-------------|-----------|--------|-------------|-----------|--------|--|
| Variable    | Exp (Coef.) | Std. Err. | t      | Exp (Coef.) | Std. Err. | t      |  |
| Age         | 1.0101****  | 0.0002    | 53.96  | 1.0067****  | 0.0003    | 24.27  |  |
| Male        | 1.2405****  | 0.0041    | 52.68  | 1.2781****  | 0.0068    | 35.97  |  |
| Engabil     | 1.1548****  | 0.004     | 35.94  | 1.1176****  | 0.006     | 18.45  |  |
| Metrores    | 1.0197**    | 0.0058    | 3.37   | 1.0758****  | 0.0064    | 11.49  |  |
| Nocit       | 0.8581****  | 0.0044    | -35.11 | 0.8998****  | 0.0044    | -15.92 |  |
| Education   |             |           |        |             |           |        |  |
| Somehs      | 1.0377****  | 0.0048    | 7.74   | 1.0217*     | 0.0072    | 2.96   |  |
| Hsgrad      | 1.1305****  | 0.0054    | 22.61  | 1.1027****  | 0.0084    | 11.67  |  |
| Somecoll    | 1.2592****  | 0.0063    | 36.81  | 1.1745****  | 0.0102    | 15.73  |  |
| Collgrad    | 1.4571****  | 0.0101    | 37.34  | 1.4741****  | 0.0151    | 25.7   |  |
| Occupation  |             |           |        |             |           |        |  |
| Mgrprorl    | 1.5657****  | 0.0095    | 47.00  | 1.6501****  | 0.019     | 26.38  |  |
| Service     | 1.1152****  | 0.0074    | 14.72  | 1.1147****  | 0.0164    | 6.62   |  |
| Salesoff    | 1.3225****  | 0.0083    | 33.56  | 1.3477***   | 0.0177    | 16.84  |  |
| Cnstexmn    | 1.4513****  | 0.0079    | 46.92  | 1.3928****  | 0.0161    | 20.57  |  |
| Prtrmtmv    | 1.2611****  | 0.0071    | 32.77  | 1.3909****  | 0.016     | 20.67  |  |
| Selfemp     | 0.9582**    | 0.0124    | -3.45  | 1.0656***   | 0.0172    | 3.69   |  |
| Constant    | 4.0905****  | 0.0116    | 121.17 | 3.8044***   | 0.0213    | 62.85  |  |

Source: 2000 5% PUMS.

CA Model Adj R-Sq = .2243; TX Model Adj R-Sq=0.1848.

<sup>\*\*\*\*</sup>Significant at the .0001 level; \*\*\*Significant at the .0005 level;

<sup>\*\*</sup>Significant at the .001 level; \*Significant at the .005 level;

(California resident = "1"; Texas resident = "0") are estimated for four subgroups (native-born, foreign-born, foreign-born with U.S. citizenship, and foreign-born without U.S. citizenship) (Table 5, 6, 7, and 8). Largely due to the higher cost of living in California, all four groups have a "California wage advantage" (see Figure 2). However, the fact that the California wage advantage is 1.7 times larger in the case of the native-born than that of the foreign-born supports the argument regarding the more disadvantageous position of immigrants in California. The same is true with respect to naturalized citizenship, where the California wage advantage is twice greater among the naturalized citizens compared to those lacking this status. These four additional models provide further support to the findings from the initial analysis that being an immigrant and non-citizen reduce Mexican-origin workers' wages, and that these penalties are greater in California than in Texas.

Nevertheless, the analysis is based on an assumption that California and Texas have equal living expenses and state minimum wages. Considering that the analysis does not control differences in cost of living and state minimum wages (\$6.75 in California and \$5.15 in Texas), the analysis may appear to violate the assumption. However, although I acknowledge the problem, I do not dismiss these latest results due to a long controversy among labor market researchers regarding these two components (living expenses and state minimum wages). First, researchers have not agreed on the definitions of living expenses (e.g., auto, public transportation, air conditioner, and heater) as most living expenses are fluid. Second, cross-state wage differences are naturally difficult to take into account. For instance, those in California are more likely to work in service industries while those in Texas are more likely to work in construction. Furthermore, considering the fact that many of Mexican immigrants are involved in food service industries and agriculture, state minimum

Table 5. The Hourly Wage Difference of Native-Born Mexican American Workers in California and Texas: Multiple Linear Regression Results for the Dependent Variable "the Natural Logarithm of the Hourly Wage"

| Variable   | Exp (Coef.) | Std. Err. | t      |
|------------|-------------|-----------|--------|
| Age        | 1.0166****  | 0.0001    | 109.99 |
| Male       | 1.2018****  | 0.0032    | 60.88  |
| Engabil    | 1.0638****  | 0.0038    | 41.21  |
| Metrores   | 1.1031****  | 0.0045    | 10.15  |
| CA         | 1.1715****  | 0.0035    | 17.62  |
| Education  |             |           |        |
| Somehs     | 1.1177****  | 0.0044    | 10.97  |
| Hsgrad     | 1.2928****  | 0.0047    | 34.38  |
| Somecoll   | 1.4943****  | 0.005     | 58.67  |
| Collgrad   | 2.0160****  | 0.007     | 74.21  |
| Occupation |             |           |        |
| Mgrpror1   | 1.6060****  | 0.0077    | 58.79  |
| Service    | 1.1906****  | 0.0068    | 17.54  |
| Salesoff   | 1.3917****  | 0.0071    | 40.4   |
| Cnstexmn   | 1.5051****  | 0.0072    | 53.16  |
| Prtrmtmv   | 1.3849****  | 0.0066    | 36.39  |
| Selfemp    | 0.9577***   | 0.0102    | -4.5   |
|            |             |           |        |
| Constant   | 2.3065****  | 0.0092    | 123.54 |

\*\*\*\*Significant at the .0001 level.

Adj R-Sq = .2923.

N=101,626.

\*\*\*Significant at the .001 level.

Table 6. The Hourly Wage Difference of Foreign-Born Mexicans in California and Texas: Multiple Linear Regression Results for the Dependent Variable "the Natural Logarithm of the Hourly Wage"

| Variable   | Exp (Coef.) |        | Std. Err. | t      |
|------------|-------------|--------|-----------|--------|
| Age        | 1.0107****  |        | 0.0001    | 70.98  |
| Male       | 1.2411****  |        | 0.0035    | 61.29  |
| Engabil    | 1.1754***   |        | 0.0033    | 49.19  |
| Metrores   | 1.0400****  |        | 0.0043    | 9.11   |
| CA         | 1.1022****  |        | 0.0034    | 28.32  |
| Education  |             |        |           |        |
| Somehs     | 1.0412****  |        | 0.004     | 10.06  |
| Hsgrad     | 1.1376****  |        | 0.0046    | 28.17  |
| Somecoll   | 1.2664****  |        | 0.0053    | 44.19  |
| Collgrad   | 1.4925****  |        | 0.0084    | 47.5   |
| Occupation |             |        |           |        |
| Mgrprorl   | 1.6099****  |        | 0.0084    | 56.5   |
| Service    | 1.1120****  |        | 0.0067    | 15.86  |
| Salesoff   | 1.3392****  |        | 0.0075    | 39.13  |
| Cnstexmn   | 1.4305****  |        | 0.0069    | 51.67  |
| Prtrmtmv   | 1.2981****  |        | 0.0064    | 40.72  |
| Selfemp    |             | 0.9898 | 0.0101    | -1.01  |
| Constant   | 3.1245****  |        | 0.0092    | 124.53 |

N=119,740.

<sup>\*\*\*\*</sup>Significant at the .0001 level.

Adj R-Sq = .2028.

Table 7. The Hourly Wage Difference of Foreign-Born with Citizenship in California and Texas: Multiple Linear Regression Results for the Dependent Variable "the Natural Logarithm of the Hourly Wage"

| Variable   | Exp (Coef.) |        | Std. Err. | t     |
|------------|-------------|--------|-----------|-------|
| Age        | 1.0079****  |        | 0.0003    | 26.2  |
| Male       | 1.3150****  |        | 0.0064    | 42.61 |
| Engabil    | 1.1499****  |        | 0.0072    | 19.28 |
| Metrores   | 1.0779****  |        | 0.0079    | 9.46  |
| CA         | 1.1621****  |        | 0.0065    | 23.05 |
| Education  |             |        |           |       |
| Somehs     | 1.0503****  |        | 0.0084    | 5.86  |
| Hsgrad     | 1.1515****  |        | 0.0088    | 16.05 |
| Somecoll   | 1.3140****  |        | 0.0093    | 29.51 |
| Collgrad   | 1.6465****  |        | 0.0135    | 36.99 |
| Occupation |             |        |           |       |
| Mgrprorl   | 1.6109****  |        | 0.0172    | 27.8  |
| Service    | 1.1427****  |        | 0.0158    | 8.42  |
| Salesoff   | 1.3799****  |        | 0.0164    | 19.63 |
| Cnstexmn   | 1.5075****  |        | 0.0162    | 25.29 |
| Prtrmtmv   | 1.3685****  |        | 0.0152    | 20.67 |
| Selfemp    |             | 0.9823 | 0.0174    | -1.03 |
| Constant   | 3.3050****  |        | 0.0208    | 57.62 |

Adj R-Sq = .2278.

N=33,767.

<sup>\*\*\*\*</sup>Significant at the .0001 level.

Table 8. The Hourly Wage Difference of Foreign-Born without Citizenship in California and Texas: Multiple Linear Regression Results for the Dependent Variable "the Natural Logarithm of the Hourly Wage"

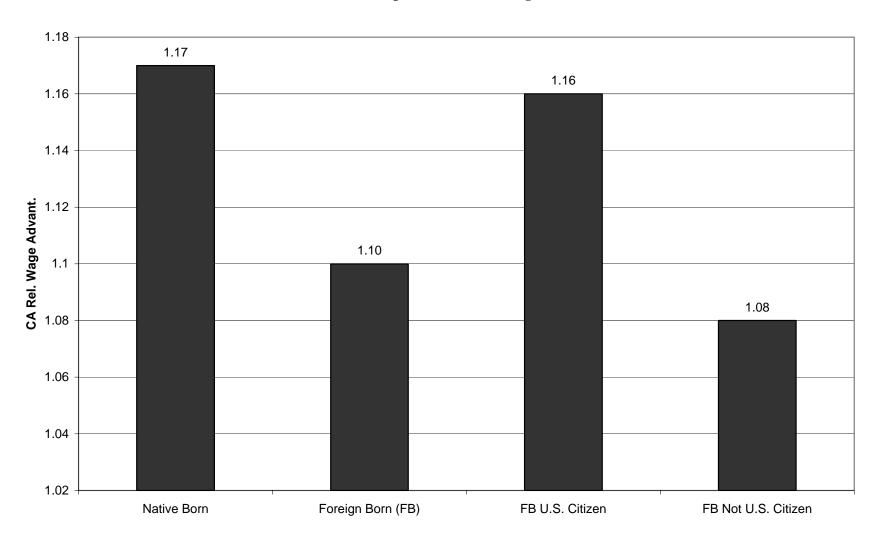
| Variable   | Exp (Coef.) |        | Std. Err. | t      |
|------------|-------------|--------|-----------|--------|
| Age        | 1.0095****  |        | 0.0002    | 52.54  |
| Male       | 1.2234****  |        | 0.0042    | 48.16  |
| Engabil    | 1.1419****  |        | 0.0038    | 35.38  |
| Metrores   | 1.0284****  |        | 0.0051    | 5.51   |
| CA         | 1.0793****  |        | 0.004     | 19.07  |
| Education  |             |        |           |        |
| Somehs     | 1.0311****  |        | 0.0045    | 6.76   |
| Hsgrad     | 1.1177****  |        | 0.0054    | 20.82  |
| Somecoll   | 1.2004****  |        | 0.0067    | 27.24  |
| Collgrad   | 1.3493****  |        | 0.011     | 27.2   |
| Occupation |             |        |           |        |
| Mgrprorl   | 1.5898****  |        | 0.0102    | 45.67  |
| Service    | 1.1076****  |        | 0.0073    | 13.94  |
| Salesoff   | 1.3139****  |        | 0.0085    | 32.16  |
| Cnstexmn   | 1.4056****  |        | 0.0076    | 44.7   |
| Prtrmtmv   | 1.2726****  |        | 0.007     | 34.32  |
| Selfemp    |             | 0.9972 | 0.0124    | -0.23  |
| Constant   | 3.3846****  |        | 0.0105    | 116.51 |

N=85,972.

<sup>\*\*\*\*</sup>Significant at the .0001 level.

Adj R-Sq = .1537.

Figure 2. California Relative Wage Advantage Over Texas from OLS Regressions by Nativity and the U.S. Citizenship Status of the Foreign Born



wages do not apply to all of the Mexican-origin workers. Third, there are a number of regional variations within the states as both California and Texas are relatively large states. Fourth, my interest is on comparing the California-Texas wage gaps between the native- and foreign-born and the naturalized citizens and non-naturalized citizens with the understanding that cost of living and minimum wage differences between the states are fairly consistent across models (subgroups). Hence, this last set of findings does have validity.

As increases in the supply of immigrant labor force have a strong impact on the earnings of immigrants themselves (especially those without U.S. citizenship) (Borjas, 1987a, 1987b, 1994, 1995; Topel, 1994), frequency distributions of Mexican immigrants are also taken into account (see Table 9). First, the foreign-born represent a much larger share of the Mexican-origin population in California than in Texas. Three-fifths (60.2%) of the Mexican-origin population in California is foreign-born, whereas nearly three-fifths (56.4%) in Texas are native-born. Second, while there are 1.22 native-born Mexican Americans in California to every one native-born Mexican American in Texas, the respective ratio is 2.39 with respect to the foreign-born population. Third, the size of the California immigrant population is larger than that of the Texas immigrant population regardless of period of U.S. entry and naturalization status. However, it is clear that among the Mexican immigrant population, those that first immigrated to the United States since 1990 (the most recent immigrants) account for a larger share of the Texas immigrant population (38.7%) compared to the California immigrant population (29.9%). This perhaps signifies a change in the destination of Mexican immigrants away from California and toward Texas.

The higher costs associated with foreign-born status and the lack of naturalization status in California appear to reflect the harsher labor market conditions in which increases in

Table 9. Distribution of the Mexican-Origin Population Used in the Study Sample by Nativity and Period of Entry and Naturalization Status among the Foreign-Born by State

| Nativity Status:                                 | California | Texas  | California-Texas<br>Population Ratio |
|--|------------|--------|--------------------------------------|
| U.SBorn  | 39.80%     | 56.40% | 1.22                                 |
| Foreign-Born                                     | 60.20%     | 43.60% | 2.39                                 |
| Total Mexican-Origin Population                  | 140,598    | 81,063 | 1.73                                 |
| Foreign-Born Population by Period of U.S. Entry: |            |        |                                      |
| Came to the U.S. < 1970                          | 8.60%      | 8.00%  | 2.57                                 |
| Came to the U.S. in 1970-1979                    | 24.90%     | 21.40% | 2.78                                 |
| Came to the U.S. in 1980-1989                    | 36.70%     | 31.90% | 2.75                                 |
| Came to the U.S. in 1990-2000                    | 29.90%     | 38.70% | 1.85                                 |
| Total Foreign-Born Population                    | 84,616     | 35,331 | 2.39                                 |
| Foreign-Born Population Naturalization Status:   |            |        |                                      |
| U.S. Naturalized Citizens                        | 28%        | 28.60% | 2.35                                 |
| Not U.S. Naturalized Citizens                    | 72%        | 71.5%  | 2.41                                 |
| Total Foreign-Born Population                    | 84,616     | 35,331 | 2.39                                 |

the supply of immigrants, especially those who arrived after 1990, may have contributed to a lowering of the average hourly wage of the Mexican immigrant population in California. On the other hand, the smaller costs based on the immigrant status and the lack of U.S. citizenship in Texas may be partly explained by the smaller immigrant population which suggests a less intensive wage competition among the foreign-born in Texas.

# Length of U.S. Residence and the Cost of Being a Mexican Immigrant and Non-Citizen

The literature notes that immigrants' wages differ depending on their length of stay in the United States, because immigrants generally improve their wage returns to their human capital with time spent in this country. Furthermore, the literature suggests that the social and economic configurations of immigrants differ with respect to the period in which immigrants come to the United States. Thus, in addition to the costs associated with foreign-born status and the lack of naturalized status, the time when a Mexican worker came to the United States needs to be taken into account.

Table 10 examines variations in costs associated with foreign-born status by four periods of entry into the United States. The results show two important patterns. First, the most recent immigrants—those who arrived during the last decade (**imm9000**)—had the lowest wages relative to the native-born, all else equal, in both states. Second, the most recent immigrants in California had hourly wages that were 14 percent below their native-born counterparts, with the respective difference being only 6 percent in Texas. The 2.3 times higher relative gap of **imm9000** in California than **imm9000** in Texas (14% to 6%) is statistically significant, although cross-state wage differences are not controlled (t = 9.51 > 1.96, p < .05, two-tailed).

Table 10. The Cost of Being a Mexican Immigrant in California and Texas: Multiple Linear Regression Results for the Dependent Variable "the Natural Logarithm of the Hourly Wage": With Length of U.S. Residence

**Model 3A: CA Whole Sample Model 3B: TX Whole Sample** N = 140,381N = 80.985Variable Exp (Coef.) Exp (Coef.) Std. Err. t Std. Err. t  $1.01\overline{10^{****}}$ 1.0136\*\*\*\* 0.0002 88.84 0.0002 60.13 Age 1.2420\*\*\*\* 0.0043 50.89 Male 1.2170\*\*\*\* 0.0032 61.67 0.004 33.12 1.0998\*\*\*\* 0.0056 17.1 Engabil 1 1407\*\*\*\* 1.0486\*\*\*\* 1.0985\*\*\*\* 23.95 Metrores 0.0045 10.59 0.0039 Length of U.S. Residence 0.007 -3.15-4.98 Imm0069 0 9782\*\* 0 9493\*\*\*\* 0.0104 0.9835\*\*\* 0.0069 2.95 Imm7079 0.0047 -3.541.0206\*\* Imm8089 0.9426\*\*\*\* 0.0043 -13.81 1.0068 0.0061 1.1 Imm9000 0.8621\*\*\*\* 0.0049 -30.1 0.9357\*\*\*\* 0.0064 -10.33 Education Somehs 1 0535\*\*\*\* 0.0044 11.79 1.0417\*\*\*\* 0.0059 6.94 1.1779\*\*\*\* 25.34 Hsgrad 0.0047 35.14 1.1656\*\*\*\* 0.0061 0.005 59.02 1.3265\*\*\*\* 0.0066 43.07 Somecoll 1 3391\*\*\*\* Collgrad 1.6954\*\*\*\* 0.007 75.27 1.8045\*\*\*\* 0.0088 67.44 **Occupation** Mgrpror1 58.08 33.1 1 5633\*\*\*\* 0.0077 1.5803\*\*\*\* 0.0138 Service 1.1274\*\*\*\* 0.0068 17.61 1.1130\*\*\*\* 0.0132 8.12 1.3347\*\*\*\* Salesoff 0.0071 39.41 21.5 1 3245\*\*\*\* 0.0134 0.0072 0.0131 25.59 1.4621\*\*\*\* 52.8 1.3974\*\*\*\* Cnstexmn 1.2633\*\*\*\* **Prtrmtmv** 0.0066 1.3731\*\*\*\* 0.013 24.49 35.59 0.0102 -4.53 1.009 0.0124 0.72 Selfemp 0 9548\*\*\*\* 3.3729\*\*\*\* 0.0098 123.94 3.0678\*\*\*\* Constant 0.016 70.19

Source: 2000 5% PUMS.

<sup>\*\*</sup>Significant at the .005 level; \*\*\*Significant at the .0005 level; \*\*\*\*Significant at the .0001 level.

CA Model Adj R-Sq = .2753; TX Model Adj R-Sq = .2520

Their highest cost of being an immigrant who arrived since 1990 can be partly accounted for by their shorter duration of stay in the U.S. and lower levels of human capital (e.g., English language proficiency and U.S. labor market experiences), although the multivariate analysis takes such compositional differences into account. Furthermore, the large cost among the most recent immigrants may be at least partly associated with the nation wide anti-immigrant sentiments leading initially to the passage of IRCA and to further animus against immigrants especially in California in the 1990s. It is also likely that the most recent arrivals are also less likely to be naturalized citizens, an increasingly debilitating factor in wage attainment in the post-IRCA period.

The patterns are relatively different for those who have been in the country longer.

Mexican immigrants who arrived before 1990 display relatively lower costs for their immigrant status, suggesting that a combination of their human capital improvement with duration of stay in the United States and possibly higher rates of naturalization enabled them to reduce their wage penalties attached to their immigrant status.

In California, foreign-born Mexicans who arrived before 1970 (**imm0069**) and between 1970 and 1979 (**imm7079**) had hourly wages that were only 2 percent lower than those of their native-born counterparts. In Texas, on the other hand, **imm0069** and **imm7079** are associated with 5% percent lower and 2 percent higher wages, respectively. This finding appears to suggest shifts in the Mexican labor market core from Texas to California. The literature notes that compared to California, Mexican immigrant workers in Texas faced a severe internal wage competition before 1970. Starting from the bracero program between 1942 and 1964, immigrant workers had lower wages than the native-born, especially in agricultural industries (Jenkins, 1977; cited in Valdes, 1995; Sandos and Cross, 1983;

Valdes, 1995). The cross-state differences in **imm0069** (t = 2.37 > 1.96, p < .05, two-tailed) and **imm7079** (t = 4.42 > 1.96, p < .05, two-tailed) are also statistically significant, although between-state wage differences are not controlled.

On the other hand, foreign-born Mexicans who arrived California between 1980 and 1989 (imm8089) had 6 percent lower wages than those of their native-born counterparts. It is likely that the higher cost of being an immigrant attached to this cohort than the costs attached to earlier arrivals (imm0069 and imm7079) reflects the inception of the harsher treatment toward Mexican immigrants in the state, which led to the implementation of anti-immigrant propositions during the last decade and accordingly immigrant status became a major determinant of internal wage differences among the Mexican-origin workers.

The last set of models (Table 11) assess whether the significance of length of stay in the United States holds for non-citizens as well. Needless to say, length of stay in the United States and citizenship status are positively correlated. While the initial analysis (Model 2A and 2B) does not control immigrants' duration of U.S. residence, Model 4A and 4B present the cost of being a non-citizen controlling immigrants' length of U.S. residence.

While the cost of being a non-citizen is associated with 14 percent lower hourly wages in California and 10 percent lower wages in Texas in Model 2A and 2B, the foreign-born without U.S. citizenship have 11 percent and 8 percent lower wages in California and Texas, respectively, compared to the naturalized citizens in Model 4A and 4B, holding immigrants' length of stay in the U.S. constant. The cross-state difference in the **nocit** variable is statistically significant, although between-state wage differences are not controlled (t = 4 > 1.96, p < .05, two-tailed).

Table 11. The Cost of Being a Mexican Non-Citizen in California and Texas: Multiple Linear Regression Results for the Dependent Variable "the Natural Logarithm of the Hourly Wage": Length of U.S. Residence Controlled

|                             | <b>Model 4A: CA Immigrant-Only Sample</b> N= 84,447 |             |        | <b>Model 4B: TX Immigrant-Only Sample</b> N= 35,293 |             |        |  |
|-----------------------------|---|-------------|--------|---|-------------|--------|--|
| Variable                    | Exp (Coef.)   | Std. Err. t |        | Exp (Coef.)   | Std. Err. t |        |  |
| Age                         | 1.0060**  | 0.0002      | 26.77  | 1.0044**  | 0.0003      | 13.62  |  |
| Male                        | 1.2433**  | 0.0041      | 53.58  | 1.2816**  | 0.0068      | 36.49  |  |
| Engabil                     | 1.1177**  | 0.0041      | 27.19  | 1.0902**  | 0.0062      | 13.93  |  |
| Metrores                    | 1.0253**  | 0.0058      | 4.26   | 1.0799**  | 0.0063      | 12.12  |  |
| Cost of Being a Non-Citizen |   |             |        |   |             |        |  |
| Nocit                       | 0.8884**  | 0.0045      | -26.54 | 0.9220**  | 0.0068      | -11.92 |  |
| Imm7079                     | 0.9565**  | 0.0073      | -6.1   | 1.02  | 0.0118      | 1.68   |  |
| Imm8089                     | 0.8905**  | 0.0077      | -15.06 | 0.9797  | 0.0121      | -1.71  |  |
| Imm9000                     | 0.7926**  | 0.0087      | -26.6  | 0.8902**  | 0.0132      | -8.79  |  |
| Education                   |   |             |        |   |             |        |  |
| Somehs                      | 1.0369**  | 0.0048      | 7.62   | 1.0253*   | 0.0072      | 3.46   |  |
| Hsgrad                      | 1.1302**  | 0.0054      | 22.67  | 1.1034**  | 0.0084      | 11.78  |  |
| Somecoll                    | 1.2517**  | 0.0062      | 36     | 1.1764**  | 0.0102      | 15.93  |  |
| Collgrad                    | 1.4734**  | 0.01        | 38.66  | 1.5067**  | 0.0151      | 27.14  |  |
| Occupation                  |   |             |        |   |             |        |  |
| Mgrprorl                    | 1.5411**  | 0.0095      | 45.58  | 1.6432**  | 0.0189      | 26.25  |  |
| Service                     | 1.1197**  | 0.0074      | 15.36  | 1.1211**  | 0.0164      | 6.99   |  |
| Salesoff                    | 1.3051**  | 0.0083      | 32.16  | 1.3418**  | 0.0177      | 16.65  |  |
| Cnstexmn                    | 1.4452**  | 0.0079      | 46.7   | 1.3940**  | 0.0161      | 20.7   |  |
| Prtrmtmv                    | 1.2538**  | 0.007       | 32.15  | 1.3888****  | 0.0159      | 20.66  |  |
| Selfemp                     | 0.9582*   | 0.0123      | -3.47  | 1.0634*   | 0.0172      | 3.59   |  |
| Constant                    | 5.2982**  | 0.0156      | 106.82 | 4.2742**  | 0.0264      | 55.1   |  |

<sup>\*</sup>Significant at the .0005 level; \*\*Significant at the .0001 level.

CA Model Adj R-Sq = .2348; TX Model Adj R-Sq = .1912

The slight reduction in the cost associated with a non-citizen status in Model 4 is accounted for by the fact that length of U.S. residence can decrease the cost of being a non-citizen. The significance of duration of stay in the United States is supported by the fact that three controlled variables, **imm7079**, **imm8089**, and **imm9000**, are all statistically significant. This pattern reflects the fact that immigrants' longer duration of stay in the U.S. has a combined effect of higher naturalization rates and human capital improvement (e.g., English language proficiency and more U.S. labor market experiences). Table 11 shows that length of U.S. residence is important for non-citizens to reduce the wage penalty associated with their lack of U.S. citizenship status.

# **Human Capital Attributes and Other Factors**

Model 1 and Model 2 strongly suggest that high levels of human capital have much to do with higher expected wages earned by Mexican-origin workers. Age is positively associated with hourly wages, as age is regarded as a substitute factor measuring working experience. English language proficiency is strongly and positively associated with hourly wages. English language ability is particularly important for immigrants, as it is correlated with levels of education.

Hourly wages are also consistently associated with educational attainment. However, among immigrants, the location where the education took place is important. Indeed, the effect of education on wages is highly associated with nativity status. Namely, wage returns to education are consistently smaller in the immigrant-only sample (Model 2A and 2B) than in the whole sample (Model 1A and Model 1B), suggesting that schooling has a smaller impact on the earnings of the foreign-born than on those of the native-born (Chiswick *et al.*, 1997). This is particularly the case regarding the two college-related variables.

There is a wide range of literature addressing the skill-based technological change in which the demand for low-skill workers has fallen relative to workers with high-level skills (McCall, 2000b). The globalization of the U.S. economy is accompanied by changes in the U.S. wage structure in which workers with less education face lower wages and earnings (Borjas, 1999). Thus, college completion and more advanced studies (e.g., professional schools) are likely to be increasingly crucial determinants of the socioeconomic attainment of Mexican-origin workers.

The non-human capital variables are also significantly related to hourly wages. For example, males consistently have higher wages than their female counterparts across the states. However, the gender wage gap is greater among the foreign-born. The literature points out gender-related structural barriers in the Mexican-origin population, due to the numerical predominance of males (Marcelli and Cornelius, 2001; Roberts *et al.*, 1999; Wallace, 1986). In fact, Saenz (2004a) reports a gender gap in labor force participation rates among Mexicans; while 71.4 percent of foreign-born males were in the labor force in 2000, only 47.7 percent of females were part of the labor force.

In addition, metropolitan area residents attain higher wages than those living in nonmetropolitan areas both in the first and second sets of models. However, wage returns to metropolitan area residence are higher in Texas than in California. This finding appears to reflect the higher population concentration of Mexican-origin workers in metropolitan areas in California (Wallace, 1986). In fact, research suggests that traditional metropolitan areas in California (e.g., Los Angeles) have been facing labor market saturation (Durand *et al.*, 2000; Krissman, 2000; Suro and Singer, 2002).

Five occupational categories are also included in the analysis. Wage returns are highest for those who work for professional, management and related occupations (**mgrprorl**) both in the whole Mexican-origin sample and immigrant-only sample. The literature suggests possible bifurcated wages between a relatively small portion of Mexican-origin workers that have specialized skills (Alarcon, 1999; Castells, 1996; Clark, 1998; Keely, 1974; Kritz, 1987; Pellegrino, 2001; Roberts *et al.*, 1999)<sup>16</sup> and disproportionate number of workers in the low-skilled and low-wage labor force (Massey and Schnabel, 1983; Portes and Truelove, 1987; Roos and Hennessy, 1987; Saenz and Torres, 2003; Waldinger, 1989a, 1989b, 1996).<sup>17</sup>

Wage returns to sales and office occupations (salesoff) are higher than service occupations (service) across the models. It is likely that sales and office occupations require higher levels of human capital, such as English-language proficiency and education. Wage returns for construction, extraction, and maintenance occupations (cnstexmn) are higher in California, but wage returns for production, transportation, and material moving occupations (prtrmtmv) are higher in Texas. Furthermore, the median hourly wages in cnstexmn are estimated to be higher than service and salesoff variables. There is only slight statistical evidence for the advantage of self-employment. As Borjas (1986) and Spencer and Bean (1999) point out, Mexican self-employment tends to be associated with disadvantaged labor market outcomes. However, one possible explanation for this observation is that the self-employed tend to work many hours which brings down their hourly wages.<sup>18</sup>

<sup>&</sup>lt;sup>16</sup> Alarcon (1999) notes the arrival of a large number of highly educated permanent residents and temporary workers during the early 1990s. Alarcon examines the processes by which Indian and Mexican engineers and scientists find employment in the high-technology companies in Silicon Valley.

<sup>&</sup>lt;sup>17</sup> Espenshade and Goodis (1985) report that as of 1980, Mexican immigrant workers made up nearly 50 percent of the low-skilled manufacturing labor force in the standard metropolitan statistical area (SMSA).

<sup>&</sup>lt;sup>18</sup> I also tested interaction effects to see whether the linear associations between human capital and other control variables and the hourly wage differ based on the nativity status and naturalization status. The results are presented in appendix.

This chapter presented findings related to the assessment of the cost of being a Mexican immigrant and being a Mexican non-citizen in California and Texas, and a series of supplementary analyses. The next chapter presents a summary of the major findings, theoretical contributions of this study to Latino/a demography, implications of the findings to the immigrant policies, shortcomings and future directions of this study.

## CHAPTER V

## **CONCLUSIONS**

For the comprehensive examination of the heterogeneity of Mexican-origin workers, this study focused on internal hourly wage differences within this group. Instead of estimating the Mexican-Anglo wage differences, which is the major approach guided by the ethnic labor market perspectives, I examined the labor market experiences of Mexican immigrants and how social context matters. Specifically, I focused on the cost of being an immigrant (the hourly wage differences between the native- and foreign-born) and the cost of being an immigrant without U.S. citizenship (the wage differences between naturalized immigrants and non-naturalized immigrants). By comparing the two largest concentration states of Mexican-origin workers in the United States, California and Texas, I also examined the impact of the state-specific labor market characteristics on Mexican immigrant workers' wages.

The results consistently show that foreign-born status and the lack of U.S. citizenship are correlated with lower hourly wages across the states. Reflecting anti-immigrant policies and sentiments, Mexican immigrants in California bear a higher cost for being foreign-born than their counterparts in Texas. Furthermore, non-U.S. naturalized foreign-born Mexicans face dual disadvantages with respect to wages, especially for those living in California. Partly, larger population concentration of immigrants, especially non-citizens, could be a source of intensive within-group labor market competition among the foreign-born workers.

The cost of being an immigrant was further examined in terms of immigrants' length of U.S. residence. Results show that immigrants who arrived during the last decade faced significant declines in wages across the states. Furthermore, the greater cost attached to this

cohort in California confirms the finding from the initial analysis that Mexican immigrants faced harsher social contexts in California in the post-IRCA period, in which foreign-born status became a stronger determinant of wages over human capital differences. U.S.-naturalized citizenship status is another central factor associated with foreign-born workers' wages, controlling for the passage of time in the United States.

Findings also support the importance of other individual attributes on wages.

Regarding the selected human capital resources, education and English-language proficiency are associated with higher wage returns. Especially, college completion appears to be a principal factor of within-group wage differences. Lack of education will continue to block Mexican immigrants with lower levels of education, especially non-U.S. citizens, from access to better-paying jobs. Moreover, the results show significant wage differences between professional and non-professional occupational categories.

Major theoretical and empirical contributions of this study to Latino/a demography are noted. First, the findings show that nativity status (native- and foreign-born) and U.S. naturalized citizenship status (foreign-born with citizenship and non-citizen foreign-born status) are two major determinants of within-group wage differences, which are often missing in the labor market analysis of Mexican-origin workers (Bean *et al.*, 1988; Bradshaw and Frisbie, 1983; Saenz, 2004b; Semyonov, 1988). Further analysis of the different labor market experiences within this group is encouraged, as the Mexican-origin population includes a number of historically disadvantaged foreign-born as well as continuing inflows of new immigrants. Second, findings show the importance of taking into account the effect of broad social contexts on wages beyond different types of labor market categories. As past studies tend to rely on small-scale case studies limited to specific occupations and industries,

further analysis on the impact of regional differences on labor market experiences of Mexican-origin workers are needed.

The results of this study raise several important immigrant policy implications.

Unlike the anti-immigrant sentiments against Mexican workers in the background of IRCA and California propositions, the findings strongly suggest that Mexican immigrants, particularly non-citizens, faced harsher labor market conditions after the implementation of IRCA. As the results show the importance of education and English language proficiency on wages, social welfare policies (e.g., public education for Limited-English-Proficiency (ELP) Mexicans) have to be reconsidered to avoid further internal wage gap based on lower levels of education and language problem. There is a great possibility that a disproportionate number of foreign-born Mexicans, especially non-citizens, stand in a disadvantaged labor market position due to a lack of human capital resources.

Furthermore, the negative impact of IRCA on non-citizen Mexican workers' wages indicates that current discussion on the revival of temporary guest worker program and a legalization program possibly result in wage penalties against non-citizen workers. Future immigrant policies should be carefully reformed by taking into account whether future policies really bring benefits to immigrant workers who are already in this country.

Several shortcomings of this study are also noted. First, the analysis did not control for the differences in living expenses and state minimum wages between California and Texas. Second, as the major focuses of this study were immigrant status and naturalization status, other individual attributes such as human capital resources were treated as control variables and thus not widely discussed. As almost all variables included in the analysis have

significant effects on wages, further analysis has to be conducted to closely estimate the effect of various attributes on wages.

The third limitation is attributed to the data set used to conduct the analysis. In separating the immigrant population into the U.S.-naturalized citizens and non-citizens, the latter contains a wide variety of persons including undocumented immigrants and international college students. As undocumented workers may exert a large impact on the wages of other individuals (Bean *et al.*, 1988; Briggs, 1983), the inseparable undocumented portion in the immigrant population in the PUMS data is certainly an unavoidable problem in this study. Also, skilled immigrants may not be permanent residents or naturalized citizens, who may hold temporary visas (the H-1B). Although I am interested in assessing the labor market outcomes for immigrants working in professional occupations, few, if any data sources, allow us to determine whether they remain in the United States for employment or return to their countries (*see* Bayer, 1968).

The PUMS data are widely used in earnings studies as they provide a large, nationally representative sample of all sectors of the labor force. Nonetheless, considering that there is a certain portion of foreign-born Mexicans involved in the underground economy and informal sectors, the cost of being a non-citizen could be much greater than the findings of this study suggest.

Fourth, I attributed the higher costs associated with foreign-born status and noncitizenship status in California to the anti-immigrant hostilities due to a higher population concentration of immigrants in that state. However, other possible factors associated with higher costs in California, industrial changes taking place between California and Texas, have not been discussed in this study. For example, the literature notes that immigrants in California faced declining wages during the 1990s due to a broad restructuring of the economy in California (Durand *et al.*, 2000), while Texas experienced an effective labor demand due to rural industrialization.

Finally, future directions of this study are noted. This study estimated the cost of being an immigrant and being a non-citizen in California and Texas, two major concentration states in the Southwest. Further analysis needs to be conducted for testing the generalization of these costs in much broader social and economic contexts, such as in different standard Census regions and the entire United States. Particularly, future research needs to examine the labor market experiences of Mexican-origin workers in new-destination areas, places primarily in the South and Midwest where Mexicans, especially immigrants, have settled over the last decade.

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#### **APPENDIX**

## **INTERACTION EFFECTS**

The literature shows that immigrant status and naturalization status have an effect on the degree that human capital affects wages. In fact, sociological theories often imply that a linear association between a predictor and a dependent variable is affected by the third variable (Blalock, 1965; Aiken and West, 1991). For example, a linear association between the natural logarithm of the hourly wage and education may differ depending on a Mexican worker's English language proficiency. Interaction is a term in a statistical model in which the effect of two or more independent variables is not simply additive. Testing for the presence of interaction is one way to reduce the effects of unexplained variables. The interaction is tested by including a cross-product term of  $x_1$  and  $x_2$  under consideration (e.g., English language proficiency and education) in a multiple regression (Allison, 1977). Thus, for a response variable y and two independent variables  $x_1$  and  $x_2$  an additive model would be:

$$y = ax_1 + bx_2 + \text{error},$$

while

$$y = ax_1 + bx_2 + c(x_1 \times x_2) + \text{error},$$

is an example of a model with an interaction between variables  $x_1$  and  $x_2$ .

The interaction term in the full model (all variables included in the model plus an interaction term) tests the following null and alternative hypotheses:

H<sub>0</sub>:  $\beta_1 = \beta_2 = \beta_3 = \cdots = \beta_k$  (all slopes are equal, or linear relations between the outcome variable and independent variable are same across values of a third variable).

Ha: at least one slope differs.

If the t-value for the interaction is significant, or if the null hypothesis is rejected, we assume that slopes are not equal, and thus there is interaction. If we fail to reject the null hypothesis, we assume equal slopes and drop the interaction term (there is not significant evidence that the change is not the same for the slopes). This model is called "reduced model" or "parallel line model" as we refit the model without interaction.

The analysis of human capital attributes and other factors shows that metropolitan area residence, employment in management, professional and related occupations, Englishlanguage ability and education are strong factors associated with higher wage returns. Therefore, I examine whether the linear associations between the natural logarithm of the hourly wage and these variables differ based on the nativity difference and immigrants' possession of U.S. citizenship. Namely, the following interactions are tested: (1) immigrant status and metropolitan area residence (imm\*metrores) and (2) immigrant status and management, professional and related occupations (**imm\*mgrprorl**) in Model 1A and 1B, and (3) self-reported English language proficiency and education (engabil\*somehs, engabil\*hsgrad, engabil\*somecoll, and engabil\*collgrad), (4) non-naturalization status and metropolitan area residence (**nocit\*metrores**), and (5) non-naturalization status and management, professional and related occupations (**nocit\*mgrprorl**) in Model 2A and 2B. However, two of the interaction terms, **imm\*metrores** and **nocit\*mgrprorl** are dropped from the models as they have strong collineality with **imm** and **nocit**, respectively, and thus make these two interaction terms statistically insignificant. Selected findings are presented below and in Table 12.

Regarding **imm\*mgrprorl** (t = -3.69, p < 0.0002 in Model 1A and t = -3.07, p < 0.0021 in Model 1B), the wage difference between foreign-born and native-born (**imm**)

Table 12. Selected Results of Interaction Tests

|   | California | Texas  |
|---|------------|--------|
| Cost of being a Mexican immigrant employed in management, professional, and related occupations | 9%**       | 5%*    |
| Wage penalty of Mexican immigrants who have high school diploma but do not speak English        | 17%**      | 11%    |
| Wage penalty of Mexican immigrants who have bachelor's degree but do not speak English          | 48%***     | 50%*** |
| The cost of being a Mexican non-citizen in metropolitan areas                                   | 17%***     | 13%**  |

<sup>\*\*\*</sup>Significant at the 0.0001 level.

Source: 2000 5% PUMS.

employed in management, professional and related occupations (**mgrprorl**) are estimated. In Model 1A, when the person is a foreign-born worker (**imm** = 1) employed in management, professional, and related occupations (**mgrprorl** = 1),

lghrwage = 
$$1.13468$$
 (intercept) +  $0.46655$  (**mgrprorl**) -  $0.05630$  (**imm**) -  $0.03171$  (**imm\*mgrprorl**) =  $1.51322$ , and

when the person is a native-born worker ( $\mathbf{imm} = 0$ ) employed in management, professional, and related occupations ( $\mathbf{mgrprorl} = 1$ ),

lghrwage = 
$$1.13468$$
 (intercept) +  $0.46655$  (mgrprorl) +  $0$  (imm) +  $0$  (imm\*mgrprorl) =  $1.60123$ 

$$\exp^{1.60123 - 1.51322} = \exp^{0.08801} = 1.0920$$

<sup>\*\*</sup>Significant at the 0.0005 level.

<sup>\*</sup>Significant at the 0.005 level.

In California, a Mexican immigrant who is employed in management, professional and related occupations has a 9 percent lower hourly wage than a native-born Mexican American who is employed in this occupational category.

In Model 1B, on the other hand, an immigrant who is employed in this occupational category has a 5 percent lower hourly wage than a native-born employed in this occupational category: when the person is a foreign-born worker (**imm** = 1) employed in management, professional, and related occupations (**mgrprorl** = 1),

lghrwage = 1.09299 (intercept) + 0.46639 (**mgrprorl**) – 0.01191 (**imm**) – 0.03628 (**imm\*mgrprorl**) = 1.51119, and

when the person is a native-born worker ( $\mathbf{imm} = 0$ ) employed in management, professional, and related occupations ( $\mathbf{mgrprorl} = 1$ ),

lghrwage = 1.09299 (intercept) + 0.46639 (mgrprorl) + 0 (imm) + 0 (imm\*mgrprorl) = 1.55938

$$\exp^{1.55938 - 1.51119} = \exp^{0.04819} = 1.0494$$

These wage differences (9% in Model 1A and 5% in 1B) are statistically significant because the interaction term **imm\*mgrprorl** is statistically significant. The interaction test shows that being employed in management, professional, and related occupations does not cancel the cost of being an immigrant across both states, confirming the major argument of this study that being an immigrant is a central factor associated with internal wage gaps within the Mexican-origin population.

Five interaction terms included in the immigrant-only sample are assessed in the same manner, and three selected findings are reported. Regarding the **engabil\*hsgrad** interaction in Model 2A, immigrants who speak English and whose highest educational attainment is

high school graduation (**engabil** = 1 and **hsgrad** = 1) have 17 percent ( $\exp^{0.15603}$ ) higher wages than immigrants whose highest educational attainment is high school graduation but do not speak English (**engabil** = 0 and **hsgrad** = 1). However, the variable is not statistically significant in Model 2B (t = 1.16, p = 0.2457).

Regarding the **engabil\*collgrad** variable in Model 2A, immigrants who speak English and whose highest educational attainment is college completion (**engabil** = 1 and **collgrad** = 1) have 48 percent (exp<sup>0.39192</sup>) higher wages than immigrants whose highest educational attainment is college completion but do not speak English (**engabil** = 0 and **collgrad** = 1). Model 2B also shows a similar pattern; immigrants who speak English and whose highest educational attainment is college completion have 50 percent (exp<sup>0.40613</sup>) higher wages than immigrants whose highest educational attainment is college completion but do not speak English. The interaction test shows that wage returns to education are significantly affected by the immigrant's English language ability.

The last interaction term included in the immigrant-only models is **nocit\*metrores**. Descriptive statistics (Table 2) shows that about 90 percent of the naturalized citizens and non-naturalized citizens were metropolitan area residents in California. In Texas, 67 percent of naturalized citizens and 75 percent of non-naturalized citizens were metropolitan area residents. As metropolitan area residence is positively associated with hourly wages, I assess the different effect of metropolitan area residence on wages between naturalized citizens and non-naturalized citizens. In Model 2A, non-citizens (**nocit** = 1) in metropolitan areas (**metrores** = 1) have 17 percent (exp<sup>0.15745</sup>) lower wages than naturalized citizens in metropolitan areas. In Model 2B, on the other hand, non-citizens in metropolitan areas have 13 percent (exp<sup>0.11795</sup>) lower wages than naturalized citizens in metropolitan areas. These

results indicate that residing in metropolitan areas does not cancel the cost of being a noncitizen both in California and Texas.

The results support the literature discussing that immigrant status and naturalization status have an effect on the degree that human capital affects wages. Mexican immigrants with the same types of occupations as native-born Mexican Americans attain lower wages than their native-born counterparts. The same is true of language ability; those who speak English attain higher wages than their non-English-speaking, but equally well-educated counterparts. It is likely that the ability to speak English is also an important status marker. The results of interaction tests confirm the major argument of this study that being a foreign-born and being a non-citizen are the two major factors of the internal wage gaps among the Mexican-origin workers.

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